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Article

On the Convergence of Randomized Block Kaczmarz Algorithm for Solving Matrix Equation

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Abstract: Randomized block Kaczmarz Method and randomized extended block Kaczmarz Method are proposed for solving matrix equation AXB = C, where the matrix A and B may be full rank or rank deficient. These methods are iterative methods without matrix multiplication, and are especially suitable for solving large-scale matrix equations. It is theoretically proved these methods converge to the solution or least square solution of the matrix equation. The numerical results show that these methods are more efficient than the existing algorithms for high-dimensional matrix equations.

Keywords: matrix equation; randomized block Kaczmarz; randomized extended block Kaczmarz; convergence

MSC: 65F10; 65F45; 65H10

1. Introduction

Consider the linear matrix equation

$$AXB = C, (1)$$

where $A \in \mathbb{R}^{m \times p}$, $B \in \mathbb{R}^{q \times n}$ and $C \in \mathbb{R}^{m \times n}$. Such problems arise in linear control and filtering theory for continuous or discrete-time large-scale dynamical systems. If AXB = C is consistent, $X^* = A^{\dagger}CB^{\dagger}$ is the minimum Frobenius norm solution. If AXB = C is inconsistent, $X^* = A^{\dagger}CB^{\dagger}$ is the minimum Frobenius norm least squares solution. When the matrices A and B are small and dense, direct methods based on QR-fraction are attractive [1,2]. However, for large matrices A and B, iterative methods have attracted much attention [3–7]. Recently, Du et al. proposed the randomized block coordinate descent (RBCD) method for solving the matrix least-squares problem $\min_{X \in \mathbb{R}^{p \times q}} \|C - AXB\|_F$ without strong convexity assumption in [8]. This method requires that matrix B is full row rank. Wu et al. [9] introduced two kinds of Kaczmarz-type methods to solve consistent matrix equation AXB = C: relaxed greedy randomized Kaczmarz (ME-RGRK) and maximal weighted residual Kaczmarz (ME-MWRK). Although the row and column index selection strategy is time-consuming, the ideas of these two methods are suitable for solving large-scale consistent matrix equations.

In this paper, randomized Kaczmarz method [10] and randomized extended Kaczmarz method [11] are used to solve consistent and inconsistent matrix equations (1) by the product of matrix and vector.

All the results in this paper hold in the complex field. But for the sake of simplicity, we only discuss it in the real number field.

In this paper, we denote A^T , A^\dagger , r(A), R(A), $\|A\|_F = \sqrt{\operatorname{trace}(A^TA)}$ and $\langle A, B \rangle_F = \operatorname{trace}(A^TB)$ as the transpose, the Moore-Penrose generalized inverse, the rank of A, the column space of A, the Frobenius norm of A and the inner product of two matrices A and B, respectively. For an integer $n \geq 1$, let $[n] = \{1, 2, \cdots, n\}$. We use I to denote the identity matrix whose order is clear from the context. In addition, for a given matrix $G = (g_{ij}) \in \mathbb{R}^{m \times n}$, $G_{i,:}$, $G_{:,j}$, $\sigma_{\max}(G)$ and $\sigma_{\min}(G)$, are used to denote ith row, jth column, the maximum singular value and the smallest nonzero singular value of G respectively. Let E_k denote the expected value conditional on the first k iterations, that is, $\mathbb{E}_k[\cdot] = \mathbb{E}[\cdot|i_0,j_0,i_1,j_1,...,i_{k-1},j_{k-1}]$, where i_s and $j_s(s=0,1,...,k-1)$ are the row and the



column chosen at the sth iteration. Let the conditional expectations with respect to the random row index be $\mathbb{E}^i_k[\cdot] = \mathbb{E}[\cdot|i_0,j_0,i_1,j_1,...,i_{k-1},j_{k-1},j_k]$ and with respect to the random column index be $\mathbb{E}^j_k[\cdot] = \mathbb{E}[\cdot|i_0,j_0,i_1,j_1,...,i_{k-1},j_{k-1},i_k]$. By the law of total expectation, it holds that $\mathbb{E}_k[\cdot] = \mathbb{E}^i_k[\mathbb{E}^j_k[\cdot]]$.

The organization of this paper is as follows. In Section 2, we will discuss Kaczmarz method (ME-RBK and ME-PRBK) for finding the minimal F-norm solution ($A^{\dagger}CB^{\dagger}$) of the consistent matrix Eq. (1). In Section 3, we discuss the extended Kaczmarz method (IME-REBK) for finding the minimal F-norm least-squares solution of the matrix Eq. (1). In Section 4, some numerical examples are provided to illustrate the effectiveness of our new methods. Finally, some brief concluding remarks are described in Section 5.

2. The Randomized Block Kaczmarz Method for Consistent Equation

At the kth iteration, Kaczmarz method selects randomized a row $i \in [m]$ of A and does an orthogonal projection of the current estimate matrix $X^{(k)}$ onto the corresponding hyperplane $H_i = \{X \in \mathbb{R}^{p \times q} | A_{i,:} XB = C_{i,:} \}$, that is

$$\min_{X \in \mathbb{R}^{p \times q}} \frac{1}{2} \|X - X^{(k)}\|_F^2 \ s.t. \ A_{i,:} XB = C_{i,:}.$$
 (2)

The Lagrangian function of the conditional optimization problem (2) is

$$L(X,Y) = \frac{1}{2} ||X - X^{(k)}||_F^2 + \langle Y, A_{i,:}XB - C_{i,:} \rangle,$$
(3)

where $Y \in \mathbb{R}^{1 \times n}$ is Lagrangian multiplier. By the matrix differentiation, we get the gradient of L(X,Y) and set $\nabla L(X,Y) = 0$ for finding the stationary matrix:

$$\begin{cases}
\nabla_X L(X,Y)|_{X^{(k+1)}} = X^{(k+1)} - X^{(k)} + A_{i,:}^T Y B^T = 0, \\
\nabla_Y L(X,Y)|_{X^{(k+1)}} = A_{i,:} X^{(k+1)} B - C_{i,:} = 0.
\end{cases}$$
(4)

By the first equation of (4), we have $X^{(k+1)} = X^{(k)} - A_{i,:}^T Y B^T$. Substituting this into the second equation of (4), we can get $Y = -\frac{1}{\|A_{i,:}\|_2^2} (C_{i,:} - A_{i,:} X^{(k)}) (B^T B)^{\dagger}$. So the projected randomized block Kacmarz (ME-PRBK) for solving AXB = C iterates as

$$X^{(k+1)} = X^{(k)} + \frac{A_{i,:}^{T}}{\|A_{i,:}\|_{2}^{2}} (C_{i,:} - A_{i,:} X^{(k)} B) B^{\dagger}.$$
 (5)

However, in practice, it is very expensive to calculate the pseudoinverse of large-scale matrices. Next, we generalize the average block Kaczmarz method [12] for solving linear equations to matrix equations.

At the kth step, we get the approximate solution $X^{(k+1)}$ by projecting the current estimate $X^{(k)}$ onto the hyperplane $H_{i,j} = \{X \in \mathbb{R}^{p \times q} | A_{i,:} X^{(k)} B_{:,j} = C_{i,j}\}$. Using the Lagrangian multiplier method, we can obtain the following Kaczmarz method for AXB = C:

$$X^{(k+1)} = X^{(k)} + \frac{A_{i,:}^T (C_{i,:} - A_{i,:} X^{(k)} B_{:,j}) B_{:,j}^T}{\|A_{i,:}\|_2^2 \|B_{:,j}\|_2^2}.$$
 (6)

Inspired by the idea of average block Kaczmaz algorithm for Ax = b, we consider the average block Kaczmaz method for AXB = C respect to B.

$$X^{(k+1)} = X^{(k)} + \lambda \frac{A_{i,:}^T}{\|A_{i,:}\|_2^2} \sum_{j=1}^n v_j^k \frac{(C_{i,:} - A_{i,:} X^{(k)} B_{:,j}) B_{:,j}^T}{\|B_{:,j}\|_2^2}, \tag{7}$$

where λ is stepsize and v_j^k be the weights that satisfy $v_j^k \geq 0$ and $\sum_{i=1}^n v_j^k = 1$. If $v_j^k = \frac{\|B_{i,j}\|_2^2}{\|B\|_F^2}$, then

$$X^{(k+1)} = X^{(k)} + \frac{\lambda}{\|B\|_F^2} \frac{A_{i,:}^T}{\|A_{i,:}\|_2^2} (C_{i,:} - A_{i,:} X^{(k)} B) B^T.$$
(8)

Set $\alpha = \frac{\lambda}{\|B\|_T^2} > 0$, we obtain the following randomized block Kaczmarz iteration

$$X^{(k+1)} = X^{(k)} + \frac{\alpha}{\|A_{i::}\|_{2}^{2}} A_{i::}^{T} \left((C_{i::} - (A_{i::}X^{(k)})B)B^{T} \right), k = 0, 1, 2, \dots,$$
(9)

where i is selected with probability $p_i = \frac{\|A_{i:}\|_2^2}{\|A\|_F^2}$. The cost of each iteration of this method is 4q(n+p) + p+1-2q if the square of the row norm of \ddot{A} has been calculated in advance. We describe this method as Algorithm 1, which is called the ME-RBK algorithm.

Algorithm 1 Randomized Block Kaczmarz Method for AXB = C (ME-RBK)

Input: $A \in \mathbb{R}^{m \times p}$, $B \in \mathbb{R}^{q \times n}$, $C \in \mathbb{R}^{m \times n}$, $X^{(0)} = 0 \in \mathbb{R}^{p \times q}$

- 1: **for** k = 0, 1, 2, ...,**do**
- Pick *i* with probability $p_i(A) = \frac{\|A_{i,:}\|_2^2}{\|A\|_F^2}$ Compute $X^{(k+1)} = X^{(k)} + \frac{\alpha}{\|A_{i,:}\|_2^2} A_{i,:}^T \left((C_{i,:} (A_{i,:}X^{(k)})B)B^T \right)$

Remark 1. Note that the problem of finding a solution of AXB = C can be posed as the following linear least squares problem

$$\min_{X \in \mathbb{R}^{p \times q}} \frac{1}{2} \|AXB - C\|_F^2 = \min_{X \in \mathbb{R}^{p \times q}} \frac{1}{2} \sum_{i=1}^m \|A_{i,i}XB - C_{i,i}\|_2^2.$$
 (10)

Define the component function

$$f_i(X) = \frac{1}{2} ||A_{i,:}XB - C_{i,:}||_2^2,$$

then differentiate with X to get its gradient

$$\nabla f_i(X) = A_{i,:}^T (A_{i,:} XB - C_{i,:}) B^T.$$

Therefore, the randomized block Kaczmarz method (9) is equivalent to one step of the stochastic gradient decent method [13] applied to (10) with stepsize $\frac{\alpha}{\|A_i\|_2^2}$.

First, we give the following lemma whose proof can be found in [8].

Lemma 1. [8] Let $A \in \mathbb{R}^{m \times p}$ and $B \in \mathbb{R}^{q \times n}$ be any nonzero matrix. Let

$$\mathcal{M} = \{ M \in R^{p \times q} \mid \exists Y \in \mathbb{R}^{m \times n} \text{ s.t. } M = A^T Y B^T \}.$$

For any matrix $M \in \mathcal{M}$, it holds

$$||AMB||_F^2 \ge \sigma_{\min}^2(A)\sigma_{\min}^2(B)||M||_F^2$$
.

Remark 2. $M \in \mathcal{M}$ means that $M_{:,j} \in R(A^T)$, j = 1, 2, ..., q and $(M_{i,:})^T \in R(B)$, i = 1, 2, ..., p. In fact, \mathcal{M} is well defined because $0 \in \mathcal{M}$ and $A^{\dagger}CB^{\dagger} \in \mathcal{M}$.

In the following theorem, with the idea of the RK method [10], we will prove that Algorithm 1 converges to the least *F*-norm solution of AXB = C.

Theorem 1. Assume $0 < \alpha < \frac{2}{\|B\|_2^2}$. If Eq. (1) is consistent, the sequence $\{X^{(k)}\}$ generated by the ME-RBK method starting from the initial matrix $X^{(0)} \in \mathbb{R}^{p \times q}$ in which $X^{(0)}_{:,j} \in R(A^T)$, $j = 1, 2, \ldots, q$ and $(X^{(0)}_{i,:})^T \in R(B)$, $i = 1, 2, \ldots, p$, converges linearly to $A^{\dagger}CB^{\dagger}$ in mean square form. Moreover, the solution error in expectation for the iteration sequence $X^{(k)}$ obeys

$$E\left[\left\|X^{(k)} - A^{\dagger}CB^{\dagger}\right\|_{F}^{2}\right] \le \rho^{k} \left\|X^{(0)} - A^{\dagger}CB^{\dagger}\right\|_{F}^{2},\tag{11}$$

where $\rho = 1 - \frac{2\alpha - \alpha^2 \|B\|_2^2}{\|A\|_F^2} \sigma_{\min}^2(A) \sigma_{\min}^2(B)$, and $i \in [m]$ picked with probability $p_i(A) = \frac{\|A_{i,:}\|_2^2}{\|A\|_F^2}$.

Proof. For k = 0, 1, 2, ..., by (9) and $A_{i,:}A^{\dagger}CB^{\dagger}B = C_{i,:}$ (consistency), we have

$$\begin{split} X^{(k+1)} - A^{\dagger}CB^{\dagger} &= X^{(k)} + \frac{\alpha}{\|A_{i,:}\|_{2}^{2}} A_{i,:}^{T} (C_{i,:} - A_{i,:}X^{(k)}B)B^{T} - A^{\dagger}CB^{\dagger} \\ &= (X^{(k)} - A^{\dagger}CB^{\dagger}) - \frac{\alpha}{\|A_{i,:}\|_{2}^{2}} A_{i,:}^{T}A_{i,:}(X^{(k)} - A^{\dagger}CB^{\dagger})BB^{T}, \end{split}$$

then

$$||X^{(k+1)} - A^{\dagger}CB^{\dagger}||_{F}^{2} = ||X^{(k)} - A^{\dagger}CB^{\dagger}||_{F}^{2} + \frac{\alpha^{2}}{||A_{i,:}||_{2}^{4}} ||A_{i,:}^{T}A_{i,:}(X^{(k)} - A^{\dagger}CB^{\dagger})BB^{T}||_{F}^{2}$$
$$- \frac{2\alpha}{||A_{i,:}||_{2}^{2}} \langle X^{(k)} - A^{\dagger}CB^{\dagger}, A_{i,:}^{T}A_{i,:}(X^{(k)} - A^{\dagger}CB^{\dagger})BB^{T} \rangle_{F}.$$

It follows from

$$\begin{split} &\frac{\alpha^2}{\|A_{i,:}\|_2^4} \|A_{i,:}^T A_{i,:}(X^{(k)} - A^{\dagger}CB^{\dagger})BB^T\|_F^2 \\ &= \frac{\alpha^2}{\|A_{i,:}\|_2^2} \|A_{i,:}(X^{(k)} - A^{\dagger}CB^{\dagger})BB^T\|_2^2 \text{ (by trace}(uu^T) = \|u\|_2^2 \text{ for any vector } u) \\ &\leq \frac{\alpha^2 \|B\|_2^2}{\|A_{i,:}\|_2^2} \|A_{i,:}(X^{(k)} - A^{\dagger}CB^{\dagger})B\|_2^2 \text{ (by } \|u^TB^T\|_2 = \|Bu\|_2 \leq \|B\|_2 \|u\|_2), \end{split}$$

and

$$\begin{split} &\frac{2\alpha}{\|A_{i,:}\|_2^2} \langle X^{(k)} - A^\dagger C B^\dagger, A_{i,:}^T A_{i,:} (X^{(k)} - A^\dagger C B^\dagger) B B^T \rangle_F \\ &= \frac{2\alpha}{\|A_{i,:}\|_2^2} \mathrm{trace} (B^T (X^{(k)} - A^\dagger C B^\dagger)^T A_{i,:}^T A_{i,:} (X^{(k)} - A^\dagger C B^\dagger) B) (\text{ by trace}(MN) = \mathrm{trace}(NM)) \\ &= \frac{2\alpha}{\|A_{i,:}\|_2^2} \|A_{i,:} (X^{(k)} - A^\dagger C B^\dagger) B\|_2^2 \end{split}$$

that

$$\|X^{(k+1)} - A^{\dagger}CB^{\dagger}\|_F^2 \le \|X^{(k)} - A^{\dagger}CB^{\dagger}\|_F^2 - \frac{2\alpha - \alpha^2 \|B\|_2^2}{\|A_{i,:}\|_2^2} \|A_{i,:}(X^{(k)} - A^{\dagger}CB^{\dagger})B\|_2^2.$$

By taking the conditional expectation, we have

$$\begin{split} E_k \left[\left\| X^{(k+1)} - A^{\dagger} C B^{\dagger} \right\|_F^2 \right] &\leq \left\| X^{(k)} - A^{\dagger} C B^{\dagger} \right\|_F^2 - E_k \left[\frac{2\alpha - \alpha^2 \|B\|_2^2}{\|A_{i,:}\|_2^2} \|A_{i,:} (X^{(k)} - A^{\dagger} C B^{\dagger}) B\|_2^2 \right] \\ &= \left\| X^{(k)} - A^{\dagger} C B^{\dagger} \right\|_F^2 - \frac{2\alpha - \alpha^2 \|B\|_2^2}{\|A\|_F^2} \|A(X^{(k)} - A^{\dagger} C B^{\dagger}) B\|_F^2. \end{split}$$

By $X^{(0)} \in \mathcal{M}$ and $A^{\dagger}CB^{\dagger} \in \mathcal{M}$, we have $X^{(0)} - A^{\dagger}CB^{\dagger} \in \mathcal{M}$. Noting $A_{i,:}^{T} = A^{T}I_{:,i}$, it is easy to show that $X^{(k)} - A^{\dagger}CB^{\dagger} \in \mathcal{M}$ by induction. Then by Lemma 1 and $0 < \alpha < \frac{2}{\|B\|_{2}^{2}}$, we can get

$$E_{k} \left[\left\| X^{(k+1)} - A^{\dagger} C B^{\dagger} \right\|_{F}^{2} \right] \leq \left\| X^{(k)} - A^{\dagger} C B^{\dagger} \right\|_{F}^{2} - \frac{2\alpha - \alpha^{2} \|B\|_{2}^{2}}{\|A\|_{F}^{2}} \sigma_{\min}^{2}(A) \sigma_{\min}^{2}(B) \left\| X^{(k)} - A^{\dagger} C B^{\dagger} \right\|_{F}^{2}$$

$$= \left(1 - \frac{2\alpha - \alpha^{2} \|B\|_{2}^{2}}{\|A\|_{F}^{2}} \sigma_{\min}^{2}(A) \sigma_{\min}^{2}(B) \right) \left\| X^{(k)} - A^{\dagger} C B^{\dagger} \right\|_{F}^{2}. \tag{12}$$

Finally, by ((12)) and induction on the iteration index k, we obtain the estimate ((11)). \Box

Remark 3. By the similar approach as used in the proof of Theorem 1, we can prove that the iterate $X^{(k)}$ generated by ME-PRBK (5) satisfies the following estimate

$$E\left[\left\|X^{(k)}-A^{\dagger}CB^{\dagger}\right\|_{F}^{2}\right] \leq \hat{\rho}^{k}\left\|X^{(0)}-A^{\dagger}CB^{\dagger}\right\|_{F}^{2},$$

where $\hat{\rho} = 1 - \frac{\sigma_{min}^2(A)\sigma_{min}^2(B)}{\|A\|_F^2\|B\|_2^2}$. The convergence factor of GRK in [14] is $\rho_{GRK} = 1 - \frac{\sigma_{min}^2(A)\sigma_{min}^2(B)}{\|A\|_F^2\|B\|_F^2}$. It is obvious that

$$\rho_{GRK} > \min_{0 < \alpha < \frac{2}{\|B\|_2^2}} \rho = 1 - \frac{\sigma_{\min}^2(A)\sigma_{\min}^2(B)}{\|A\|_F^2 \|B\|_2^2} = \hat{\rho}$$

and $\rho < \rho_{GRK}$ when $\frac{1-\sqrt{1-\frac{\|B\|_2^2}{\|B\|_F^2}}}{\|B\|_2^2} < \alpha < \frac{1+\sqrt{1-\frac{\|B\|_2^2}{\|B\|_F^2}}}{\|B\|_2^2}$. This means that the convergence factor of ME-PRBK is the smallest and the factor of ME-RBK can be smaller than that of GRK when α is properly selected.

3. The Randomized Extended Block Kaczmarz Method for Inconsistent Equation

In [11,15,16], the authors proved that the Kaczmarz method does not converge to the least squares solution of AX = b when AX = b is inconsistent. Analogically, if the matrix Eq. (1) is inconsistent, the above ME-PRBK method dose not converge to $A^{\dagger}CB^{\dagger}$. The following theorem give the error bound of the for the inconsistent matrix equation.

Theorem 2. Assume that the consistent equation AXB = C has a solution $X^* = A^{\dagger}CB^{\dagger}$. Let $\hat{X}^{(k)}$ denote the kth iterate of the ME-PRBK method applied to the inconsistent AXB = C + W for any $W \in \mathbb{R}^{m \times n}$ starting from the initial matrix $X^{(0)} \in \mathbb{R}^{p \times q}$ in which $X^{(0)}_{:,j} \in \mathbb{R}(A^T)$, $j = 1, 2, \ldots, q$ and $(X^{(0)}_{i,:})^T \in \mathbb{R}(B)$, $i = 1, 2, \ldots, p$. In exact arithmetic, it follows that

$$E\left[\left\|\hat{X}^{(k)} - A^{\dagger}CB^{\dagger}\right\|_{F}^{2}\right] \leq \rho^{k} \left\|X^{(0)} - A^{\dagger}CB^{\dagger}\right\|_{F}^{2} + \frac{1 - \rho^{k}}{1 - \rho} \frac{\|WB^{\dagger}\|_{F}^{2}}{\|A\|_{F}^{2}},\tag{13}$$

Proof. Set $H_i = \{X | A_i XB = C_i\}$, $\hat{H}_i = \{X | A_i XB = C_i + W_i\}$. Let Y denote the iterate of the PRBK method applied to the consistent AXB = C at kth step, that is

$$Y = \hat{X}^{(k)} + \frac{A_{i,:}^T}{\|A_{i::}\|_2^2} (C_{i,:} - A_{i,:} \hat{X}^{(k)} B) B^{\dagger}.$$

It follows from

$$\langle Y - A^{\dagger}CB^{\dagger}, \hat{X}^{(k+1)} - A^{\dagger}CB^{\dagger} \rangle_{F} = \langle Y - A^{\dagger}CB^{\dagger}, \frac{A_{i,:}^{I}}{\|A_{i,:}\|_{2}^{2}} W_{i}B^{\dagger} \rangle_{F}$$

$$= \operatorname{trace} \left((B^{\dagger})^{T} W_{i}^{T} \frac{A_{i,:}}{\|A_{i,:}\|_{2}^{2}} (Y - A^{\dagger}CB^{\dagger}) \right)$$

$$= \operatorname{trace} \left(W_{i}^{T} \frac{A_{i,:}YB - A_{i,:}A^{\dagger}CB^{\dagger}B}{\|A_{i,:}\|_{2}^{2}} (B^{T}B)^{\dagger} \right) = 0$$

and

$$\left\|\hat{X}^{(k+1)} - Y\right\|_{F}^{2} = \frac{1}{\|A_{i,:}\|_{2}^{4}} \left\|A_{i,:}^{T} W_{i} B^{\dagger}\right\|_{F}^{2} = \operatorname{trace}\left((B^{\dagger})^{T} W_{i}^{T} A_{i,:} A_{i,:}^{T} W_{i} B^{\dagger} \frac{1}{\|A_{i,:}\|_{2}^{2}}\right) = \frac{\left\|W_{i} B^{\dagger}\right\|_{2}^{2}}{\|A_{i,:}\|_{2}^{2}}$$

that

$$\begin{aligned} \left\| \hat{X}^{(k+1)} - A^{\dagger} C B^{\dagger} \right\|_{F}^{2} &= \left\| Y - A^{\dagger} C B^{\dagger} \right\|_{F}^{2} + \left\| \hat{X}^{(k+1)} - Y \right\|_{F}^{2} \\ &= \left\| Y - A^{\dagger} C B^{\dagger} \right\|_{F}^{2} + \frac{\left\| W_{i} B^{\dagger} \right\|_{2}^{2}}{\left\| A_{i::} \right\|_{2}^{2}}. \end{aligned}$$
(14)

By taking the conditional expectation on both sides of (14), we can obtain

$$E_{k} \left[\left\| \hat{X}^{(k+1)} - A^{\dagger} C B^{\dagger} \right\|_{F}^{2} \right] = E_{k} \left[\left\| Y - A^{\dagger} C B^{\dagger} \right\|_{F}^{2} \right] + E_{k} \left[\frac{\left\| W_{i} B^{\dagger} \right\|_{2}^{2}}{\left\| A_{i,:} \right\|_{2}^{2}} \right]$$

$$\leq \left(1 - \frac{\sigma_{\min}^{2}(A) \sigma_{\min}^{2}(B)}{\left\| A \right\|_{F}^{2} \left\| B \right\|_{2}^{2}} \right) \left\| \hat{X}^{(k)} - A^{\dagger} C B^{\dagger} \right\|_{F}^{2} + \frac{\left\| W B^{\dagger} \right\|_{F}^{2}}{\left\| A \right\|_{F}^{2}}$$

The inequality is obtained by Remark 3. Applying this recursive relation iteratively, we have

$$\begin{split} E\left[\left\|\hat{X}^{(k+1)} - A^{\dagger}CB^{\dagger}\right\|_{F}^{2}\right] &\leq \rho E\left[\left\|\hat{X}^{(k)} - A^{\dagger}CB^{\dagger}\right\|_{F}^{2}\right] + \frac{\left\|WB^{\dagger}\right\|_{F}^{2}}{\left\|A\right\|_{F}^{2}} \\ &\leq \rho^{2} E\left[\left\|\hat{X}^{(k-1)} - A^{\dagger}CB^{\dagger}\right\|_{F}^{2}\right] + (\rho+1)\frac{\left\|WB^{\dagger}\right\|_{F}^{2}}{\left\|A\right\|_{F}^{2}} \\ &\leq \cdots \\ &\leq \rho^{k+1}\left\|\hat{X}^{(0)} - A^{\dagger}CB^{\dagger}\right\|_{F}^{2} + (\rho^{k} + \cdots + \rho + 1)\frac{\left\|WB^{\dagger}\right\|_{F}^{2}}{\left\|A\right\|_{F}^{2}} \\ &= \rho^{k+1}\left\|X^{(0)} - A^{\dagger}CB^{\dagger}\right\|_{F}^{2} + \frac{1 - \rho^{k+1}}{1 - \rho}\frac{\left\|WB^{\dagger}\right\|_{F}^{2}}{\left\|A\right\|_{F}^{2}}. \end{split}$$

This completes the proof. \Box

Next, we use the idea of randomized extended Kaczmarz method (see [16–18] for details) for solving the least squares solution of the inconsistent Eq. (1). At each iteration, $Z^{(k)}$ is the kth iterate of

ME-RBK applied to $A^TZB^T = 0$ with the initial guess $Z^{(0)}$, and $X^{(k)}$ is one-step ME-RBK update for $AXB = C - Z^{(k)}$. we can get the following randomized extended block Kaczmarz iteration

$$\begin{cases}
Z^{(k+1)} = Z^{(k)} - \frac{\alpha}{\|A_{:,j}\|_{2}^{2}} A_{:,j} (((A_{:,j}^{T} Z^{(k)}) B^{T}) B), \\
X^{(k+1)} = X^{(k)} + \frac{\alpha}{\|A_{i,:}\|_{2}^{2}} A_{i,:}^{T} ((C_{i,:} - Z_{i,:}^{(k+1)} - (A_{i,:} X^{(k)}) B) B^{T}),
\end{cases} (15)$$

where $\alpha > 0$ is the step size, and i and j are selected with probability $p_i = \frac{\|A_{i,i}\|_2^2}{\|A\|_F^2}$ and $\hat{p}_j(A) = \frac{\|A_{i,j}\|_2^2}{\|A\|_F^2}$, respectively. The cost of each iteration of this method is 4n(q+m)+m+1-2n-q for updating $Z^{(k)}$ and (4q+1)(n+p)+1-2q for updating $X^{(k)}$ if the square of the row norm and the column norm of A have been calculated in advance. We describe this method as Algorithm 2, which is called the ME-REBK algorithm.

Algorithm 2 Randomized Extended Block Kaczmarz Method for AXB = C (ME-REBK)

Input: $A \in R^{m \times p}$, $B \in R^{q \times n}$, $C \in R^{m \times n}$, $X^{(0)} = 0 \in R^{p \times q}$, $Z^{(0)} = C$

- 1: **for** k = 0, 1, 2, ..., do
- 2:
- 3:
- 4:
- Pick j with probability $\hat{p}_{j}(A) = \frac{\|A_{:j}\|_{2}^{2}}{\|A\|_{F}^{2}}$ Compute $Z^{(k+1)} = Z^{(k)} \frac{\alpha}{\|A_{:j}\|_{2}^{2}} A_{:,j}^{-1}((A_{:,j}^{T}Z^{(k)})B^{T})B)$ Pick i with probability $p_{i}(A) = \frac{\|A_{i,j}\|_{2}^{2}}{\|A\|_{F}^{2}}$ Compute $X^{(k+1)} = X^{(k)} + \frac{\alpha}{\|A_{i,j}\|_{2}^{2}} A_{i,j}^{T}((C_{i,j} Z_{i,j}^{(k+1)} (A_{i,j}X^{(k)})B)B^{T})$

Theorem 3. Assume $0 < \alpha < \frac{2}{\|B\|_2^2}$. Let $\{Z^{(k)}\}$ denote the kth iterate of ME-RBK applied to $A^TZB^T = 0$ starting from the initial matrix $Z^{(0)} \in R^{m \times n}$ in which $Z^{(0)}_{:,j} \in C_{:,j} + R(A), j = 1, 2, ..., n$ and $(Z^{(0)}_{i,:})^T \in R^{m \times n}$ $(C_{i,:})^T + R(B^T), i = 1, 2, ..., m$. Then $\{Z^{(k)}\}$ converges linearly to $C - AA^{\dagger}CB^{\dagger}B$ in mean square form, and the solution error in expectation for the iteration sequence $X^{(k)}$ obeys

$$E\left[\left\|Z^{(k)} - (C - AA^{\dagger}CB^{\dagger}B)\right\|_{F}^{2}\right] \le \rho^{k} \left\|Z^{(0)} - (C - AA^{\dagger}CB^{\dagger}B)\right\|_{F}^{2},\tag{16}$$

where the jth column of A is selected with probability $\hat{p}_j(A) = \frac{\|A_{:,j}\|_2^2}{\|A\|_2^2}$.

Proof. The approach of proof is the same as that used in the proof of Theorem 1, so we omit it. \Box

Theorem 4. Assume $0<\alpha<\frac{2}{\|B\|_2^2}$. The sequence $\{X^{(k)}\}$ is generated by the ME-REBK method for AXB=Cstarting from the initial matrix $X^{(0)} \in R^{p \times n}$ and $Z^{(0)} \in R^{m \times n}$, where $X^{(0)}_{:,j} \in R(A^T)$, j = 1, 2, ..., q, $(X^{(0)}_{i,:})^T \in R(B)$, i = 1, 2, ..., p $Z^{(0)}_{:,j} \in C_{:,j} + R(A)$, j = 1, 2, ..., n and $(Z^{(0)}_{i,:})^T \in (C_{i,:})^T + R(B^T)$, i = 1, 2, ..., n

$$E\left[\left\|X^{(k)} - A^{\dagger}CB^{\dagger}\right\|_{F}^{2}\right] \leq \frac{(1+\varepsilon)^{k+1} - (1+\varepsilon)}{\varepsilon^{2}} \frac{\alpha^{2} \|B\|_{2}^{2} \rho^{k}}{\|A\|_{F}^{2}} \left\|Z^{(0)} - (C - AA^{\dagger}CB^{\dagger}B)\right\|_{F}^{2} + (1+\varepsilon)^{k} \rho^{k} \left\|X^{(0)} - A^{\dagger}CB^{\dagger}\right\|_{F}^{2}, \tag{17}$$

where $i \in [m]$, $j \in [p]$ are picked with probability $p_i(A) = \frac{\|A_{i,:}\|_2^2}{\|A\|_E^2}$ and $\hat{p}_j(A) = \frac{\|A_{:,j}\|_2^2}{\|A\|_E^2}$, respectively.

Proof. Let $X^{(k)}$ denote the kth iterate of ME-REBK method for AXB = C, and $\tilde{X}^{(k+1)}$ be the one-step Kaczmarz update for the matrix equation $AXB = AA^{\dagger}CB^{\dagger}B$ from $X^{(k)}$, i.e.,

$$\tilde{X}^{(k+1)} = X^{(k)} + \frac{\alpha}{\|A_{i,:}\|_{2}^{2}} A_{i,:}^{T} (A_{i,:} A^{\dagger} C B^{\dagger} B - A_{i,:} X^{(k)} B) B^{T}.$$

We have

$$\tilde{X}^{(k+1)} - A^{\dagger}CB^{\dagger} = X^{(k)} - A^{\dagger}CB^{\dagger} - \frac{\alpha}{\|A_{i::}\|_{2}^{2}} A_{i::}^{T} A_{i::} (X^{(k)} - A^{\dagger}CB^{\dagger})BB^{T}$$

and

$$X^{(k+1)} - \tilde{X}^{(k+1)} = \frac{\alpha}{\|A_{i,:}\|_2^2} A_{i,:}^T (C_{i,:} - Z_{i,:}^{(k+1)} - A_{i,:} A^{\dagger} C B^{\dagger} B) B^T.$$

For any $\varepsilon > 0$, by triangle inequality and Young's inequality, we can get

$$\begin{split} \left\| X^{(k+1)} - A^{\dagger} C B^{\dagger} \right\|_{F}^{2} &= \left\| (X^{(k+1)} - \tilde{X}^{(k+1)}) + (\tilde{X}^{(k+1)} - A^{\dagger} C B^{\dagger}) \right\|_{F}^{2} \\ &\leq (\left\| X^{(k+1)} - \tilde{X}^{(k+1)} \right\|_{F} + \left\| \tilde{X}^{(k+1)} - A^{\dagger} C B^{\dagger} \right\|_{F}^{2} \\ &\leq \left\| X^{(k+1)} - \tilde{X}^{(k+1)} \right\|_{F}^{2} + \left\| \tilde{X}^{(k+1)} - A^{\dagger} C B^{\dagger} \right\|_{F}^{2} \\ &+ 2 \left\| X^{(k+1)} - \tilde{X}^{(k+1)} \right\|_{F} \left\| \tilde{X}^{(k+1)} - A^{\dagger} C B^{\dagger} \right\|_{F}^{2} \\ &\leq (1 + \frac{1}{\varepsilon}) \left\| X^{(k+1)} - \tilde{X}^{(k+1)} \right\|_{F}^{2} + (1 + \varepsilon) \left\| \tilde{X}^{(k+1)} - A^{\dagger} C B^{\dagger} \right\|_{F}^{2}. \end{split} \tag{18}$$

By taking the conditional expectation on the both sides of (18), we have

$$E_{k}\left[\left\|X^{(k+1)} - A^{\dagger}CB^{\dagger}\right\|_{F}^{2}\right] \leq (1 + \frac{1}{\varepsilon})E_{k}\left[\left\|X^{(k+1)} - \tilde{X}^{(k+1)}\right\|_{F}^{2}\right] + (1 + \varepsilon)E_{k}\left[\left\|\tilde{X}^{(k+1)} - A^{\dagger}CB^{\dagger}\right\|_{F}^{2}\right]. \tag{19}$$

It follows from

$$\begin{split} \left\| X^{(k+1)} - \tilde{X}^{(k+1)} \right\|_{F}^{2} &= \left\| \frac{\alpha}{\|A_{i,:}\|_{2}^{2}} A_{i,:}^{T} (C_{i,:} - Z_{i,:}^{(k+1)} - A_{i,:} A^{\dagger} C B^{\dagger} B) B^{T} \right\|_{F}^{2} \\ &= \frac{\alpha^{2}}{\|A_{i,:}\|_{2}^{2}} \operatorname{trace} \left(B (C_{i,:} - Z_{i,:}^{(k+1)} - A_{i,:} A^{\dagger} C B^{\dagger} B)^{T} (C_{i,:} - Z_{i,:}^{(k+1)} - A_{i,:} A^{\dagger} C B^{\dagger} B) B^{T} \right) \\ &\leq \frac{\alpha^{2} \|B\|_{2}^{2}}{\|A_{i::}\|_{2}^{2}} \left\| C_{i,:} - Z_{i,:}^{(k+1)} - A_{i,:} A^{\dagger} C B^{\dagger} B \right\|_{2}^{2} \end{split}$$

that

$$\begin{split} E_k \left[\left\| X^{(k+1)} - \tilde{X}^{(k+1)} \right\|_F^2 \right] &\leq \alpha^2 \|B\|_2^2 E_k^j E_k^i \left[\frac{\left\| C_{i,:} - Z_{i,:}^{(k+1)} - A_{i,:} A^{\dagger} C B^{\dagger} B \right\|_2^2}{\|A_{i,:}\|_2^2} \right] \\ &= \alpha^2 \|B\|_2^2 E_k^j \left[\frac{1}{\|A\|_F^2} \sum_{i=1}^m \left\| C_{i,:} - Z_{i,:}^{(k+1)} - A_{i,:} A^{\dagger} C B^{\dagger} B \right\|_2^2 \right] \\ &= \frac{\alpha^2 \|B\|_2^2}{\|A\|_F^2} E_k \left[\left\| Z^{(k+1)} - (C - A A^{\dagger} C B^{\dagger} B) \right\|_F^2 \right]. \end{split}$$

It yields

$$E\left[\left\|X^{(k+1)} - \tilde{X}^{(k+1)}\right\|_{F}^{2}\right] \leq \frac{\alpha^{2}\|B\|_{2}^{2}}{\|A\|_{F}^{2}} E\left[\left\|Z^{(k+1)} - (C - AA^{\dagger}CB^{\dagger}B)\right\|_{F}^{2}\right]$$

$$\leq \frac{\alpha^{2}\|B\|_{2}^{2}}{\|A\|_{F}^{2}} \rho^{k+1} \left\|Z^{(0)} - (C - AA^{\dagger}CB^{\dagger}B)\right\|_{F}^{2} \text{ (by Lemma 3)}. \tag{20}$$

By $X^{(0)} - A^{\dagger}CB^{\dagger} \in \mathcal{M}$, we have $X^{(k)} - A^{\dagger}CB^{\dagger} \in \mathcal{M}$. Then, by using Theorem 1, we can get

$$\begin{split} E_{k}[\|\tilde{X}^{(k+1)} - A^{\dagger}CB^{\dagger}\|_{F}^{2}] &= E_{k} \left[\left\| X^{(k)} - A^{\dagger}CB^{\dagger} - \frac{\alpha}{\|A_{i,:}\|_{2}^{2}} A_{i,:}^{T} A_{i,:} (X^{(k)} - A^{\dagger}CB^{\dagger}) BB^{T} \right\|_{F}^{2} \right] \\ &\leq \rho \left\| X^{(k)} - A^{\dagger}CB^{\dagger} \right\|_{F}^{2}, \end{split}$$

then

$$E\left[\left\|\tilde{X}^{(k+1)} - A^{\dagger}CB^{\dagger}\right\|_{F}^{2}\right] \leq \rho E\left[\left\|X^{(k)} - A^{\dagger}CB^{\dagger}\right\|_{F}^{2}\right]. \tag{21}$$

Combining (19),(20) and (21) yields

$$\begin{split} E\left[\left\|X^{(k+1)} - A^{\dagger}CB^{\dagger}\right\|_{F}^{2}\right] &\leq (1 + \frac{1}{\varepsilon})E\left[\left\|X^{(k+1)} - \tilde{X}^{(k+1)}\right\|_{F}^{2}\right] + (1 + \varepsilon)E\left[\left\|\tilde{X}^{(k+1)} - A^{\dagger}CB^{\dagger}\right\|_{F}^{2}\right] \\ &\leq (1 + \frac{1}{\varepsilon})\frac{\alpha^{2}\|B\|_{2}^{2}\rho^{k+1}}{\|A\|_{F}^{2}}\left\|Z^{(0)} - (C - AA^{\dagger}CB^{\dagger}B)\right\|_{F}^{2} \\ &\quad + (1 + \varepsilon)\rho E\left[\left\|X^{(k)} - A^{\dagger}CB^{\dagger}\right\|_{F}^{2}\right] \\ &\leq (1 + \frac{1}{\varepsilon})\frac{\alpha^{2}\|B\|_{2}^{2}\rho^{k+1}}{\|A\|_{F}^{2}}\left\|Z^{(0)} - (C - AA^{\dagger}CB^{\dagger}B)\right\|_{F}^{2}\left[1 + (1 + \varepsilon)\right] \\ &\quad + (1 + \varepsilon)^{2}\rho^{2}E\left[\left\|X^{(k-1)} - A^{\dagger}CB^{\dagger}\right\|_{F}^{2}\right] \\ &\leq \cdots \\ &\leq (1 + \frac{1}{\varepsilon})\frac{\alpha^{2}\|B\|_{2}^{2}\rho^{k+1}}{\|A\|_{F}^{2}}\left\|Z^{(0)} - (C - AA^{\dagger}CB^{\dagger}B)\right\|_{F}^{2}\sum_{i=0}^{k}(1 + \varepsilon)^{i} \\ &\quad + (1 + \varepsilon)^{k+1}\rho^{k+1}\left\|X^{(0)} - A^{\dagger}CB^{\dagger}\right\|_{F}^{2} \\ &= \frac{(1 + \varepsilon)^{k+2} - (1 + \varepsilon)}{\varepsilon^{2}}\frac{\alpha^{2}\|B\|_{2}^{2}\rho^{k+1}}{\|A\|_{F}^{2}}\left\|Z^{(0)} - (C - AA^{\dagger}CB^{\dagger}B)\right\|_{F}^{2} \\ &\quad + (1 + \varepsilon)^{k+1}\rho^{k+1}\left\|X^{(0)} - A^{\dagger}CB^{\dagger}\right\|_{F}^{2}. \end{split}$$

This completes the proof. \Box

Remark 4. Replacing B^T in (15) with B^{\dagger} , we obtain the following projection-based randomized extended block Kaczmarz mathod (ME-PREBK) iteration

$$\begin{cases}
Z^{(k+1)} = Z^{(k)} - \frac{\alpha}{\|A_{:,j}\|_{2}^{2}} A_{:,j} (((A_{:,j}^{T} Z^{(k)}) B^{T}) (B^{\dagger})^{T}), \\
X^{(k+1)} = X^{(k)} + \frac{\alpha}{\|A_{i,:}\|_{2}^{2}} A_{i,:}^{T} ((C_{i,:} - Z_{i,:}^{(k+1)} - (A_{i,:} X^{(k)}) B) B^{\dagger}),
\end{cases} (22)$$

4. Numerical Experiments

In this section, we will present some experiment results of the proposed algorithms for solving various matrix equations, and compare them with ME-RGRK and ME-MWRK in [9] for consistent matrix equations and RBCD in [8] for inconsistent matrix equations. All experiments are carried out by using MATLAB (version R2020a) on a DESKTOP-8CBRR86 with Intel(R) Core(TM) i7-4712MQ CPU @2.30GHz 2.29GHz, RAM 8GB and Windows 10.

All computations are started from the initial guess $X^{(0)}=0$, the step size $\alpha=\frac{1.5}{\|B\|_2^2}$, and terminated once the relative error (RE) of the solution, defined by

$$RE = \frac{\|X^{(k)} - X^*\|_F^2}{\|X^*\|_F^2}$$

at the the current iterate $X^{(k)}$, satisfies $RE < 10^{-6}$ or exceeds maximum iteration K = 50000, where $X^* = A^{\dagger}CB^{\dagger}$. We report the average number of iterations (denoted as "IT") and the average computing time in seconds (denoted as "CPU") for 20 trials repeated runs of the corresponding method. Three examples are tested, and A and B are generated as follows.

- Type I: For given m, p, q, n, the entries of A and B are generated from a standard normal distribution, i.e., A = randn(m, p), B = randn(q, n).
- Type II: Like [14], for given m, p, and $r_1 = rank(A)$, we construct a matrix A by $A = U_1D_1V_1^T$, where $U_1 \in R^{m \times r_1}$ and $V_1 \in R^{p \times r_1}$ are orthogonal columns matrices, $D \in R^{r_1 \times r_1}$ is a diagonal matrix whose first r 2 diagonal entries are uniformly distributed numbers in $[\sigma_{\min}(A), \sigma_{\max}(A)]$, and the last two diagonal entries are $\sigma_{\max}(A), \sigma_{\min}(A)$. The entries of B are generated by similar method with parameters q, n, $r_2 = rank(B)$.
- Type III: The real-world sparse data come from the Florida sparse matrix collection [19].

4.1. Consistent Matrix Equation

Given A, B, we set $C = AX^*B$ with $X^* = randn(p,q)$ to construct a consistent matrix equation. In Tables 1–3, we report the average IT and CPU of the ME-RGRK, ME-MWRK, ME-RBK and ME-PRBK methods for solving consistent matrix. In the following tables, the item '>' represents that the number of iteration steps exceeds the maximum iteration (50000), and the item '-' represents that the method does not converge.

From these tables, we can see that the ME-RBK and ME-PRBK methods vastly outperform the ME-RGRK and ME-MWRK methods in terms of both IT and CPU times regardless of whether the matrices A and B are full column/row rank or not. As the increasing of matrix dimension, the CPU time of ME-RBK and ME-PRBK methods is increasing slowly , while the running time of ME-RGRK and ME-MWRK increases dramatically.

In addition, when the matrix size is small, the ME-PRBK method is competitive, because of the pseudoinverse is less expensive and the number of iteration steps is small. When the matrix size is large, the matrix is large, the ME-RBK method is more challenging because it does not need to calculate the pseudoinverse (see the last line in Table 1).

Table 1. IT and CPU of ME-RGRK, ME-MWRK, ME-RBK and ME-PRBK for the consistent matrix equations with Type I.

m	р	9	n		ME-RGRK	ME-MWRK	ME-RBK	ME-PRBK
100	40	40	100	IT CPU	49707 0.78	27579 2.01	8609.9 0.21	1152.8 0.04
40	100	100	40	IT CPU	> >	49332.7 2.15	6449.2 0.34	1507.2 0.08
500	100	100	500	IT CPU	> >	32109 57.61	4049.4 0.53	1866.1 0.19
1000	200	300	2000	IT CPU	> >	> >	39941.2 7.01	4450.4 0.72

Table 2. IT and CPU of ME-RGRK, ME-MWRK, ME-RBK and ME-PRBK for the consistent matrix equations with Type II.

m	р	r_1	$[\sigma_{\min}(A), \sigma_{\max}(A)]$	q	n	r_2	$[\sigma_{\min}(B), \sigma_{\max}(B)]$		ME-RGRK	ME-MWRK	ME-RBK	ME-PRBK
100	40	20	[1,2]	40	100	40	[1,2]	IT CPU	4361.0 0.06	1987.1 0.17	503.1 0.01	332.4 0.008
100	40	20	[1,5]	40	100	20	[1,5]	IT CPU	22423.2 0.33	6439.8 0.57	9307.6 0.20	1056.3 0.02
1000	200	100	[1,2]	100	1000	50	[1,2]	IT CPU	20055.5 78.45	7047.8 70.56	2587.4 0.42	1674.3 0.23
1000	100	50	[1,5]	200	1000	200	[1,5]	IT CPU	> >	> >	28024.4 4.72	2833.6 0.42

Table 3. IT and CPU of ME-RGRK, ME-MWRK, ME-RBK and ME-PRBK for the consistent matrix equations with Type III.

\overline{A}	В		ME-RGRK	ME-MWRK	ME-RBK	ME-PRBK
divorce	ash 219^T	IT CPU	43927.8 1.15	14164.4 1.35	10993.4 0.36	3873.5 0.13
divorce	ash219	IT CPU	40198.7 0.63	17251.4 0.80	11557.4 0.43	3124.7 0.11
ash219	ash958 ^T	IT CPU	> >	> >	6042.3 1.04	2267.0 0.36
ash219	ash958	IT CPU	> >	> >	5745.4 1.22	2114.2 0.42

4.2. Inonsistent Matrix Equation

To construct an inconsistent matrix equation, we set $C = AX^*B + R$, where X^* and R are random matrices which are generated by $X^* = randn(p,q)$ and $R = \delta * randn(p,q)$, $\delta \in (0,1)$. Numerical results of the RBCD, IME-REBK and IME-PREBK methods are listed in Table 4–6. From these tables, we can see that the IME-PREBK method is better than the RBCD method in terms of IT and CPU time, especially when the $\frac{\sigma_{\max}}{\sigma_{\min}}$ is large (see the last line in Table 5). The IME-REBK method is not competitive for B with full row rank because it needs to solve two equations. However, when B has not full row rank, the RBCD method does not converge, while the IME-REBK and IME-PREBK methods do.

Table 4. IT and CPU of RBCD, IME-REBK and IME-PREBK for the inconsistent matrix equations with Type I.

m	р	q	n		RBCD	IME-REBK	IME-PREBK
100	40	40	100	IT CPU	17212.2 0.69	21270.5 1.57	2469.4 0.19
40	100	100	40	IT CPU	- -	24708.3 2.38	2174.6 0.21
500	100	100	500	IT CPU	6059.1 4.97	7352.8 7.28	2512.3 2.70
1000	200	300	2000	IT CPU	14209.2 152.32	12490.4 246.56	5183.5 99.68

Table 5. IT and CPU of RBCD, IME-REBK and IME-PREBK for the inconsistent matrix equations with Type II.

m	р	r_1	$[\sigma_{\min}(A), \sigma_{\max}(A)]$	q	п	r_2	$[\sigma_{\min}(B), \sigma_{\max}(B)]$		RBCD	IME-REBK	IME-PREBK
100	40	20	[1,2]	40	100	40	[1,2]	IT CPU	1035.9 0.04	762.2 0.05	384.5 0.03
100	40	20	[1,5]	40	100	20	[1,5]	IT CPU	-	16224.7 1.23	1507.8 0.12
1000	200	100	[1,2]	100	1000	50	[1,2]	IT CPU	-	4067.5 24.23	2217.8 13.67
1000	200	100	[1,5]	100	1000	100	[1,5]	IT CPU	> >	48269.6 365.78	4328.0 30.25

Table 6. IT and CPU of RBCD, IME-REBK and IME-PREBK for the inconsistent matrix equations with Type III.

A	В		RBCD	IME-REBK	IME-PREBK
divorce	ash 219^T	IT CPU	> >	20026.3 2.26	4308.4 0.48
divorce	ash219	IT CPU	- -	19199.1 1.49	4026.2 0.31
ash219	ash958 ^T	IT CPU	22313.4 15.69	10823.5 19.03	2561.8 5.89
ash219	ash958	IT CPU	-	10020.7 15.83	2363.5 4.72

5. Conclusions

In this paper we have proposed a randomized block Kaczmarz algorithm for solving the consistent matrix equation and its extended version for the inconsistent case. Theoretically, we have proved the proposed algorithms converge linearly to the unique minimal F-norm solution or least-squares solution (i.e., $A^{\dagger}CB^{\dagger}$) without requirements on A and B have full column/row rank. Numerical results show the effectiveness of the algorithms. Since the proposed algorithms only require one row or one column of A at each iteration without matrix-matrix product, it is suitable for the scenarios where the matrix A is too large to fit in memory or matrix multiplication is considerably expensive.

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