

Article

Science vs. Conspiracy-Theory about Covid-19: Need for Cognition and Openness to Experience Increase Belief in Conspiracy-Theoretical Postings in Social Media

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Abstract

In the context of Covid-19 virus containment, there is a lack of acceptance of preventive measures in the population. The present work investigated which factors influence the belief in of scientific propositions compared belief in conspiracy theories. The focus here was on the determinants of conspiracy beliefs in the context of Covid-19 related media content. Using an online questionnaire ($N = 175$), results indicate that scientific compared to conspiracy theoretical media content led to higher acceptance. Furthermore, Need for Cognition (NFC-K), a conspiracy theoretical worldview (CMQ), and openness to experience (NEO-FFI) were positively associated with conspiracy beliefs derived from Facebook postings. In addition, a conspiracy theoretical worldview was negatively associated with belief in scientific media content. Furthermore, agreeableness was unrelated to conspiracy beliefs, although it was positively associated with conspiracy theoretical worldview. The results imply promising persuasion strategies for reducing conspiracy theoretical beliefs and to increase the acceptance of preventive measures.

Keywords: Covid-19; conspiracy theories; Need for Cognition; agreeableness; openness to experience; social media

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1 Introduction

As a socially stressful event, the Covid-19 (Coronavirus Disease 2019) pandemic triggers feelings of powerlessness and being overwhelmed, as well as stress (Constantinou et al., 2021; Georgiou et al., 2020). According to Swami and colleagues (2014), this leads to a preference for quick and easy explanations for uncertainties. These can be found in conspiracy theoretical media content. On the one hand, conspiracy theories grow more frequently in pandemic contexts (Douglas et al., 2017). On the other hand, crisis situations such as pandemics increase the recourse to conspiracy theoretical explanations (Goertzel, 1994). This development constitutes according to the World Economic Forum 2013 a major threat to society (Howell, 2013).

Conspiracy beliefs with respect to the origin and maintenance of Covid-19 impair scientifically based social action more (Lewandowsky et al., 2013). This is reflected, for example, in reduced vaccination acceptance or reduced adoption of evidence-based prevention and treatment approaches in general (Bogart et al., 2010; Kata, 2010) as well as in relation to Covid-19 (Constantinou et al., 2021; Corbu et al., 2021; Desta & Mulugeta, 2020; Jovančević & Milićević, 2020; Romer & Jamieson, 2020). To counteract this problem, the development of an empirically confirmed framework of the determinants of belief in conspiracy theories is desirable. To facilitate the interpretation of the results the special emphasis in this study is on the examination of determinants of the belief in conspiracy theories contrasted with the determinants of belief in scientifically grounded media content.

In the digital age, social media and its reach are of great importance. Increased exposure to Covid-19-related stress-inducing media content presumably creates a downward spiral in the emergence of conspiracy theoretical beliefs (cf. Corbu et al., 2021; Quattrociocchi et al., 2011). Social media is furthermore known for

using algorithms that control consumption patterns of media content. So-called echo chambers lead to movement within homogeneous information bubbles on social media (Del Vicario et al., 2016). This is most likely to instigate and perpetuate conspiracy beliefs. High exposure to conspiracy theoretical content contributes to its dissemination and acquirement (cf., Bessi et al., 2014; Bessi et al., 2015; Mocanu et al., 2015).

Currently more than four billion people, and thus more than half of the world's population, are using social media (Kemp, 2020; Newman, 2020). In addition, the importance of social media as a source of news and information has increased, according to recent statistics, especially due to the Covid-19 pandemic (Newman, 2020). Both science-based media news and fake news spread quickly and easily on social media (Quattrociocchi et al., 2011). Facebook seems to be the main channel for spreading false information related to conspiracy theories (Newman, 2020).

2 Theoretical Background

2.1 Affirmation of meaning frameworks and the big-cause effect

People are meaning-makers who capture the external world by their mental representations. If their mental representation of the external world is disrupted by contradictory evidence, people experience a deep concern with the unexpected incongruity which activates their meaning maintenance system motivating them to employ alternative frameworks of meaning (Heine et al., 2006).

The emergence of the Covid-19 pandemic threatens the existing framework of meaning because it contradicts the predictability of events and causes a loss of control, which enhances the responsiveness to conspiracy theories (Van Prooijen, 2012). The pandemic represents a global change which has overwhelming consequences for many people and disrupts their belief system. In correspondence with the big-cause effect a readiness to explain big events by big causes is triggered. For example, the occurrence of devastating damage

triggers the tendency to apply a big cause more than the occurrence of less devastating damage (Van Prooijen, 2012). Therefore, the Covid-19 pandemic is likely to be explained by conspiracy theories in accordance with the meaning maintenance system and the big-cause effect. In summary, loss of control, feelings of uncertainty because of disruption of belief systems, and threat to the social order which accompany the emergence of the Covid-19 pandemic facilitate the employment of conspiracy theories which offer an alternative meaning after threatening the predominant meaning maintenance system. The occurrence of the big-cause effect does the rest. As a consequence, belief in conspiracy theories is likely to be intensified.

2.2 Beliefs and attitudes as tools of the meaning-maker

Meanings are carried by beliefs and attitudes. Much research focuses on attitude formation and change. On the basis of previous definitions (e.g., Allport, 1929), Eagly and Chaiken (1993, p. 1) define the concept of attitude as a psychological tendency that is expressed by evaluating a particular entity with some degree of favor or disfavor (see also Eagly & Chaiken, 2008). Examples of these entities or attitude objects are ideas and ideologies. The opinions on ideas and ideologies is reflected in the strength of attitudes and beliefs about them (Brotherton et al., 2013; Bruder et al., 2013; Goertzel, 1994; Swami et al., 2010; 2011; 2012; 2013; 2014). Attitudes and the beliefs on which they are built up serve as frameworks of personal meaning. Therefore, attitudes and beliefs towards Covid-19 related media content serve as building blocks for the attribution of meaning to this media content.

The concept of persuasion describes how attitude change occurs through social influence (cf., Hovland et al., 1953). It is defined as attitude change in response to verbal messages (Kruglanski & Stroebe, 2005, p. 344). Early research on persuasion was reported by Hovland and his co-workers at Yale University. Petty and Cacioppo (1986) in their Elaboration Likelihood Model as well as Eagly and Chaiken (1993) in their Heuristic-Systematic Model proposed influential dual process models of persuasion which were widely applied in recent

research on attitude change. Both models contrast a fast route of attitude change with a more time-consuming route which is based on the systematic elaboration of arguments.

Both models have much in common, but the Heuristic-Systematic Model emphasizes the use of heuristics more than the Elaboration Likelihood Model. Basically, the ELM distinguishes between two routes of persuasion processes depending on the likelihood of elaboration. A distinction is made between the central route, which requires systematic information processing, and the peripheral route, which is based on heuristics and the availability of superficial cues. On the central route, a critical weighing of argumentative content takes place by investing high cognitive effort. This is referred to as high elaboration likelihood. The peripheral route stands for the processing of heuristic cues as cognitive shortcuts neglecting the quality of arguments and requiring little cognitive effort.

2.3 Attitude formation in social media

2.3.1 Scientific theory vs. conspiracy theory

In general, scientific and conspiracy theoretical media content can be distinguished, which oppose each other on a continuum of verifiability (Del Vicario et al., 2016). Scientific media content is mostly characterized by arguments based on empirical sources (Bessi et al., 2015). Conspiracy theoretical media content, on the contrary, presents explanations lacking an empirical basis. Moreover, conspiracy theories use often unsubstantiated assumptions as explanations of public events and societal actions (Douglas et al., 2019). Recent studies show that conspiracy beliefs are increasingly spreading in the context of Covid-19 on social media (Allington et al., 2020; Georgiou et al., 2020; Quinn et al., 2020). However, it is not clear which factors influence belief toward scientific versus conspiracy theoretical Covid-19 related media content.

Scientific compared to conspiracy theoretical media content should foster stronger belief in scientific arguments. This prediction is derived from the activation of the central route of the ELM: Empirically

grounded argument structures of scientific media content are likely to increase the likelihood of elaboration and, thus, influence belief in scientific arguments positively. In contrast conspiracy theoretical content is likely to elicit less systematic elaboration of arguments. In correspondence with this assumption conspiracy media content are related to emotional rather than analytical processing (Douglas et al., 2017). From this reasoning H1 was derived with respect to Facebook postings.

H1: Belief in the scientific posting is higher than belief in the conspiracy theoretical posting.

2.3.2 Need for Cognition

According to the ELM, the probability of elaboration depends on the processing motivation and ability, respectively (Petty & Cacioppo, 1986). The higher the processing motivation/ability, the more likely systematic elaboration of arguments occurs. Processing motivation/ability can be measured by an individual's need for cognition (NFC) which represents individual differences in how much and how readily the person thinks about the arguments contained in a message (Cacioppo et al., 1996). According to the theoretical framework of NFC, high NFC should enhance the persuasive impact of high-quality arguments. From the theoretical framework of NFC H2 was derived.

H2: High NFC increases the belief in scientific theories about Covid-19 related postings, whereas it decreases the belief in corresponding conspiracy theories.

2.3.3 Conspiracy theoretical worldview

Studies show that the belief in conspiracy theories is based on increased exposure to such media content (Bessi et al., 2014; Bessi et al., 2015; Mocanu et al., 2015). Moreover, social media trigger the spread of conspiracy theoretical misinformation more than face-to-face communication (cf. Quattrociocchi et al., 2011). Research indicates that a conspiracy theoretical worldview underlies the beliefs of conspiracists (Brotherton et al., 2013; Bruder et al., 2013; Goertzel, 1994; Swami et al., 2010, 2011, 2013). This assumption was also confirmed in the context of Covid-19 related CTs (Allington et al., 2020; Georgiou et al., 2020).

However, the meaning of social media and the comparison to scientific news in this context remains unclear. According to the review by Wang et al. (2019), social media particularly contributes to the spread of health-related conspiracy theoretical misinformation. This was shown by a negative association between conspiracy theoretical Covid-19 related media content and the acceptance of government health policies (Allington et al., 2020; Freeman et al., 2020). In accordance with this reasoning, H3 was formulated.

H3: A conspiracy theoretical worldview decreases the belief in scientific Covid-19 related postings, while it increases the belief in conspiracy theoretical postings (Georgiou et al., 2020; Swami et al., 2011, 2013).

2.3.4 Big 5 variables

Much personality research is guided by the five-factor (Big 5) model of personality by McCrae and Costa (1987), which in general has been shown to be robust and valid (see also Goldberg, 1990). Two of the five factors – agreeableness and openness to experience – have been connected with conspiracy beliefs (Goreis & Voracek 2019). Previous research revealed inconsistent findings on the relationship between openness to

experience and agreeableness on the one hand and conspiracy beliefs on the other hand (Bruder et al., 2013; Hollander, 2017; Kim et al., 2013; Orosz et al., 2016; Swami et al., 2010; 2011, 2013).

A recent meta-analysis by Goreis and Voracek (2019) indicated that the average correlations between agreeableness and openness to experience on the one hand and conspiracy beliefs on the other hand were close to zero. Because the results were heterogeneous across samples moderator analyses were conducted indicating with respect to openness to experience that samples with larger proportions of men and samples consisting of younger participants exhibited higher correlations of openness to experience with conspiracy beliefs.

In addition, the association between agreeableness and conspiracy beliefs was higher among samples which included a larger proportion of older participants.

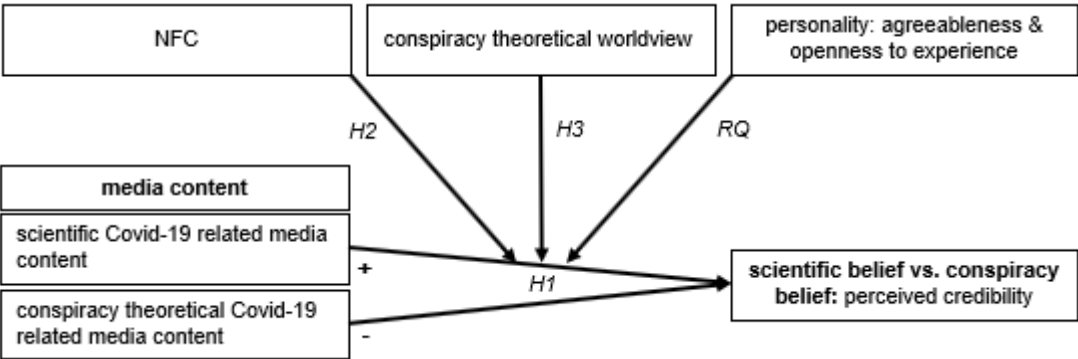
In summary, the results on the association between openness to experience/ agreeableness and belief in conspiracy theories are inconsistent. Therefore, the following research questions were formulated.

RQs: How are the personality traits (a) agreeableness and (b) openness to experience associated with belief in scientific or conspiracy theories embedded in Covid-19 related media content?

The structure of hypotheses and research questions is illustrated in *Figure 1*.

Figure 1

Structure of hypotheses and research questions



3 Method

3.1 Study design

The survey consisted of an online questionnaire.³ Data collection spanned a 5-week period. Participants took a maximum of 15 minutes to complete the survey. The central variables to be collected in this study were the belief in scientific postings and in conspiracy postings. In addition, NFC, conspiracy theoretical worldview, and the Big 5 personality traits agreeableness and openness to experience were measured.

3.2 Materials

3.2.1 Preliminary study of media content

The aim of the preliminary study was to the selection of the media content which was operationalized by two fictitious Facebook postings. The focus is on Facebook postings because Facebook is considered a key medium in spreading fake news (Silverman, 2016). The content of the postings dealt with the effectiveness and mandatory use of FFP masks (Press and Information Office of the Federal Government, 2021). The FFP mask requirement represents an example of a Covid-19-related topic that elicits conflicting opinions in society.

In the preliminary study, the participants read a total of six fictitious Covid-19-related postings which differed in their argument quality and source credibility. A within-subjects design was employed including six postings. Three fictitious scientific (see APPENDIX A1⁴ to A3) and conspiracy theoretical (see APPENDIX A4 to A6) Covid-19-related postings were included. The construction of the postings assumed that high argument quality and high source credibility are characteristics of a scientific posting (Del Vicario et al., 2016; Douglas et al., 2017).

³ The study and its aims were preregistered but due to the blind review the link can only be provided after the review process is completed.

In contrast, these same variables are assumed to be low for conspiracy theoretical postings. Participants rated the six postings on two items, source credibility of the postings on a scale from 1 (*not at all credible*) to 5 (*very credible*) and argument quality of the postings from 1 (*very weak*) to 5 (*very strong*).

The sample of the preliminary study included predominantly young adults with a mean age of 29 years ($SD = 13$ years). The 71% of the 90 participants were female. The highest level of education attained by most participants was Abitur (German high-school diploma, (49%) or university degree (40%). In addition, most of the participants were currently studying (57%) or working (34%).

Table B1 (see APPENDIX B) summarizes the mean ratings of credibility and argument quality of the six postings. For the main study scientific media content was represented by Posting Team A (see APPENDIX A1) exhibiting both the highest mean argument quality ($M = 4.20$, $SD = 0.82$) and source credibility ($M = 4.10$, $SD = 0.79$). Due to a lack of normal distribution by a Shapiro-Wilk test ($W_s \geq .74$, $p < .001$), we used the nonparametric Wilcoxon test to compare means. Regarding both assessments, Posting Team A differed significantly from both scientific Posting Team D and Posting Team E (see APPENDIX B2). Posting Team B represented the conspiracy theoretical posting in the main study (see APPENDIX A4). Posting Team B achieved the lowest source credibility ($M = 1.63$, $SD = 0.73$), which was significantly different from Posting Team C ($z = -2.22$, $p = .026$, $r = .44$) and Posting Team F ($z = -2.15$, $p = .031$, $r = .29$). In addition, posting Team B's argument quality ($M = 1.79$, $SD = 0.73$) was on average assessed low. In summary, based on the results of the preliminary study, Posting Team A served as the scientific media content and Posting Team B as the conspiracy theoretical media content in the main study.

3.2.2 Belief in conspiracy theories

Based on previous research (Brotherton et al., 2013; Bruder et al., 2013; Goertzel, 1994; Swami et al., 2010; 2011; 2013; 2014) belief in conspiracy theories and scientific theories, respectively, was measured by

four items (*How credible/ believable/ plausible/ convincing do you find the posting?*). Responses were obtained on a 7-point Likert scale (1 = *strongly disagree* to 7 = *strongly agree*). The dependent variable was the sum across the four items for CT and ST, respectively. Based on the current sample, the internal consistency of both scales was very high ($\alpha_{CT} = .95$ and $\alpha_{SCI} = .96$).

Additionally, we conducted a confirmatory factor analysis of the ratings of CT and ST using MPlus 8.6 (Muthén & Muthén, 2017). We used the mean and variance adjusted unweighted least squares method was used (ULSMV), that has to be shown as default estimator for models containing ordinal outcomes. Above, the ULSMV was proven to be robust regarding model violations (Kline, 2005). The model fit was assessed using four statistics: (a) the chi-square test statistic, (b) the Comparative Fit Index (CFI; an acceptable fit is inferred if the CFI is 0.90 or higher), (c) the Tucker–Lewis index (TLI; an acceptable fit is inferred if the TLI is 0.90 or higher) and (d) the Root Mean Square Error of Approximation (RMSEA; an acceptable fit is inferred if the RMSEA is equal to 0.08 or smaller). The CFA revealed a good fit: chi-square, $p < .05$, CFI = .996, TLI = .995, RMSEA = .058. Therefore, the CT items were located on the first factor and the ST items on the second factor.

3.2.3 Need for Cognition

Need for Cognition was measured by the questionnaire originally developed by Cacioppo and Petty (1982). The German NFC short scale by Beißert et al., 2014) consists of only four items (e.g., *I would prefer complex to simple problems.*). The response scale of the *NFC-short scale* ranges from 1 (*strongly disagree*) to 7 (*strongly agree*). Beißert and colleagues (2014) stated that the calculation of internal consistencies as a reliability measure for the NFC-short scale is not appropriate because of its shortness. In correspondence with the findings by Beißert et al. (2014), the internal consistency of the short scale in the current sample was low

($\alpha = .42$). An alternative is the assessment of the retest reliability of the scale. Using the test-retest method, Beißert and colleagues (2014) obtained sufficiently high reliability coefficient of the short scale ($r_{tt} = .78$).

3.2.4 Conspiracy theoretical worldview

The conspiracy theoretical worldview was assessed using the German version of the *Conspiracy Mentality Questionnaire* (CMQ, Bruder et al., 2013). The CMQ asks participants to assess their tendency toward a conspiracy theoretical worldview on five items (e.g., *I think government agencies closely monitor all citizens.*) employing an 11-point response scale (*0% certainly not, 10% extremely unlikely, ..., 100% certainly*). Bruder et al. consider the CMQ to be reliable ($\alpha = .82$). The reliability estimate of the CMQ was also high in the current sample ($\alpha = .88$).

3.2.5 Agreeableness

The Big 5 personality trait agreeableness was measured by six items of the respective subscale of the 30-item short version of the German *NEO Five-Factor Inventory* (NEO-FFI-30; Körner et al., 2008). The positive pole of agreeableness refers to trust, cooperation and politeness. The negative pole of antagonism includes facets of arrogance, aggressiveness and manipulateness (McCrae & Costa, 1987). A sample item is *Some people think I am cold and calculating.* The response scale ranges from 1 (*strongly disagree*) to 5 (*strongly agree*). The agreeableness scale revealed sufficient internal consistency in the current sample ($\alpha = .69$). Comparable results were reported by Körner et al. ($\alpha = .75$).

3.2.6 Openness to experience

Openness to experience includes a receptiveness for new ideas, values, and feelings, originality, imaginativeness, and broad interests (McCrae & Costa, 1987). It was measured by 6 items of the respective subscale of the 30-item short version of the German *NEO Five-Factor Inventory* (NEO-FFI-30; Körner et al., 2008). A sample item is *Poetry impresses me little or not at all.* The response scale ranges from 1 (*strongly*

disagree) to 5 (*strongly agree*). The scale revealed sufficient reliability both in the current sample ($\alpha = .72$) and in the calibration sample ($\alpha = .67$).

3.2.7 Demographic variables

We included the following demographic variables: Gender (female, male, diverse), age (in years), highest educational status (no graduation, German “Hauptschulabschluss”, German “Mittlere Reife”, German “Fachabitur”, German “Abitur”, completed apprenticeship, completed university studies, other), current occupation (homemaker, in apprenticeship, studying, in employment, retired, unemployed, other), and marital status (single, in a relationship, married, separated, divorced, widowed) were included in the questionnaire. Furthermore, we applied quality controls for data collection (e.g., to check data quality, please indicate the answer option “strongly disagree” here).

3.3 Sample

Using the snowball sampling technique, we distributed the invitation link for the online questionnaire via social media (WhatsApp, Telegram, Facebook) including information flyers about the study. We explicitly recruited participants in Telegram groups of the German movement “Querdenken” (i.e. a German protest movement including pandemic skeptics, anti-vaxxers, and anti-lockdown protesters). This movement opposes anti-Corona measures of the German government and is open-minded about conspiracy theories. Participant recruitment took place via a study portal of the University XXX. Overall, the sample was quite young representing a broad demographic background of participants.

According to a power analysis by G*Power 3.1 (Faul et al., 2009), the optimal sample size for an assumed medium effect ($F^2 = .15$) was 89 participants. In total, we recruited 274 participants. Data set cleaning, led to the exclusion 95 participants because of minority status refused consent to participate, failed quality

control, and incomplete data. In addition, four additional participants were eliminated from the sample because of extreme scores.

After data set cleaning the final sample consisted of 175 participants, 118 participants of whom were female, 56 participants male, and 1 participant diverse. The mean age was 29 years ($SD = 12$ years). Half of the participants had a high school diploma (German “Abitur”, 54%) as their highest educational qualification, and one third had a university degree (30%). Most of the participants were students (59%) or employees (34%). 41% of the participants were in a romantic relationship without marriage, 38% were single, 19% were married, and $\leq 1\%$ were divorced or separated.

3.4 Statistical analysis

In the first step descriptive analyses were performed. Since the independent variable in this work was not based on a normal distribution according to Kolmogorov-Smirnov-Test and Shapiro-Wilk-Test, we used robust methods for the main hypothesis-based analyses. The computations of the model assumptions can be understood in the provided datasets and syntax. For testing H1, employed nonparametric Wilcoxon tests which tests differences between two dependent samples as nonparametric equivalent to the t-test for dependent samples were employed to compare the belief in perceived scientific and conspiracy theoretical media content. In addition, linear regression analysis to examine the further hypotheses and research questions. All models were controlled for gender, age, highest educational status, current occupation, and marital status which were included as covariates. Finally, tested significance of differences between correlations in magnitude were examined by Fisher's z-test. All statistical analysis were performed using IBM SPSS 24 statistical software (IBM Corp, 2020)⁵

⁵ The data set is available at XXX.

4 Results

4.1 Descriptive analysis

The variables NFC, conspiracy theoretical worldview, agreeableness, and openness to experience were approximately normally distributed which was shown in the visualization of the distribution and skewness ($vs \leq |0.45|$). Further descriptive statistics are summarized in *table 1*. Correlational results indicated as expected that a significant negative association between belief in scientific theories and belief in conspiracy theories occurred explaining 47.6% of total variance ($r = -.69, p < .001$). Overall, 155 participants (89%) indicated that their belief in the scientific theory was stronger than their belief in the conspiracy theory.

Table 1

Means, standard deviations, und correlations

Variable	M	SD	1	2	3	4	5	6	Note. M = mean, SD = standard deviation. df = 173.* p < .05. ** p < .01.
Criteria:									
1. Belief in scientific theories	5.51	1.37	–						
2. Belief in conspiracy theories	2.55	1.61	-.69**	–					
Moderators:									
3. Need for Cognition	4.65	0.73	-.06	.20**	–				

4. Conspiracy theoretical worldview	5.30	2.10	-.44**	.55**	.05	–	
5. Openness to experience	3.61	0.66	-.03	.16*	.37**	-.00	–
6. Agreeableness	3.94	0.58	.14	-.08	.07	-.22**	.15*

4.2
Research
questions

The

research questions refer to the association between Big 5 variables and belief in scientific and conspiracy postings. Specifically, agreeableness and openness to experience were included. Results indicate that agreeableness was neither related to belief in the scientific posting nor to belief in the conspiracy posting. Agreeableness did neither affect the belief in scientific theories (all $ps \geq .061$) nor the belief in conspiracy theories (all $ps \geq .324$), even after taking into account control variables. But agreeableness and conspiracy theoretical worldview were associated negatively, $r = -.22$, $p = .004$, $R^2 = .048$.

Finally, openness to experience was positively associated with belief in the conspiracy posting, $F(1,173) = 4.27$, $p = .040$, $R^2 = .024$, but did not predict belief in the scientific posting. While the effect for the conspiracy posting remained significant after controlling for gender ($p = .028$) and age ($p = .048$), it was no longer significant after controlling for highest educational status, current occupation, and marital status (all $ps \geq .071$). The effect for the scientific posting remained nonsignificant even after taking into account control variables (all $ps \geq .630$). The correlations of openness to experience with both belief variables did not differ significantly, $t(174) = -1.32$, $p = .095$.

4.3 Hypotheses tests

In correspondence with H1, belief in the scientific posting ($M = 5.51$, $SD = 1.37$) was significantly higher than belief in the conspiracy theoretical posting ($M = 2.55$, $SD = 1.61$), $z = -9.20$, $p < .001$, $r = -.69$. Note that we employed the nonparametric Wilcoxon test to examine the hypothesis because a Shapiro-Wilk

test revealed a violation of the assumption of normal distribution for both variables belief in the conspiracy theoretical and in the scientific posting ($W = 0.85, p < .001$).

Linear regression analysis including bootstrapping was employed to examine H2 and H3. *Table 2* summarizes the prediction of belief in the scientific posting by NFC, a conspiracy theoretical worldview, agreeableness, and openness to experience. The results of the respective regression analysis with belief in conspiracy theories as outcome variable are summarized in *Table 3*.

Table 2

Regression analysis: Belief in scientific theories as criterion

Model	β	SE	BCa 95% CI ^a		<i>t</i>	<i>p</i>	<i>R</i> ²
			LL	UL			
1 (Intercept)		0.67	4.80	7.19	8.98	< .001	
Need for Cognition	-0.06	0.14	-0.38	0.17	-0.78	.436	0.004
2 (Intercept)		0.25	6.50	7.56	27.70	< .001	
Conspiracy theoretical worldview	-0.44	0.05	-0.39	-0.18	-6.49	< .001	0.196
3 (Intercept)		0.71	2.93	5.42	5.92	< .001	
Agreeableness	0.14	0.18	0.004	0.67	1.89	.061	0.020
4 (Intercept)		0.58	4.68	6.74	9.85	< .001	
Openness to experience	-0.03	0.16	-0.36	0.25	-0.35	.729	0.001

Note. CI = confidence interval; LL = lower limit; UL = upper limit. *df* = 174.

^a Confidence interval and standard deviation via BCa-Bootstrapping with 10,000 BCa-samples.

Table 3

Regression analysis: perceived credibility of the conspiracy theoretical posting as criterion

Model	β	SE	BCa 95% CI ^a		t	p	R ²
			LL	UL			
1	(Intercept)	0.77	-1.12	1.99	0.61	.543	
	Need for Cognition	0.20	0.12	0.80	2.73	.007	0.041
2	(Intercept)	0.28	-0.18	0.88	1.18	.241	
	Conspiracy theoretical worldview	0.55	0.30	0.53	8.62	< .001	0.300
3	(Intercept)	0.84	1.85	4.91	4.04	< .001	
	Agreeableness	-0.08	-0.58	0.17	-0.99	.324	0.006
4	(Intercept)	0.67	-0.13	2.51	1.77	.078	
	Openness to experience	0.16	0.01	0.75	2.07	.040	0.024

Note. CI = confidence interval; LL = lower limit; UL = upper limit. *df* = 174.

^a Confidence interval and standard deviation via BCa-Bootstrapping with 10,000 BCa-samples.

H2 which focused on Need for Cognition was not confirmed because NFC did not predict the belief in the scientific posting. Even after controlling for the effects of age, gender, highest educational status, current occupation, and marital status, no significant effect of NFC emerged (all *ps* ≥ .436). In contrast, NFC predicted the belief in the conspiracy posting, $F(1,173) = 7.47, p = .007, R^2 = .041$. The magnitude of the correlations between NFC and belief in the scientific posting ($r = -.06$) and NFC and belief in the conspiracy posting ($r = .20$), respectively, was significantly different, $t(174) = -1.92, p = .028$, indicating the occurrence of appreciable differences in the magnitude of correlations. Thus, the

higher the NFC, the higher the belief in conspiracy whereas NFC and belief in scientific theories were unrelated. This effect remained significant even when the control variables were held constant (all p s $\leq .027$).

In correspondence with H3, conspiracy theoretical worldview predicted negatively belief in the scientific posting, $F(1,173) = 42.10$, $p < .001$, $R^2 = .196$, and positively belief in the conspiracy posting, $F(1,173) = 74.31$, $p < .001$, $R^2 = .300$. Even after statistically removing the effects of control variables, conspiracy theoretical worldview was negatively linked with the scientific posting ($p < .001$). When the control variables were included in the regression model, the influence of conspiracy theoretical worldview on belief in conspiracy theory remained significant ($p < .001$). The magnitude of the correlations between a conspiracy theoretical worldview on the one hand and the belief in scientific postings and the belief in conspiracy postings, respectively, on the other hand differed significantly, $t(174) = -8.49$, $p < .001$.

In summary, the stronger the conspiracy theoretical worldview, the weaker the belief in scientific theories and especially the stronger the belief in conspiracy theories.

5 Discussion

The main purpose of the study was to examine determinants of the belief in conspiracy theoretical media content contrasted with the belief in scientifically grounded media content. The focus was on belief in conspiracy postings with respect to Covid-19 related media content.

A review of prior research suggests that three basic motives motivate belief in conspiracy theories: the epistemic motive, the existential motive, and the social motive (Douglas et al., 2017). In this research the focus is on the epistemic motive which explains conspiracy beliefs by feelings of uncertainty which are likely to be aroused by the Covid-19 crisis (Constantinou et al., 2021; Douglas et al., 2017; Georgiou et al., 2020) and which contrast with a desire for certainty. The epistemic motive represents the process of sense making (Heine et al., 2006). Conspiracy beliefs fulfill the sense-making function because they answer questions like ‘who is responsible?’ and ‘which factors cause the threat to the social order?’.

The results confirmed H1 which proposes that belief in the scientific posting is higher than belief in the conspiracy theoretical posting. This hypothesis was derived from higher argument quality of the scientific posting in comparison with the conspiracy posting. The preliminary study confirmed that the argument quality of the scientific posting was higher than the argument quality of the conspiracy posting.

Contrary to H2, NFC enhanced the belief in conspiracy theoretical media content. Furthermore, NFC did not affect the belief in the scientific posting. On the surface, these results seem to contradict the reasoning

by Petty and Cacioppo (1986) which seems to imply that high NFC enhances the critical weighing of arguments on the cognitive level. The results of the preliminary study indicate argument quality was higher for the scientific compared to the conspiracy theoretical posting. Other studies suggest in accordance with the viewpoint by Petty and Cacioppo (1986) a positive relationship between conspiracy theoretical views and low cognitive reflection (Alper et al., 2020) and high intuitive instead of analytical reasoning (Swami et al., 2014).

It is interesting to compare the pattern of results for NFC and belief in media content on the one hand and for openness to experience and media content on the other hand. Remarkably, NFC and openness to experience both foster the belief in the conspiracy theoretical posting (cf., *table 1*). Furthermore, NFC and openness to experience correlate significantly positively, $r = .37$, $p < .01$, indicating 14% of common variance between both constructs. Note that previous research by Berzonsky and Sullivan (1992) has already found this result.

Openness to experience refers to curiosity which includes an interest in understanding new perspectives. Conspiracy theories tend to focus on such new perspectives. High openness to experience is likely to facilitate the interest in unusual and unique ideas. Conspiracy theories incorporate frequently such unusual and unique ideas. Therefore, the positive association between openness to experience and belief in conspiracy postings corresponds with the focus on new perspectives inherent in openness to experience. Nevertheless, the emergence of a reliable positive association between openness to experience and belief in conspiracy contradicts in tendency the meta-analytic results by Goreis and Voracek (2019). The sample characteristics of the sample employed in the current study are only partially correspondent with sample characteristics which are associated with a higher association between openness to experience and conspiracy beliefs. Whereas

predominantly male samples reported more positive associations, the current sample was predominantly female. Furthermore, younger participants reported more positive associations which is in correspondence with the current sample.

Because NFC overlaps considerably with openness to experience the same reasoning might be applied to the positive association between NFC and belief in the conspiracy posting. The content of openness to experience which overlaps with NFC seems to enable people who express high NFC to approach conspiracy postings with positive interest although the component of analytical reasoning seems to favor the belief in the scientific posting. Interestingly, Cacioppo et al. (1996) speculate about two facets of NFC referring to the positive association between openness to experience and NFC. The explanation of the positive association between NFC and belief in the conspiracy posting hinges both on the component of openness to experience which seems to be inherent in NFC and the assumption that the facet of openness to experience implied by NFC has a stronger effect on the evaluation of the conspiracy posting than the emphasis on analytical reasoning implied by NFC which is likely to reduce the belief in the conspiracy posting. Further research is needed to clarify these issues. But an advantage of the proposed explanation for the positive association between NFC and conspiracy belief is that it is able to account for the positive association of both NFC and openness to experience with conspiracy belief and the positive association between NFC and openness to experience.

H3 proposes that a conspiracy theoretical worldview decreases the belief in the scientific Covid-19 related posting, whereas it increases the belief in the conspiracy posting. The results correspond with H3 because they indicate that a conspiracy theoretical worldview increases the belief in the conspiracy theoretical media content whereas it decreases the acceptance of the scientific media content.

In addition, the association between worldview and belief in the conspiracy posting is stronger than the association between worldview and belief in the scientific posting indicating that a similarity effect occurs meaning that the conspiracy theoretical worldview fits into the content of the conspiracy posting. The conspiracy theoretical worldview also inhibits the belief in the scientific posting, but the similarity effect is stronger than the inhibition effect.

Therefore, the conspiracy theoretical worldview strongly facilitates the belief in the conspiracy posting and inhibits somewhat the belief in the scientific posting. This pattern of results indicates that a conspiracy theoretical worldview tends to generalize strongly on a similarity gradient, whereas the inhibition of the belief in the scientific posting is somewhat weaker. In general, these results correspond with prior research on the association between conspiracy theoretical worldview and belief in conspiracy theoretical media content (Allington et al., 2020; Brotherton et al., 2013; Bruder et al., 2013; Georgiou et al., 2020; Goertzel, 1994; Swami et al., 2010, 2011, 2013).

6 Limitations and outlook

In summary, the following factors were positively associated with the belief in the Covid-19 conspiracy theoretical posting: Conspiracy theoretical worldview, NFC, openness to experience, and age. In addition, males scored higher than females.

The present study has some limitations: First, the sample composition did not achieve as much heterogeneity as expected. Despite a large sample size, the representativeness of the sample is limited because younger participants and women are overrepresented. In addition, data were collected by self-report, which may be biased by social desirability and other response sets.

In general, research on conspiracy beliefs relies heavily on self-reports elicited by verbal items. Nevertheless, future studies might employ alternative measurements of belief and attitude. For example, nonreactive measures and methods of indirect measurement (cf., Greenwald & Banaji, 1995) might complement verbal measures.

Finally, conspiracy belief and scientific belief, respectively, were measured by the assessment of single postings. Nevertheless, the reliability of the belief measures was high. To increase generalizability of findings across stimulus materials future studies should include manifold postings.

The dissemination of conspiracy theories constitutes a social problem in the context of the containment of the Covid-19 pandemic. Previous research demonstrated a negative association between conspiracy theoretical beliefs and the acceptance of empirically based preventive measures for pandemic containment (Constantinou et al., 2021; Corbu et al., 2021; Desta & Mulugeta, 2020; Jovančević & Milićević, 2020). In addition, Romer and colleagues (2020) demonstrated a negative association between conspiracy beliefs and vaccination propensity with respect to Covid-19.

This implies the need to limit the influence of conspiracy theoretical postings. Fake news spread faster and broader in social media than the truth (Vosoughi et al., 2018). People are meaning seekers (Heine et al., 2006) who focus on alternative theories if their meaning maintenance system is called into question. Therefore, it is urgent to support the acceptance of science-based news as well as the acceptance of preventive measures on Covid-19. Such an endeavor would profit from including the central route of persuasion proposed by the ELM.

Another way to improve the acceptance of scientific content among conspiracists would be the dissemination of short and understandable explanatory videos. During the Covid-19 pandemic, the use of such videos was already demonstrated in Germany by May Thi Nguyen-Kim, a science journalist who was awarded a Federal Cross of Merit (WDR, 2020). Whether such videos provide convincing arguments instead of conspiracy theoretical arguments could be tested by follow-up studies.

In general, scientific and conspiracy narratives about the causes of Covid-19 compete with each other on social media (Bessi et al., 2015). The present study offers new insights into the message characteristics which are likely to make sure that the scientific narrative surpasses the conspiracy narrative about Covid-19 in social media by identifying factors that determine conspiracy theoretical beliefs and informing about applied implications. The results enable the derivation of measures of social action against the current Covid-19 pandemic. For example, an opinion attack on conspiracy theoretical worldviews seems to be promising. Whereas such an attack is located on the ideological level, additional anti-conspiracy communications which are promising are located on the individual level.

The ELM by Petty and Cacioppo (1986) considers the relevance of argument quality in persuasive communications and distinguishes between two routes of persuasion depending on the likelihood of elaboration. The central route which takes the quality of arguments into account is contrasted with the peripheral route which puts less emphasis on argument quality. The central route requires a careful weighing of argumentative content.

The findings of the present study indicate that both NFC and openness to experience are positively associated with a high probability of elaboration in terms of the ELM. Because both determinants are positively related to systematic cognitive processing and preference for novel viewpoints, in a first step it is necessary to revise the stereotype of conspiracists as superficial people. In a second step, the use of counter-arguments should take into account that many conspiracists are ready to elaborate arguments carefully in accordance with high elaboration likelihood. Although some of these conspiracists may indeed fulfill the stereotype, others focus on arguments in conspiracy messages and their argumentative refutation and are fascinated by the disclosure of new perspectives which are beyond traditional thinking including scientific thinking. Therefore, conspiracists frequently focus on the central route of information processing.

In summary, it is likely that an argumentative discussion, which gives time to the refutation of conspiracy theories in detail and includes counterarguments, is a successful antidote. In addition, the emphasis on new perspectives which are derived from scientific thinking is likely to be especially convincing for people who are fascinated by conspiracy theories.

Declarations: We declare that our study was approved by the local ethical committee of XXX. We declare that there are no potential conflicts of interests, and that informed consent was given by all participants.

Conflicts of Interest: none.

Data availability: The datasets generated and analyzed during the current study are available in the XXX repository. Note that for the blind review process we covered the link.

References

- Alby, T. (2007). *Web 2.0. Konzepte, Anwendungen, Technologien*. (2nd Edition). Hanser.
- Allington, D., Duffy, B., Wessely, S., Dhavan, N. & Rubin, J. (2020). Health-protective behaviour, social media usage and conspiracy belief during the COVID-19 public health emergency. *Psychological Medicine*, 1–7. <https://doi.org/10.1017/S003329172000224X>
- Allport, G. W. (1929). The composition of political attitudes. *American Journal of Sociology*, 35(2), 220-238. <https://doi.org/10.1086/214980>
- Alper, S., Bayrak, F. & Yilmaz, O. (2020). Psychological correlates of COVID-19 conspiracy beliefs and preventive measures: Evidence from Turkey. *Current Psychology*, 1–10. <https://doi.org/10.1007/s12144-020-00903-0>
- Beißert, H., Köhler, M., Rempel, M. & Beierlein, C. (2014). Eine deutschsprachige Kurzsкала zur Messung des Konstrukts Need for Cognition: Die Need for Cognition Kurzsкала (NFC-K). *GESIS-working papers*, 32, 1–28. <https://nbn-resolving.org/urn:nbn:de:0168-ssoar-403157>
- Berzonsky, M.D. & Sullivan, C. (1992). Social-cognitive aspects of identity style: Need for cognition, experiential openness, and introspection. *Journal of Adolescent Research*, 7, 140-155. <https://doi.org/10.1177/074355-489272002>
- Bessi, A., Coletto, M., Davidescu, G. A., Scala, A., Caldarelli, G. & Quattrociocchi, W. (2015). Science vs conspiracy: Collective narratives in the age of misinformation. *PloS one*, 10(2), e0118093. <https://doi.org/10.1371/journal.pone.0118093>

- Bessi, A., Scala, A., Rossi, L., Zhang, Q. & Quattrociocchi, W. (2014). The economy of attention in the age of (mis)information. *Journal of Trust Management*, 1(12). <https://doi.org/10.1186/s40493-014-0012-y>
- Bogart, L. M., Wagner, G., Galvan, F. H. & Banks, D. (2010). Conspiracy beliefs about HIV are related to antiretroviral treatment nonadherence among African American men with HIV. *Journal of Acquired Immune Deficiency Syndromes*, 53(5), 648–655. <https://doi.org/10.1097/QAI.0b013e3181c57dbc>
- Brotherton, R., French, C. C. & Pickering, A. D. (2013). Measuring belief in conspiracy theories: the generic conspiracist beliefs scale. *Frontiers in Psychology*, 4(279), 1–15. <https://doi.org/10.3389/fpsyg.2013.00279>
- Bruder, R., Haffke, P., Neave, N., Nouripanah, N. & Imhoff, R. (2013). Measuring individual differences in generic beliefs in conspiracy theories across cultures: Conspiracy Mentality Questionnaire. *Frontiers in Psychology*, 4(225), 1–15. <https://doi.org/10.3389/fpsyg.2013.00279>
- Bundesamt für Verfassungsschutz. (2021). *Neuer Phänomenbereich "Verfassungsschutzrelevante Delegitimierung des Staates"*. <https://www.verfassungsschutz.de/SharedDocs/kurzmeldungen/DE/2021/2021-04-29-querdenker.html>
- Cacioppo, J. T. & Petty, R. E. (1982). The need for cognition. *Journal of Personality and Social Psychology*, 42, 116–131.
- Cacioppo, J.T., Petty, R.E., Feinstein, J.A. & Jarvis, B.G. (1996). Dispositional differences in cognitive motivation: The life and times of individuals varying in need for cognition. *Psychological Bulletin*, 119, 197-253.

-
- Choi, D.-H. & Shin, D.-H. (2017). Exploring political compromise in the new media environment: The interaction effects of social media use and the Big Five personality traits. *Personality and Individual Differences*, 107, 163–171. <https://doi.org/10.1016/j.paid.2016.11.022>
- Cohen, A., Stotland, E. & Wolfe, D. (1955). An experimental investigation of need for cognition. *Journal of Abnormal and Social Psychology*, 51(2), 291–294. <https://doi.org/10.1037/h0042761>
- Cohen, J. (1988). *Statistical Power Analysis for the Behavioral Sciences* (2nd ed.). Lawrence Erlbaum Associates.
- Constantinou, M., Gloster, A. & Karkla, M. (2021). I won't comply because it is a hoax: Conspiracy beliefs, lockdown compliance, and the importance of psychological flexibility. *Journal of Contextual Behavioral Science*, 20, 46–51. <https://doi.org/10.1016/j.jcbs.2021.03.001>
- Corbu, N., Negrea-Nusuioc, E., Udrea, G. & Radu, L. (2021). Romanians' willingness to comply with restrictive measures during the covid-19 pandemic: Evidence from an online survey. *Journal of Applied Communication Research*. <https://doi.org/10.1080/00909882.2021.1912378>
- Costa, P. T., Jr., & McCrae, R. R. (1985). The NEO Personality Inventory. Odessa, FL: Psychological Assessment Resources.
- Del Vicario, M., Bessi, A., Zollo, F., Petroni, F., Scala, A. & Caldarellia, G. (2016). The spreading of misinformation online. *PNAS*, 113(3), 554–559. <https://doi.org/10.1073/pnas.1517441113>
- Desta, T. T. & Mulugeta, T. (2020). Living with COVID-19-triggered pseudoscience and conspiracies. *International Journal of Public Health*, 65(6), 713–714. <https://doi.org/10.1007/s00038-020-01412-4>

- Douglas, K. M., Sutton, R. M. & Cichocka, A. (2017). The Psychology of Conspiracy Theories. *Current Directions in Psychological Science*, 26(6), 538–542. <https://doi.org/10.1177/0963721417718261>
- Douglas, K. M., Uscinski, J. E., Sutton, R. M., Cichocka, A., Nefes, T., Ang, C. S. & Deravi, F. (2019). Understanding Conspiracy Theories. *Political Psychology*, 40(S1), 3–35. <https://doi.org/10.1111/pops.12568>
- Eagly, A.H. & Chaiken, S. (1993). The psychology of attitudes. Yale University Press.
- Edwards, A. L. (1957). *The social desirability variable in personality assessment and research*. Dryden Press.
- Faul, F., Erdfelder, E., Buchner, A. & Lang, A.-G. (2009). Statistical power analyses using G*Power 3.1: Tests for correlation and regression analyses. *Behavior Research Methods*, 41(4), 1149–1160. <https://doi.org/10.3758/BRM.41.4.1149>
- Freeman, D., Waite, F., Rosebrock, L., Petit, A., Causier, C., East, A., Jenner, L., Teale, A.-L., Carr, L., Mulhall, S., Bold, E. & Lambe, S. (2020). Coronavirus conspiracy beliefs, mistrust, and compliance with government guidelines in England. *Psychological Medicine*, 1–13. <https://doi.org/10.1017/S0033291720001890>
- Furnham, A. (2021). Just world beliefs, personal success and beliefs in conspiracy theories. *Current Psychology*, 1–7. <https://doi.org/10.1007/s12144-021-01576-z>
- Georgiou, N., Delfabbro, P. & Balzan, R. (2020). COVID-19-related conspiracy beliefs and their relationship with perceived stress and pre-existing conspiracy beliefs. *Personality and Individual Differences*, 166, 110201. <https://doi.org/10.1016/j.paid.2020.110201>

- Glaser, S. (2020). Telegram: Admist Propaganda of Violence and "Infowars". *Jugenschutz.net*.https://www.vielfalt-mediathek.de/wp-content/uploads/2020/12/jugenschutz_net_report_telegram_amidst_propaganda_of_violence_and_infowars_vielfalt_mediathek.pdf
- Goertzel, T. (1994). Belief in Conspiracy Theories. *Political Psychology* (15), 731–742. <https://doi.org/10.2307/3791630>
- Goldberg, L. R. (1990). An alternative "description of personality": The Big-Five factor structure, *Journal of Personality and Social Psychology*, 59(6), 1216–1229. <https://doi.org/10.1037/0022-3514.59.6.1216>
- Goreis, A. & Voracek, M. (2019). A Systematic Review and Meta-Analysis of Psychological Research on Conspiracy Beliefs: Field Characteristics, Measurement Instruments, and Associations With Personality Traits. *Frontiers in Psychology*, 10(205), 1–13. <https://doi.org/10.3389/fpsyg.2019.00205>
- Greenberg, J., Solomon, S. & Pyszczynski, T. (1997). Terror Management Theory of Self-Esteem and Cultural Worldviews: Empirical Assessments and Conceptual Refinements. *Advances in Experimental Social Psychology*, 29, 61–139. [https://doi.org/10.1016/S0065-2601\(08\)60016-7](https://doi.org/10.1016/S0065-2601(08)60016-7)
- Greenwald, A.G. & Banaji, M.R. (1995). Implicit social cognition: Attitudes, self-esteem, and stereotypes. *Psychological Review*, 102, 4-27.
- Grzesiak-Feldman, M. (2013). The Effect of High-Anxiety Situations on Conspiracy Thinking. *Current Psychology*, 32(1), 100–118. <https://doi.org/10.1007/s12144-013-9165-6>
- Hollander, B. A. (2017). Partisanship, individual differences, and news media exposure as predictors of conspiracy beliefs. *Journalism & Mass Communication Quarterly* 95, 691–713. <https://doi.org/10.1177/1077699017728919>

- Hovland, C. I., Janis, I. L. & Kelley, H. H. (1953). *Communication and persuasion*. Yale University Press, (135-142). <https://doi.org/10.1007/BF02713272>.
- Howell, L. (Februar 2013). *Digital Wildfires in a hyperconnected world*. *Global Risks Report*. World Economic Forum. <http://reports.weforum.org/global-risks-2013/risk-case-1/digital-wildfires-in-a-hyperconnected-world/>
- IBM Corp. (2020). *IBM SPSS Statistics for Windows* (Version 27.0) [Computer software]. Armonk, NY. <https://www.ibm.com/de-de/analytics/spss-statistics-software>
- Jovančević, A. & Milićević, N. (2020). Optimism-pessimism, conspiracy theories and general trust as factors contributing to COVID-19 related behavior - A cross-cultural study. *Personality and Individual Differences*, 167, 110216. <https://doi.org/10.1016/j.paid.2020.110216>
- Kata, A. (2010). A postmodern Pandora's box: anti-vaccination misinformation on the Internet. *Vaccine*, 28(7), 1709–1716. <https://doi.org/10.1016/j.vaccine.2009.12.022>
- Kemp, S. (2020). *Digital 2020 October Global Statshot Report*. We Are Social Inc. <https://wearesocial.com/blog/2020/10/social-media-users-pass-the-4-billion-mark-as-global-adoption-soars>
- Kim, Y., Hsu, S.-H. & Zúñiga, H. G. de (2013). Influence of Social Media Use on Discussion Network Heterogeneity and Civic Engagement: The Moderating Role of Personality Traits. *Journal of Communication*, 63(3), 498–516. <https://doi.org/10.1111/jcom.12034>
- Körner, A., Geyer, M., Roth, M., Drapeau, M., Schmutzer, G., Albani, C., Schumann, S. & Brähler, E. (2008). Persönlichkeitsdiagnostik mit dem NEO-Fünf-Faktoren-Inventar: Die 30-Item-Kurzversion (NEO-FFI-30) [Personality assessment with the NEO-Five-Factor Inventory: the 30-Item-Short-Version (NEO-

FFI-30)]. *Psychotherapie, Psychosomatik, medizinische Psychologie*, 58(6), 238–245.

<https://doi.org/10.1055/s-2007-986199>

Kossowska, M. & Bukowski, M. (2015). Motivated roots of conspiracies: The role of certainty and control motives in conspiracy thinking. In: Bilewicz, M., Cichocka, A. & Soral, W. (Eds.). *Psychology of Conspiracy*, 1st Edition (145–161). Routledge.

Lamberty, P. & Imhoff, R. (2018). Powerful Pharma and Its Marginalized Alternatives? *Social Psychology*, 49(5), 255–270. <https://doi.org/10.1027/1864-9335/a000347>

Lewandowsky, S., Oberauer, K. & Gignac, G. E. (2013). NASA faked the moon landing--therefore, (climate) science is a hoax: an anatomy of the motivated rejection of science. *Psychological Science*, 24(5), 622–633. <https://doi.org/10.1177/0956797612457686>

Marchlewska, M., Cichocka, A. & Kossowska, M. (2018). Addicted to answers: Need for cognitive closure and the endorsement of conspiracy beliefs. *European Journal of Social Psychology*, 48(2), 109–117. <https://doi.org/10.1002/ejsp.2308>

McCrae, R.R. Costa, P.T. (1987). Validation of the five-factor model of personality across instruments and observers. *Journal of Personality and Social Psychology*, 52, 81-90.

Mocanu, D., Rossi, L., Zhang, Q., Karsai, M. & Quattrociocchi W.. (2015). Collective attention in the age of (mis)information. *Computers in Human Behavior*, 51, 1198–1204. <https://doi.org/10.1016/j.chb.2015.01.024>

Newman, N. (2020). *Reuters Institute Digital News Report 2020*. https://reutersinstitute.politics.ox.ac.uk/sites/default/files/2020-06/DNR_2020_FINAL.pdf

- Orosz, G., Krekő, P., Paskuj, B., Tóth-Király, I., Bothe, B. & Roland-Lévy, C. (2016). Changing Conspiracy Beliefs through Rationality and Ridiculing. *Frontiers in Psychology*, 7(1525), 1–9. <https://doi.org/10.3389/fpsyg.2016.01525>
- Petty, R. E. & Cacioppo, J. T. (1986). The elaboration likelihood model of persuasion. *Advances in Experimental Social Psychology*, 19, 123-205. [https://doi.org/10.1016/S0065-2601\(08\)60214-2](https://doi.org/10.1016/S0065-2601(08)60214-2)
- Press and Information Office of the Federal Government. (2021). *Masken: wie sie schützen, wo sie getragen werden müssen*. <https://www.bundesregierung.de/breg-de/aktuelles/regelung-zu-masken-1842704>
- Quattrociocchi, W., Conte, R. & Lodi, E. (2011). Opinions Manipulation: Media, Power and Gossip. *Advances in Complex Systems*, 14(04), 567–586. <https://doi.org/10.1142/S0219525911003165>
- Quinn, E. K., Fazel, S. S. & Peters, C. E. (2020). The Instagram Infodemic: Cobranding of Conspiracy Theories, Coronavirus Disease 2019 and Authority-Questioning Beliefs. *Cyberpsychology, Behavior and Social Networking*. <https://doi.org/10.1089/cyber.2020.0663>
- Romer, D. & Jamieson, K. H. (2020). Conspiracy theories as barriers to controlling the spread of COVID-19 in the U.S. *Social Science & Medicine*, 263, 113356. <https://doi.org/10.1016/j.socscimed.2020.113356>
- Silverman, C. (2016). *This analysis shows how viral fake election news stories outperformed real news on facebook*. *Buzzfeed*, November 16, 2016. Retrieved May 22, 2017, from https://www.buzzfeed.com/craig-silverman/viral-fake-election-news-outperformed-real-news-on-facebook?utm_term=.lgQvmj974#.qqXqL1AJV.

-
- Swami, V., Chamorro-Premuzic, T. & Furnham, A. (2010). Unanswered questions: A preliminary investigation of personality and individual difference predictors of 9/11 conspiracist beliefs. *Applied Cognitive Psychology*, 24(6), 749–761. <https://doi.org/10.1002/acp.1583>
- Swami, V., Coles, R., Stieger, S., Pietschnig, J., Furnham, A., Rehim, S. & Voracek, M. (2011). Conspiracist ideation in Britain and Austria: evidence of a monological belief system and associations between individual psychological differences and real-world and fictitious conspiracy theories. *British Journal of Psychology*, 102(3), 443–463. <https://doi.org/10.1111/j.2044-8295.2010.02004.x>
- Swami, V. & Furnham, A. (Hrsg.). (2014). Political paranoia and conspiracy theories. In: van Prooijen, J.-W. & van Lange, P.A.M. (Eds.) *Power, Politics, and Paranoia: Why People are Suspicious of their Leaders* (218-236). Cambridge University Press.
- Swami, V., Pietschnig, J., Tran, U. S., Nader, I. W., Stieger, S. & Voracek, M. (2013). Lunar Lies: The Impact of Informational Framing and Individual Differences in Shaping Conspiracist Beliefs About the Moon Landings. *Applied Cognitive Psychology*, 27(1), 71–80. <https://doi.org/10.1002/acp.2873>
- Swami, V., Voracek, M., Stieger, S., Tran, U. S. & Furnham, A. (2014). Analytic thinking reduces belief in conspiracy theories. *Cognition*, 133(3), 572–585. <https://doi.org/10.1016/j.cognition.2014.08.006>
- Tajfel, H. & Turner, J. C. (1985). The Social Identity Theory of Intergroup Behavior. In: Worchel, S. & Austin, W.G. (Eds.), *Psychology of Intergroup Relations*, 2nd Edition, (7-24). Nelson Hall.
- Tversky, A. & Kahnemann, D. (1974). Judgment under Uncertainty: Heuristics and Biases. *Science* (185), 1124–1131. <https://10.1126/science.185.4157.1124>

-
- Vosoughi, S., Roy, D., & Aral, S. (2018). The spread of true and false news online. *Science*, 359 (6380), 1146-1151. 10.1126/science.aap9559.
- Wang, Y., McKee, M., Torbica, A. & Stuckler, D. (2019). Systematic Literature Review on the Spread of Health-related Misinformation on Social Media. *Social Science & Medicine*, 240, 112552. <https://doi.org/10.1016/j.socscimed.2019.112552>
- Wason, P. C. (1968). Reasoning about a rule. *The Quarterly journal of experimental psychology*, 20(3), 273–281. <https://doi.org/10.1080/14640746808400161>
- WDR (2020, October 1). *Bundesverdienstkreuz für Drosten, Pusch und May Thi Nguyen*. <https://www1.wdr.de/nachrichten/themen/coronavirus/bundesverdienstkreuz-corona-krise-drosten-pusch-100.html>
- Webster, D. M. & Kruglanski, A. W. (1994). Individual Differences in Need for Cognitive Closure. *Journal of Personality and Social Psychology*, 67(6), 1049–1062. <https://doi.org/10.1037/0022-3514.67.6.1049>
- Wood, M. J. & Gray, D. (2019). Right-wing authoritarianism as a predictor of pro-establishment versus anti-establishment conspiracy theories. *Personality and individual differences*, 138, 163–166. <https://doi.org/10.1016/j.paid.2018.09.036>