Modeling spatial local variations in intermarriage

Albert Esteve [1], Coro Chasco [2,3]* and Antonio López-Gay [4,5]

[1] Centre d'Estudis Demogràfics, Universitat Autònoma de Barcelona, Carrer de Ca n'Altayó, Edifici E2, 08193 Bellaterra, Barcelona, Spain; aesteve@ced.uab.cat

[2] Department of Applied Economics, Universidad Autónoma de Madrid, C/Francisco Tomás y Valiente, 5, 28049 Madrid, Spain; coro.chasco@uam.es

[3] ECEMIN Research Group, Nebrija University, C/Sta. Cruz de Marcenado, 27, 28015 Madrid, Spain

[4] Department of Geography, Universitat Autònoma de Barcelona Campus de Bellaterra, Edifici B, Carrer de la Fortuna, s/n, 08193 Bellaterra, Barcelona, Spain; tlopez@ced.uab.cat

[5] Centre d'Estudis Demogràfics, Universitat Autònoma de Barcelona, Carrer de Ca n'Altayó, Edifici E2, 08193 Bellaterra, Barcelona, Spain

* Correspondence: coro.chasco@uam.es

Abstract: We utilized all Spanish marriage records available at the municipality level from 2005-2007 to model spatial variations in intermarriage. We constructed a spatial regime zero inflated Poisson model and grouped-data probit model, with spatially lagged regressors, to predict the absolute and relative presence of intermarriage between Spaniards and migrants based on structural characteristics of the local marriage markets and their neighboring areas (i.e., relative group size, homogeneity of national origins, and sex ratio indicators). Our models do not assume collapsibility of the marriage market. Instead, they incorporate the local dimension of the marriage market and examine the association between intermarriage and structural variables at the spatial local level. The model also investigates intermarriage variation by size of place. The local characteristics of the marriage markets are robust indicators of both the absolute and relative importance of intermarriage, but their impact varies by size of municipality. The relative size of the migrant community positively impacts intermarriage. The homogeneity of the origins of migrants is negatively related to it. The impact of sex ratios in the migrant and native communities on intermarriage is not uniform across all municipalities and is not always related to more intermarriage.

Keywords: Intermarriage; migration; local markets; Poisson model; Probit model; spatial autocorrelation; spatial heterogeneity; Spain

1. Introduction

Social scientists have paid little attention to the variation of intermarriage across local marriage markets. Data constraints may explain the lack of spatial awareness in intermarriage research. Datasets seldom provide the geographical detail and sample density required for in-depth spatial analyses. Thus, researchers feel constrained to assume collapsibility of the marriage market (i.e., one country, one marriage market) at the risk of under- or over-estimating the social distance between groups based on national levels of intermarriage ([1]). The collapsibility bias primarily arises from two well-established facts. First, most individuals find their future spouses/partners in the places where they live, work, study, relax and conduct their main activities ([2]). Second, demographic, ethnic and racial compositions vary from place to place ([3-4]). Not all individuals are exposed to the same opportunities.

In this paper, we move beyond the collapsibility assumption to investigate the relationship between intermarriage and space. We develop spatial models to predict the absolute and relative number of intermarried couples as a function of the structural characteristics of local marriage markets. Our models combine several features not commonly found in previous research. First, we neither assume collapsibility of the marriage market nor define standardized geographic areas based on an equal number of inhabitants or any

other equalizing criteria (e.g., same area). Instead, we utilize spatial analysis techniques to incorporate the characteristics of area (i.e., municipalities in Spain) and the corresponding neighboring areas as determinants of intermarriage. Second, we construct a set of structural variables to portray marriage markets at the local scale in terms of the relative size of the migrant community, homogeneity of migrant origins, and sex imbalances in migrant and native groups. Third, we examine intermarriage variations by size of place, from small towns to large cities.

This paper addresses two main questions. a) Is intermarriage predictable from the structural characteristics of local marriage markets? b) How is intermarriage related to relative group size, homogeneity of migrant origin, and sex ratios, and how does this relationship vary by size of place?

Our models are applied to intermarriage between foreign national women (migrant women) and Spanish men and intermarriage between foreign national men (migrant men) and Spanish women. By foreign nationals we refer to men and women that do not have Spanish citizenship. We model the absolute number and relative frequency of intermarried couples in Spain from 2005 to 2007 in all its 8,111 municipalities. For this purpose, we use full counts of marriage micro-data from the Spanish Vital Registration System from 2005-2007. This period is particularly relevant for the study of intermarriage in Spain, since it falls right at the peak of the first major stage of international migration in the country. To predict the absolute number of intermarried couples, we use a zero inflated Poisson model with spatially lagged independent variables. For the relative frequency (intermarriage rate), we use a probit model for grouped data with spatially lagged independent variables. Both models include the same explanatory variables but differ in the way the endogenous variable has been defined.

2. Background and hypotheses

Conceptually, patterns of assortative mating result from the interplay among three forces: individual preferences, third-party influences (i.e., church, state, and family), and the structural context of opportunities ([5-6]). Most individuals in modern society are assumed to have the freedom to choose their partners because third parties have lost most of their historical influence on partner choice ([7-8]). However, for some ethnic groups, this may not be the case (e.g., migrants coming from countries with different religions). In a context of individual choice, assortative mating becomes a matter of individual preferences and structural opportunities. Researchers have developed various alternatives to control for opportunities, but they seldom formulate hypotheses regarding its direct influence on assortative mating (some exceptions are [9-11] or, more recently, [12-13]). This paper examines the effect of structural constraints on intermarriage. We develop a local model of intermarriage that combines three characteristics. First, the model avoids collapsibility of the marriage market by defining marriage markets from a local perspective but considers the influence of neighboring areas. Second, the model includes structural variables constructed from the entire population and not exclusively from the population in union. Third, the model allows for variations on intermarriage by size of place.

Following [1], collapsibility refers to the assumption that marriage markets are collapsible; the odds of intermarriage on the local scale mirror the national scale. Collapsibility implies that the availability of a uniform pool of candidates in one's place of residence and in the distance to one's place of residence (work, study, or leisure) are not relevant dimensions in the marital search process. Neither of these two suppositions applies to the real world. First, the distribution of migrant, ethnic, racial or other groups across marriage markets is far from uniform ([3],[14-15]). Second, people tend to choose spatially homogamous partners ([5],[16-19]). As a result, people are exposed to different opportunities. In this regard, in [1] it is demonstrated that the odds of intermarriage between blacks and whites were underestimated when assuming the collapsibility of the marriage market. Following this idea, our model defines marriage markets at the local scale for both the dependent (i.e., intermarried couples) and the independent variables (i.e., local

characteristics of the marriage market). However, our model also considers the influence of neighboring areas by including spatially lagged independent variables.

A common practice in intermarriage research is to control for the structural constraints of the marriage market. Typically, these constraints are constructed only from the population in union (e.g., [20-21], and more recently, [22]). For example, intermarriage research widely utilizes log-linear models. These models adjust the observed marriage patterns for differences in group size by controlling for the marginal distributions of a contingency table that classifies couples by the husband's and wife's characteristics. Consequently, the characteristics of the marriage market are exclusively inferred from the married population, overlooking the non-married population.

On the contrary, the structural variables used in our model are based on the entire population, similar to [9]. We borrow two indicators from this research: group relative size and heterogeneity origins. Relative size refers to the share of the minority group in each community. In our case, the minority group refers to the migrant population, which is the population born abroad. Heterogeneity refers to the diversity of origins in the migrant community. We expect a positive relationship between the relative size of the migrant community and intermarriage with the native population. Intermarriage in absolute and relative terms will increase as the size of the migrant community increases: more migrants means more opportunities to intermarry. From a migrant's point of view, the opposite may hold: a migrant's propensity to marry out of his or her group may decrease as the size of the migrant community increases. However, our models do not predict a migrant's probability to marry out of his or her group but the occurrence of intermarriage in each municipality.

Regarding heterogeneity in migrant origins, intermarriage will be lower inmigrant communities that are very homogeneous. The classical assumption is that homogenous communities will encourage endogamous marriages, assuming there are no major imbalances in the number of men and women of the same origin.

The demographic literature on the marriage squeeze has shown that societal sex imbalances can promote intermarriage with other groups ([23-27]). First generation migrant communities historically present skewed sex distributions because of male- or female-dominated flows of migration. However, skewed sex ratios are not exclusive in-migrant groups; they are also common in specific areas among natives. In rural areas with female out-migration, males outnumber females ([28-29]). We expect that skewed sex ratios among migrants and natives will promote intermarriage between them.

Finally, a less investigated feature of marriage markets concerns the importance of size of place. We argue that intermarriage varies by size of place and that the relationship between intermarriage and the structural constraints of the local marriage market will also differ. We assume that residential segregation is lower in small towns than in large cities that provide more opportunity for group interaction. Thus, we expect that the relationship between the size of the migrant community and intermarriage will be stronger in small municipalities than in larger ones. In general, the model we present (i) incorporates the local dimension of marriage markets (i.e., does not assume collapsibility), (ii) examines the association between intermarriage and the structural constrains of the marriage market on the local scale, and (iii) investigates if intermarriage varies by size of place.

3. International migration and intermarriage in Spain

Our models are based on intermarriage between Spanish men and women and foreign-national women and men, respectively. Spain is considered a new country of immigration ([30-32]). Between 1998 and 2008, the foreign-born population in Spain grew from 0.6 to 5.2 million international migrants coming from many countries but mainly from Latin American countries (e.g., Ecuador, Colombia, Bolivia, Argentina, Peru), Africa (Morocco) and Eastern Europe (Romania). As a result, the total population of Spain grew from 39.9 to 46.2 million during the same period. After 2008, the economic recession has slowed the number of international arrivals. In fact, the migratory balance was negative between

2010 and 2015. The demographic impact of international migration is not limited to variations in the total population. International migrants are a source of other demographic impacts on mainly fertility and nuptial dynamics. Migrant women have contributed to the rise of fertility rates in Spain and to increase the number of births in recent years ([33-34]). Foreign born mothers constituted approximately 20% of the total number of children born in Spain from 2005-2013. Regarding nuptiality, international migration has had a notorious impact on marriage dynamics in Spain ([35]). According to the 2011 census, there were 0.5 million mixed couples in Spain of whom 66% were married and 34% were cohabiting.

The Spanish vital registration system has registered more than 300,000 marriages between Spanish citizens and foreign nationals since 2000. In this year, mixed marriages between Spanish nationals and foreign-born nationals represented 6.8% of all marriages. In 2013, the intermarriage rate had reached 14.6%. Several articles have documented patterns and trends of mixed marriages in Spain on the basis of marriage registration statistics ([35][36]), which only include marriages in Spain. This aspect is a limitation on the study of marriage patterns from the migrant perspective because it does not capture marriages abroad or cohabiting couples. However, with regard to intermarriage, Esteve and Bueno have shown that the vast majority of mixed marriages between Spanish citizens and foreign nationals are registered in the Spanish civil registration system ([37]). Data from these registers show that Spanish men intermarry more than Spanish women. From 1998 to 2012, Spanish men married mostly women from Latin America and Eastern Europe. Spanish women married men from Western and Southern Europe and Latin America. This pattern has been changing in recent years because the number of Spanish women marrying Moroccans has increased. This fact reflects second-generation Moroccans (born in Spain with Spanish citizenship) marrying Moroccan citizens ([35]) and shows that the country of citizenship is not necessarily the best indicator of ethnicity.

(**a**)

(**b**)
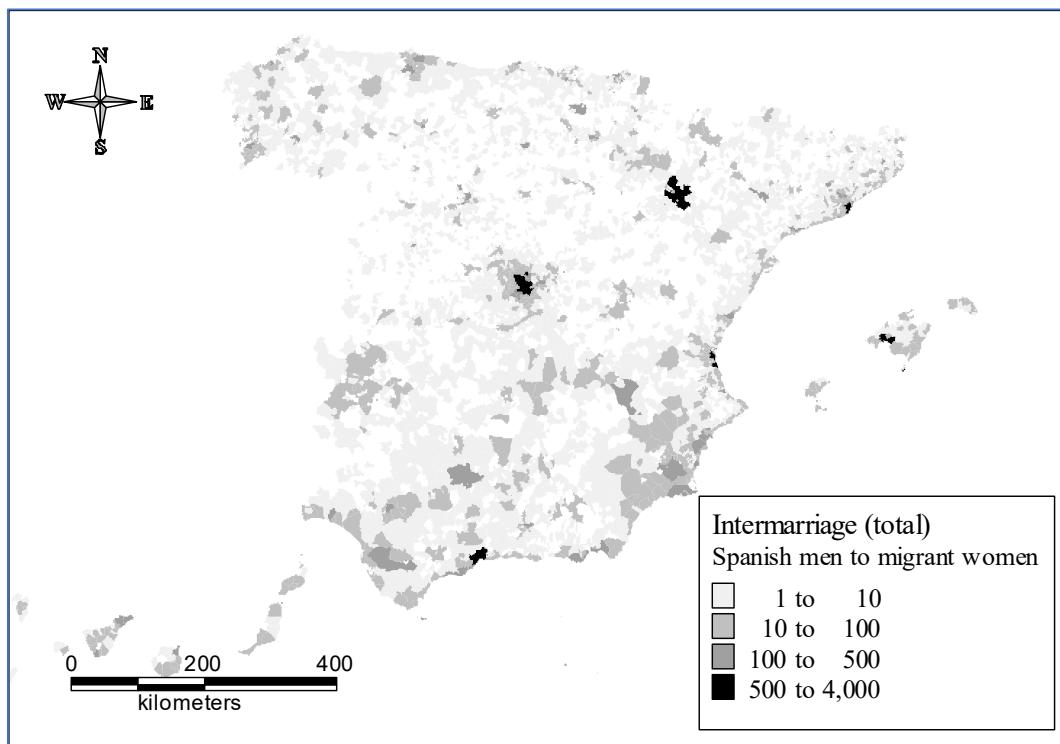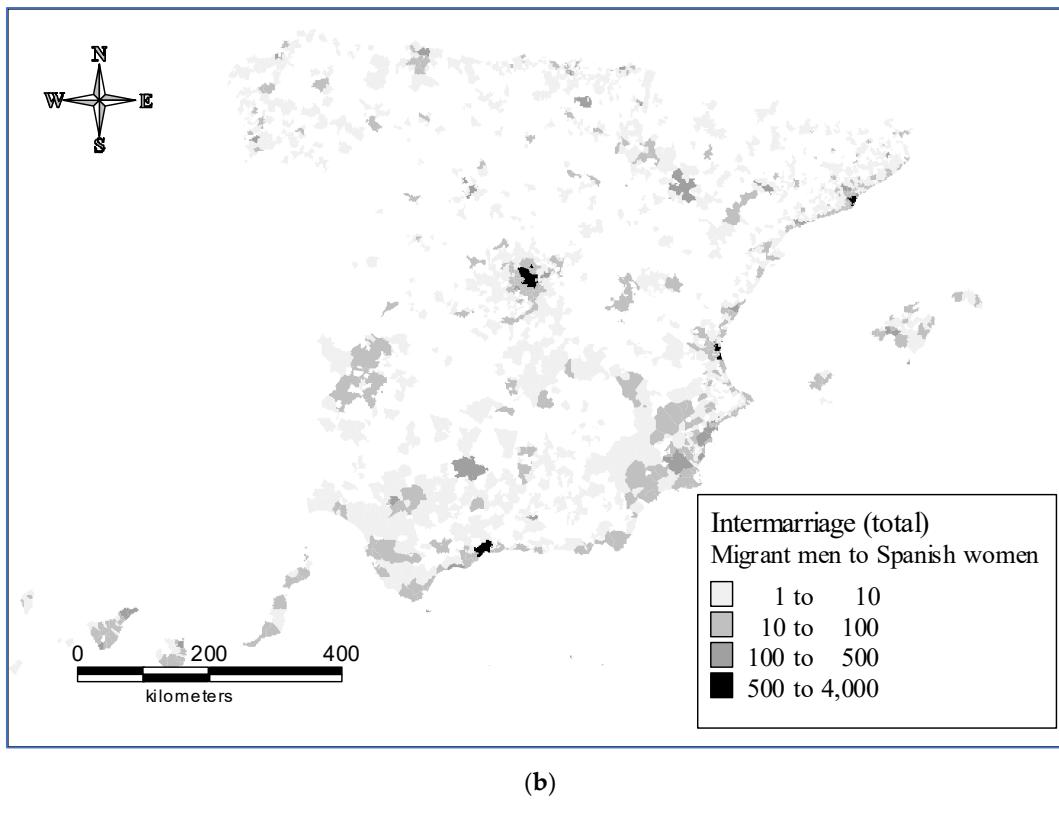
Figure 1. Spatial distribution of intermarried couples: (**a**) Spanish Men married to Migrant Women; (**b**) Spanish Women married to Migrant Men, 2005-2007.

From the migrant perspective, several studies based on census and survey data have noted that some migrant groups are more likely than others to intermarry with Spaniards. In general, African and Asian migrants are the least likely to marry Spanish nationals ([38-41]). In contrast, European and Latin Americans are the most likely to marry Spanish nationals. Most of these studies have been conducted at a national level. In general, there is a lack of knowledge concerning how intermarriage in Spain varies across regions and the role played by the structural characteristics of local marriage markets. In Figure 1, we show the distribution of intermarried couples by municipality who married in Spain from 2005-2007. We distinguish between intermarriage between Spanish men and migrant women (Figure 1a) and between Spanish women and migrant men (Figure 1b). As we explain in the coming section, these data come from the Spanish vital registration systems and basically reflect the distribution of our dependent variable, that is, the number of intermarriages by municipality. Generally, the municipality distribution of intermarried couples reflects the spatial distribution of the Spanish population, with Madrid, in the center, and Barcelona, on the northeast coast, registering the largest number of intermarried couples from 2005 to 2007. The main goal of this article is to model the spatial distribution of intermarriage considering the local characteristics of the marriage market regarding size of the population, heterogeneity of migrant origins and the sex ratios in the migrant and native communities.

4. Data and variables

4.1. Marriage and population register data

We utilized all marriage records from the MNP (Spanish vital registration statistics). The analysis is based on 42,972 heterosexual marriages between foreign national women and Spanish men and 26,310 marriages between foreign national men and Spanish women

that occurred in Spain from 2005-2007. By foreign nationals we refer to men and women that do not have Spanish citizenship. We pooled the data for these three years to increase the sample size. Marriages are geo-referenced at the municipality of residence. It must be said that the public use files of marriage microdata are yearly available from the INE (Spanish National Statistical Institute)'s website only for the capital cities and municipalities with more than 50,000 inhabitants (2% of the total number of municipalities in Spain, where 53% of the residents live). The present study has been possible due to a special request we made to the INE to obtain all marriage microdata at the municipality level, which was provided for a past period.

In 2007, Spain was divided into 8,111 municipalities, which are the smallest geo-political unit, though the final sample used in the models was reduced to 8,079 municipalities due to data availability problems. During the observed period, 3,326 municipalities registered at least one marriage between a migrant woman and a Spanish man, and 2,120 municipalities registered at least one marriage between a Spanish woman and a migrant man. This means that for the vast majority of municipalities, there were no intermarriages registered during this period and, therefore, our absolute and relative measures of intermarriage will be zero.

Although one advantage of official marriage data is the inclusion of all marriages contracted in Spain, the marriages contracted and registered in the MNP system are not representative of the unions that migrants form after migration. Our data does not include the unions of migrants who returned to their countries to marry or migrants who cohabit. The data only include unions of migrants who married in Spain. Although these data may not be representative of the unions that migrants form after migration, previous research has shown that the vast majority of unions between migrants and natives are registered in the Spanish marriage registration system. Migrants and natives are more likely to marry than cohabit, and they are more likely to marry in Spain than abroad (see [37],[42]).

We utilize all population counts at the municipal level from the "Padrón" (Population Register) to construct the dependent and independent model variables (relative group size, homogeneity of national origins, male/female availability for migrant women/men, and the sex ratio among natives). The "Padrón" offers annual counts of the population registered in each municipality by age, sex, country of birth and country of citizenship.

4.2. Construction of variables

Table 1 provides a detailed description of the dependent and independent variables included in the model and describes the statistical sources from which these variables were created.

4.2.1. The dependent variables

We use the following two measures of intermarriage as a dependent variable: the absolute number of intermarried couples (intermarriage) and the percentage of marriages between a Spanish national and a foreign national (intermarriage rate).

Intermarriage is the number (counts) of marriages in each municipality between Spanish citizens and foreign-national individuals in Spain from 2005-2007. We distinguish two types of intermarriage based on the sex of the spouses: Spanish men marrying foreign-national women and Spanish women marrying foreign-national men.

Intermarriage rate refers to the percentage of marriages between Spanish men/women and foreign national women/men of the total number of marriages contracted in each municipality from 2005 to 2007.

4.2.2. Independent variables

The model includes the following four independent or exogenous variables: (1) the relative size of the migrant population (i.e., foreign nationals); (2) the homogeneity of national origins among the migrant population; (3) the scarcity of migrant men/women for

the migrant women/men population; and (4) the sex ratio among natives. All variables were computed at the municipality level with data from the "Padrón".

Table 1. Description and source of variables utilized in the analysis.

Variables	Description	Source
Dependent variables:		
Intermarriage	Number of marriages in each municipality between Spanish citizens and foreign-national individuals in Spain from 2005-2007.	MNP (Spanish Vital Registration)
Intermarriage rate	Percentage of marriages between Spanish men/women and foreign national women/men out of the total number of marriages contracted in each municipality from 2005 to 2007.	MNP (Spanish Vital Registration)
Independent variables:		
Relative group size	Percentage of foreign national population in the municipality	"Padrón" (Population Register)
Homogeneity of national origin	Homogeneity of migrant origin in each municipality ($H = 100$ means high homogeneity among migrant origin; $H \approx 0$ means high heterogeneity among migrant origin). This indicator is the result of the summation of the squared proportion of each origin in the foreign population.	"Padrón" (Population Register)
Scarcity of men for migrant women	Proportion of foreign national women who do not have an opposite sex counterpart of the same national origin (%)	"Padrón" (Population Register)
Scarcity of women for migrant men	Proportion of foreign national men who do not have an opposite sex counterpart of the same national origin (%)	"Padrón" (Population Register)
Sex ratio among natives	Total number of Spanish men aged 15 to 64 divided by the total number of Spanish women aged 15 to 64 (%). Log transformation	"Padrón" (Population Register)

The relative size of the migrant community corresponds to the proportion of the foreign national population in each municipality. We expect this variable to be positively related to intermarriage.

The homogeneity of national origins among the migrant population equals the summation of the squared proportion of each origin. This variable equals 1 when all foreign nationals hold the same nationality and tends to 0 as the number of origins increase. We expect this variable to be negatively related to intermarriage.

The scarcity of migrant men/women for migrant women/men is equal to the proportion of foreign national women/men that do not have an opposite sex counterpart of the same national origin. This indicator expresses in relative terms the scarcity of men/women of each national origin. For example, a value of scarcity of 1 indicates that each foreign national men/woman does not have a female/male counterpart of the same origin. A value of 0.6 indicates that for 60% of foreign national men/women, there is no female/male counterpart, indicating that a man of the same origin exists for only 40% of these men/women. We expect that the scarcity of men/women of the same national origin will be directly related to the intermarriage rate.

The sex ratio among natives is the log transformation of the ratio between the number of Spanish men aged 15-64 and the number of Spanish women aged 15-64 for each municipality. We expect that the sex ratio will be directly related to the intermarriage rate.

4.2.3. Spatially lagged variables

The structural variables described above were constructed at the municipality level. Neighboring areas may exert an additional influence on local intermarriage. To capture the existence of spatial dependence in intermarriage, we constructed the spatial lags of the independent variables. Spatial dependence reflects a situation where values observed in one location or region depend on the values of neighboring observations in nearby locations. A variable spatial lag is a variable vector constructed with a (weighted) average of

values from neighboring municipalities or regions. This calculation places elements w_{ij} in the $n \times n$ spatial weight matrix $\mathbf{W}$, for n the number of municipalities, such that $\sum_{j=1}^{n} w_{ij} x_{ij}$ results in a scalar representing a linear combination of values obtained by neighboring observations ([43]). In this paper, the spatial weight matrix utilizes an inverse distance function for a 170 km distance band, such that each element (w_{ij}) is equal to $1/d_{ij}^2$ for each pair of municipalities i, j located at a distance $d_{ij} \leq 170$ km and zero otherwise (see [44] for further information). We utilize 170 km as the minimum distance at which every municipality has at least one neighbor.

We have built the spatial lags of the four independent variables: relative group size, homogeneity of national origin, scarcity of men/women for migrant women/men and sex ratio among natives.

4.2.4. Size of municipality

The municipalities were classified in nine groups based on the number of inhabitants. We utilized the classification provided by the Spanish National Statistical Institute to present the results. The classification ranges from less than 250 inhabitants to more than 100,000 (Table 2). Size of place was not included as an independent variable in the model. Instead, we specified a spatial regimes model in which we estimate a different model for each of the municipal groups to test whether the level of intermarriage and the impact of structural variables vary by size of municipality

Table 2. Descriptive statistics for the variables utilized in the models.

Variables	Size of municipality								
	0-249	250-499	500-999	1,000-2,499	2,500-4,999	5,000-9,999	10,000-24,999	25,000-99,999	>100,000
Number of municipalities (n)	2,581	1,214	1,082	1,256	712	545	426	236	59
Total marriages	1,584	3,027	6,646	21,744	32,217	52,717	96,115	152,378	239,316
Relative group size, avg. %	3.4	4.5	5.9	6.1	7.6	7.7	10.1	11.6	9.4
Sex Ratio among nationals, avg.	1.47	1.27	1.20	1.13	1.09	1.06	1.04	1.02	0.97
Spanish Men with Foreign National Women:									
Municipalities with intermarriage*	146	221	386	747	584	522	425	236	59
Intermarriages	175	281	530	1,484	1,820	3,044	6,282	10,915	18,441
Total marriages	1,584	3,027	6,646	21,744	32,217	52,717	96,115	152,378	239,316
Proportion of intermarriages	11.0	9.3	8.0	6.8	5.6	5.8	6.5	7.2	7.7
Migrant women	752	2,514	9,385	41,144	75,976	130,710	308,917	590,004	914,725
Homogeneity of national origin, avg. %**	60.1	45.4	40.2	34.6	30.6	28.1	25.2	22.1	18.6
Scarcity of men for migrant women, avg. %**	56.7	42.3	30.1	22.9	18.6	16.7	13.1	12.0	12.7
Spanish Women with Foreign National Men:									
Municipalities with intermarriage*	35	61	134	361	391	421	413	236	59
Intermarriages	38	67	151	509	793	1,377	3,375	6,625	13,375
Proportion of intermarriages	2.4	2.2	2.3	2.3	2.5	2.6	3.5	4.3	5.6
Migrant men	165	804	5,054	27,733	69,861	140,283	359,471	672,837	982,838
Homogeneity of national origin, avg. %**	57.2	45.7	38.9	34.1	30.6	27.1	25.1	22.1	18.6
Scarcity of women for migrant men, avg. %**	24.3	29.5	20.2	19.0	15.0	15.1	12.7	12.0	12.7

Source: Self elaboration based on data from Spanish Vital Registration and Population Register (MNP). *Municipalities with at least one intermarriage of the referred type from 2005-2007. ** Based on the municipalities with intermarriage.

5. Model specifications

Building a robust model to predict spatial variation of intermarriage at the local level is a complex task. In this section, we document the process of model specification and testing.

5.1. Zero Inflated Poisson (ZIP) with spatial effects

The first model estimates the absolute number of local intermarried couples as a Zero Inflated Poisson (ZIP) model, which allows controlling for the excess of zero counts corresponding to the municipalities without mixed marriages ([45]). Because local intermarriages are both counts and rare events, a Poisson distribution is the best option. The excess zeros are supposed to be generated by a separate process from the count values and they must be modeled independently. Thus, the ZIP model must estimate two models, a Poisson count model, and the logit model for predicting excess zeros. It has the following three parts:

- A Probability Mass Function (PMF), $P(y_i = 0)$, which is used to calculate the probability of observing a zero count.
- A PMF, $P(y_i = k)$, which is used to calculate the probability of observing k events, given that $k > 0$.
- A link function used to express the mean rate, λ, as a function of p regression variables **X**.

Hence, the PMF of the ZIP regression model can be specified as a mixture model of the following two component distributions:

$$P(y_i = 0) = \phi_i + (1 - \phi_i) \cdot e^{-\lambda_i}$$

$$P(y_i = k) = (1 - \phi_i)\frac{e^{-\lambda_i} \cdot \lambda_i^k}{k!}$$

(1)

where the counts y_i, for $i = 1,2, \dots, n$ observations, equals zero with probability ϕ_i and follow a Poisson distribution with mean λ_i and probability $(1 - \phi_i)$. Hence, ϕ_i is the proportion of excess zeros for the ith observation (e.g., a Spanish municipality), and $\lambda_i = e^{x_i\beta}$, for x_i the value of one of the p explanatory variables of the regression model for the ith observation and β the vector of the regression coefficients. The second part of equation (1), $P(y_i = k)$, corresponds with the standard Poisson regression model for y_i the random dependent variable that denotes the observed count of the ith observation.

Note that ZIP models assume that an ith observation is 0 with a probability ϕ_i or is a realization of a Poisson random variable, which can also be 0, with a probability $(1 - \phi_i)$. That is, zero observations arise from both the zero-component distribution, $P(y_i = 0)$ and the Poisson distribution, $P(y_i = k)$. The zero-component distribution is related to model the 'excess' or 'inflated' zeros that are observed in addition to the zeros that are expected to be observed under the assumed Poisson distribution [46]. When the data set does not have any excess zeros in the dependent variable, the value of $\phi_i = 0$ and the PMF of the ZIP model reduces to the PMF of the standard Poisson model [47].

The way of estimating the proportion of excess zeros, ϕ_i, is by estimating them through a logistic function of p regression variables **X**, as follows:

$$P(y_i' = 0) = \mu_i = \frac{e^{x_i\gamma}}{1 + e^{x_i\gamma}}$$

$$P(y_i' = 1) = (1 - \mu_i)$$

(2)

where y' is a binary 0/1 random variable which adopts the value 0 if the underlying dependent variable of counts, y is 0, and 1 in all the other cases; and γ is a vector of parameters. The estimation of this logistic model yields a vector of n fitted probabilities, $\hat{\mu}_i$, which is simply set to ϕ_i ($\hat{\mu}_i = \hat{\phi}_i$). Once estimated, we set vector $\hat{\phi}_i$ into the probability

functions of the ZIP model presented in equation (1) and estimate it by Maximum Likelihood (ML).

Hence, in a ZIP model, ϕ_i and λ_i can be explicitly expressed as a function of explanatory variables **X**, as follows:

$$logit(\phi_i) = x_i'\gamma$$
$$log(\lambda_i) = x_i'\beta \tag{3}$$

As in [48], we suppose the existence of a potential impact of social interactions—in the form of local spatial spillovers—on the probability of a municipality of having mixed marriages. This spatial version of ZIP model takes the form of a cross-sectional spatially lagged SLX model ([49-51]). It is an augmented model which incorporates the spatial lags of the explanatory variables in the expressions of equation (3) as follows:

$$dlogit(\phi_i) = x_i'\gamma + (\mathbf{W}x)_i'\theta$$
$$log(\lambda_i) = x_i'\beta + (\mathbf{W}x)_i'\xi \tag{4}$$

Additionally, we also test for the existence of spatial instability in the form of nine municipality groups based on size of place to find out whether the level of intermarriage and the impact of structural variables vary by municipality size, like in [52]. It is a so-called spatial regimes regression model. Hence, the specification of equation (4) is as follows:

$$logit(\phi_i) = x_i'\gamma^g + (\mathbf{W}x)_i'\theta^g$$
$$log(\lambda_i) = x_i'\beta^g + (\mathbf{W}x)_i'\xi^g \tag{5}$$

where g represents the different municipality groups.

5.2. Probit model for grouped data with spatial effects

The second model estimates local intermarriage rates between migrant women and Spanish men and between migrant men and Spanish women as a probit model of grouped data ([53]), because the dependent variable consists of a number of observed proportions or relative frequencies of sets of individuals who share similar characteristics (e.g., living in the same municipality). In this case, the dependent variable consists of the n number of observed proportions (p_i) of mixed marriages over the total number of marriages contracted in each municipality. We disregard other possible specifications like Tobit regression ([54]) or beta regression models ([55]) because they suppose that the dependent variable is either continuous (though censored by the modeler for somewhat reasons) or bounded between 0 and 1, excluding extreme values, respectively.

These observed proportions can be considered as the realization of an indirect utility function:

$$\pi_i = F(x_i'\beta) \tag{6}$$

where π_i is the theoretical or population proportion of the ith observation, x_i is a vector gathering a set of k variables which explains the intermarriage decision and β contains a set of parameters.

One of the functional forms most frequently used in application for the F function is the probit model, which by means of the Slutsky's theorem on convergence in probability, can be linearized. The Cumulative Distribution Function (CDF) of the standard normal distribution is expressed as $\Phi(x_i'\beta)$. Since the CDF is strictly monotonic, it is an inverse form, $z_i = \Phi^{-1}(p_i)$, which by means of a Taylor series approximation leads to the probit model for grouped data or 'gprobit' model ([56]):

$$z_i = \Phi^{-1}(p_i) = x_i'\beta + \varepsilon_i \tag{7}$$

where ε_i is a well-behaved error term normally distributed with zero mean but heteroskedastic variance, which is given by the following expression:

$$\sigma_i^2 = \frac{p_i(1 - p_i)}{n_i[\phi[\Phi^{-1}(p_i)]]^2} \tag{8}$$

where ϕ is the Probability Distribution Function (PDF) of the standard normal distribution. We can use a two-step estimator to obtain the Weighted Least Squares (WLS), which will be applied to the observations with weights proportional to $1/\sigma_i^2$ ([57-58]).

Since the gprobit model shown in equation (7) is a linear model, it is possible to specify local spatial spillovers, in the form of spatial lagged explanatory variables, and spatial regimes, for the municipality groups, leading to the following augmented gprobit model:

$$z_i = \Phi^{-1}(p_i) = x_i'\beta^g + (\mathbf{W}x)_i'\xi^g + \varepsilon_i \tag{9}$$

5. Results

5.1. Modeling the number of intermarriages

Table 3 shows the Maximum Likelihood (ML) estimations of the ZIP model for both marriages between Spanish men and migrant women (first panel) and between Spanish women and migrant men (second panel). These results can be replicated using the database and coding available from [59]. In the second column of Table 3, this model does not differentiate between size of place, i.e., all municipalities regardless of size are pooled in the same regression. Intermarriage counts are predicted based on the four independent variables and their corresponding spatial lags to control for spatial autocorrelation effects (all computations and tests, such as spatial autocorrelation Moran's *I* and Spatial Chow-Wald tests, are available from the authors upon request).

From columns three to eleven, the ZIP model tests the existence of spatial instability in the form of nine municipality groups based on size of place and indicates whether the level of intermarriage and the impact of structural variables vary by municipality size. For both types of mixed marriages, Table 3 results show heterogeneity in the coefficients across the municipality groups. Furthermore, some of the spatially lagged variables are highly significant for certain groups. Both results demonstrate the bias of the general estimation (in column 2) as well as the need to control for spatial dependence. To assess the goodness of fit of each model, we computed the Pseudo R^2 with and without the spatially lagged variables. The inclusion of the spatially lagged variables enhances the explanatory capacity of all models though the improvement is better in the larger municipalities. These results suggest that the influence of the characteristics of neighboring municipalities on intermarriage is higher in larger municipalities than in smaller ones.

Regarding the coefficients and how many of them are statistically significant, the relative size of the migrant community turns out to be the most important factor. The relative size of the migrant community, when significant, is positively related to the number of intermarried couples. All coefficients but four are significant, at least, at the 10% level. Municipalities with larger migrant communities have more intermarriage than those with smaller communities.

The impact of migrant origin homogeneity is lower than the group's relative size and becomes significant in 7 of the 9 municipality groups for intermarriage between Spanish men and migrant women and in 3 of the 9 municipality types for Spanish women and migrant men. Consistent with our expectations, the homogeneity of migrant origin, when significant, is inversely related to intermarriage. This result shows that there is less intermarriage in municipalities whose migrant communities are more homogeneous in terms of national origins.

Contrary to the previous two variables, the impact of sex ratios among both migrants and natives on intermarriage is less consistent. The sex ratio coefficients are neither uniform across sizes of municipalities nor between the two types of intermarriage. We

expected that skewed sex ratios would result in more intermarriage. This expectation is true in the groups of large cities and also in the municipalities with less than 1,000 inhabitants where migrant communities are female-dominated. In these municipalities, the scarcity of migrant men for migrant women is related to more intermarriage.

Table 3. Regression results for absolute intermarriage (Zero Inflated Poisson model).

Variables	Total	Size of municipality								
		0-249	250-499	500-999	1,000-2,499	2,500-4,999	5,000-9,999	10,000-24,999	25,000-99,999	>100,000
Spanish Men with Foreign National Women:										
Relative group size (G)	0.036***	0.016	0.019	0.015*	0.012**	0.010**	0.028***	0.019***	0.024***	0.095**
Homogeneity of national origin (H)	-0.062***	-0.009	-0.012**	-0.011**	-0.013***	-0.012***	-0.010***	-0.007**	-0.006*	0.007
Scarcity of men for migrant women (S)	-0.013***	0.010**	0.009*	0.003	-0.002	-0.005*	0.003	-0.002	0.012	0.011
Sex Ratio among natives (X)	-0.167***	-0.001	0.005	-0.011***	0.000	0.009	-0.003	-0.019**	-0.047***	-0.265***
Spatially Lagged Variables:										
Relative group size	-0.078	0.085	0.129	0.008	0.063*	0.114***	0.073***	0.024	-0.004	-0.024
Homogeneity of national origin	-0.049**	0.037	-0.005	0.004	0.005	0.003	0.012	-0.004	-0.029**	-0.055
Scarcity of men for migrant women	-0.070	0.056	0.060	-0.032	0.013	0.037**	0.023	0.006	-0.023	-0.108
Sex Ratio among natives	0.014*	0.000	0.002	0.001	0.008***	0.014***	0.009***	0.002	0.005	0.023*
Constant	23.703**	-4.859	-3.703	1.366	-1.164	-3.066***	-0.545	4.272***	8.824***	31.977***
Pseudo R^2	0.4103	0.266	0.299	0.290	0.420	0.481	0.479	0.380	0.383	0.583
SL Pseudo R^2	0.4299	0.277	0.300	0.303	0.427	0.525	0.515	0.385	0.451	0.694
Spanish Women with Foreign National Men:										
Relative group size (G)	0.033***	0.122**	0.076***	-0.015	-0.002	0.020***	0.027***	0.016***	0.023***	0.022*
Homogeneity of national origin (H)	-0.048***	0.002	-0.001	-0.005	-0.004	-0.016***	-0.018***	-0.014***	0.011	-0.001
Scarcity of women for migrant men (S)	-0.020***	0.031***	0.007*	-0.012**	-0.002	-0.008***	-0.010***	-0.008***	-0.018***	-0.024**
Sex Ratio among natives (X)	-0.258***	-0.006	0.013	-0.029***	-0.018**	-0.016*	-0.011	-0.015	-0.036**	-0.246***
Spatially Lagged Variables:										
Relative group size (WG)	0.073*	0.021	0.183**	-0.031	0.076**	0.074***	0.021	0.050***	0.039**	0.194***
Homogeneity of national origin (WH)	-0.050	-0.085	0.016	0.000	0.015	-0.018	-0.010	-0.006	-0.030*	-0.064
Scarcity of women for migrant men (WS)	-0.026	0.023	0.019	-0.024	-0.002	0.011	-0.003	-0.012	0.025*	0.023
Sex Ratio among natives (WX)	0.012	-0.001	0.033**	0.002	0.010**	0.013***	0.005	0.004	0.002	0.021
Constant	30.986***	-1.126	-10.02**	3.611*	-0.470	-2.577**	2.445*	3.826***	6.579***	26.838***
Pseudo R^2	0.216	0.127	0.161	0.309	0.341	0.449	0.506	0.415	0.439	0.569
SL Pseudo R^2	0.363	0.148	0.190	0.317	0.354	0.459	0.512	0.473	0.553	0.709

Note: *** significant at 0.01, ** significant at 0.05, * significant at 0.10. Pseudo R^2: correlation coefficient between the real and estimated dependent variables. SL Pseudo R^2: Pseudo R^2 for the model with the spatially lagged variables. Standard errors: Huber/White/sandwich robust variance estimator. Source: Self elaboration based on data from the MNP and "Padrón".

Regarding intermarriage between Spanish women and migrant men, the scarcity of migrant women for migrant men has mainly a negative effect on intermarriage,

suggesting that in municipalities with male-dominated migrant communities, Spanish women are less likely to intermarry than in municipalities with more balanced migrant communities. The explanation of this paradox may be in the origin of the migrants of the male-dominated communities. As previous research has shown on a national scale, some migrants are more likely to intermarry than others. For example, Moroccans are the least likely to marry Spanish women with no Moroccan ancestry though Moroccan migrant communities in Spain are traditionally male dominated.

The sex ratio among natives has not shown the expected impact. Sex ratios among natives were expected to increase intermarriage showing the spatially lagged variable the expected positive relationship. The excess of males in neighboring areas increases inter-marriage between Spanish males and migrant women in a determined municipality. In the case of Spanish women marrying migrant men, sex ratios among natives, when statistically significant, reduce intermarriage, meaning that in municipalities where native women are scarce, less Spanish women intermarry.

5.2. Modeling the intermarriage rate

Table 4 shows the Weighted Least Squares estimations of the gprobit model for both intermarriage rates between Spanish men and foreign women (first panel) and between Spanish women and foreign men (second panel). These results can be replicated using the database and coding available from [59]. In the second column of Table 4, this model does not differentiate among size of place, i.e., all municipalities regardless of size are pooled in the same regression. Intermarriage rates are predicted based on the four independent variables and their corresponding spatial lags to control for spatial autocorrelation effects. From columns three to eleven, the gprobit model tests the existence of spatial instability in the form of nine municipality groups based on size of place and indicates whether intermarriage rates and the impact of structural variables vary by municipality size. Table 4 results for both type of mixed marriages show heterogeneity in the coefficients across the municipality groups. Furthermore, several spatially lagged variables are highly significant for certain groups. Both results demonstrate the bias of the general estimation (in column 2) and the need to control for spatial dependence. We also report the R^2 with and without the spatially lagged variables.

Overall, modeling the intermarriage rate yields more statistically significant parameters, and they are more consistent with our initial hypotheses. The relative size of the migrant community in the municipality is the explanatory variable with the highest number of statistically significant coefficients and is the most influential to intermarriage. The relationship between relative size of the migrant community and the intermarriage rate is positive for all sizes.

The homogeneity of migrant origin is statistically significant in only 3 sizes of municipalities for intermarriage between Spanish men and migrant women and also for intermarriage between Spanish women and migrant men. Inconsistently with our expectations, the homogeneity of migrant origin, when significant, is positively related to intermarriage between migrant men and Spanish women, which occurs in rural municipalities with less than 1,000 inhabitants. This result shows that there is more intermarriage of this type in villages whose migrant communities are more homogeneous in terms of national origins.

The scarcity of migrant men for migrant women of the same origin is positively related to the share of intermarriage between Spanish men and migrant women in municipalities of sizes from 5,000 to 100,000 inhabitants. These results indicate that when migrant communities are female-dominated, the intermarriage rate between Spanish men and migrant women is higher. Regarding intermarriage between Spanish women and migrant men, the impact is negative for practically all municipality sizes indicating that in male-dominated migrant communities, the intermarriage rate between Spanish women and migrant men is lower than in municipalities with gender-balanced migrant communities.

Table 4. Regression results for the intermarriage rate (probit model for grouped data).

Variables	Total	0-249	250-499	500-999	1,000-2,499	2,500-4,999	5,000-9,999	10,000-24,999	25,000-99,999	>100,000
Spanish Men with Foreign National Women:										
Relative group size (G)	0.019***	0.010	0.005	0.009***	0.013***	0.012***	0.022***	0.018***	0.017***	0.023***
Homogeneity of national origin (H)	-0.004***	-2.3-e^{-4}	0.002	0.002	0.003**	0.001	1.0-e^{-4}	0.001	-0.005*	0.006*
Scarcity of men for migrant women (S)	0.006***	0.002	-2.3-e^{-4}	0.001	0.001	0.001	0.004***	0.004**	0.006**	0.002
Sex Ratio among natives (X)	-0.002	0.001	3.7-e^{-4}	0.010***	0.009***	0.007**	0.006*	-0.003	-0.005	-0.024***
Spatially Lagged Variables:										
Relative group size (WG)	0.039***	0.054	-0.081	0.008	0.052***	0.072***	0.071***	0.057***	0.044***	0.019
Homogeneity of national origin (WH)	-0.008	0.051	0.013	0.031***	0.016***	0.004	0.011**	0.003	-0.005	-0.018**
Scarcity of men for migrant women (WS)	0.010**	0.007	0.022	-0005	0.029***	0.035***	0.030***	0.028***	0.013	0.006
Sex Ratio among natives (WX)	0.004***	1.7-e^{-4}	-0.015***	-0.008***	-0.003**	0.002**	0.002**	0.002	0.003**	0.004***
Constant	-2.606***	-3.240	0.059	-2.459***	-3.905***⁾	-4.251***	-4.371***	-2.796***	-2.154***	0.467
R^2	0.318	0.059	0.038	0.088	0.098	0.127	0.315	0.383	0.460	0.653
SL R^2	0.359	0.139	0.180	0.147	0.118	0.161	0.350	0.419	0.525	0.724
Spanish Women with Foreign National Men:										
Relative group size (G)	0.015***	-0.021	-0.013	0.005	0.010***	0.013***	0.017***	0.013***	0.016***	0.029***
Homogeneity of national origin (H)	0.002*	0.011**	0.009**	0.002	0.005***	0.001	-0.001	-0.002	0.002	0.004
Scarcity of women for migrant men (S)	-0.003***	-0.005	-0.002	-4.9-e^{-4}	0.001	0.001	-2.4-e^{-4}	-0.002**	-0.005***	-1.4-e^{-4}
Sex Ratio among natives (X)	-0.010***	0.002	0.004	0.004	0.003	0.006**	-0.006	-0.007	-0.006	-0.033***
Spatially Lagged Variables:										
Relative group size (WG)	0.025***	-0.110	-0.077	-0.002	0.012	0.009	0.012*	0.031***	0.028***	0.003
Homogeneity of national origin (WH)	-0.010**	-0.061	-0.014	0.007	0.013*	-0.004	-0.002	-0.003	-0.010	-0.025***
Scarcity of women for migrant men (WS)	0.001	-0.006	-0.008	0.016	-5.3-e^{-6}	0.001	-0.005	-0.007	0.007	-0.008
Sex Ratio among natives (WX)	0.004***	-0.010	-0.012**	-0.007**	-0.005***	-0.002*	-1.5-e^{-4}	0.004***	0.002	0.002
Constant	-1.224***	-3.674	1.292	-1.691*	-2.213***	-2.258***	-1.198*	-1.503**	-1.552*	2.147**
Pseudo R^2	0.301	0.214	0.180	0.046	0.178	0.204	0.224	0.322	0.489	0.554
SL Pseudo R^2	0.364	0.329	0.304	0.131	0.220	0.245	0.253	0.394	0.554	0.686

Note: ** significant at 0.01, * significant at 0.05. R^2: R^2 for the model without the spatially lagged variables. SL R^2: R^2 for the model with the spatially lagged variables. Standard errors: Huber/White/sandwich robust variance estimator. Source: Self elaboration based on data from the MNP and "Padrón".

The relationship between the sex ratio among natives and the intermarriage rate is statistically significant in five municipality sizes for intermarriage between Spanish men and migrant women and significant in only two municipality sizes for intermarriage between Spanish women and migrant men. In both cases of intermarriage, the model yields the expected negative sign only for the group of large cities above 100,000 inhabitants. This result means that in municipalities where the number of native men or women is far

greater than the number of women or men, there is little intermarriage between Spanish women and migrant men or Spanish men and migrant women, respectively.

6. Discussion and conclusions

In this paper, we presented a spatial regimes model to predict the absolute number and relative importance of intermarriage in Spain between native men and migrant women and between native women and migrant men. Our models combine three main features. First, the models did not assume collapsibility of the marriage market. Instead, we examined intermarriage rates at the local level based on structural characteristics of local marriage markets and their neighboring areas. The structural characteristics of the neighboring areas were specified with spatially lagged variables that utilized a spatial weight matrix to determine the vicinity relationship among spatial units. Second, the explanatory variables in the model included the following structural characteristics at the local level: the relative size of the migrant community, the homogeneity of migrant origin, and the sex ratios in migrant and native communities. Third, a spatial regime model was implemented to examine significant differences in intermarriage by size of place.

Regarding our modeling strategy, results show that the local characteristics of the marriage markets have an impact on both the absolute and relative importance of inter-marriage and that these effects vary by size of place. We have also shown that including spatially lagged variables, which means considering the influence of neighboring areas, always improves the fit of the models, particularly among the medium and large munici-palities. The explanatory capacity of our models, measured as R^2, is higher in the largest municipalities. However, this result does not exclude finding statistically significant coefficients among the smallest municipalities as well.

The results indicate that intermarriage is more likely to occur where there are communities of migrants and where these migrants are diverse in terms of origin. The more migrants and the more diverse they are, the more intermarriage there will be. This relationship holds true for both the absolute and relative measures of intermarriage. Specifically, the relative size of the migrant community turns out to be the most important factor for both models and intermarriage types. When significant, the relative size of the migrant community is positively related to the number of intermarried couples. Hence, municipalities with larger migrant communities have more intermarriage than those with smaller communities.

However, in the relative measure of intermarriage, the impact of homogeneity of migrant origin on intermarriage provides additional insights, since it is positively related with intermarriage between foreign men and Spanish women in smaller municipalities. Residential segregation may explain this fact. In small municipalities, the presence of migrants from a same origin -mainly when they are from a Latin American country- encourages social integration. Conversely, the coexistence of migrants of different nationalities and cultures in a village, typically farmworkers, might produce higher levels of segregation.

The impact of the sex ratio in the migrant and native communities on intermarriage is not uniform across all municipality sizes and does not always have the expected positive sign. The clearest impact of sex ratios is that the relative importance of intermarriage between Spanish men and migrant women is higher in municipalities whose migrant communities are female-dominated. In contrast, male-dominated migrant communities are not associated with more intermarriage between Spanish women and migrant men. These results suggest that male-dominated migrant communities are not enough to encourage intermarriage among native women. The social cultural background of these migrants may play a role as well. For example, male-dominated migrant communities in Spain generally consist of Moroccan migrants, and research has shown that these migrants are very unlikely to marry native Spanish women with no Moroccan ancestry.

Future research should investigate further the relationship between the characteristics of the local marriage markets and the levels and opportunities for intermarriage. If

data availability were not a concern, residential segregation indicators in municipalities is an important dimension to consider; this is as important as the social and cultural backgrounds of the migrant communities who settled in each municipality. The ideal study should consider all types of unions but, more importantly, all marriages regarding the country where they were contracted. This consideration will provide us with a broader understanding of intermarriage dynamics in local communities and will enable us to disentangle whether the structural opportunities of the marriage market operate in the same manner for migrants as for natives. The importance given in this article to the local context of opportunities should not undermine the importance of individual preferences and expectations concerning marriage and intermarriage.

Author Contributions: Conceptualization, A.E. and T.L-G.; methodology, A.E. and C.C.; software, C.C. and T.L-G.; validation, A.E., C.C. and T.L-G.; formal analysis, A.E., C.C. and T.L-G.; investigation, A.E. and C.C.; resources, A.E., C.C. and T.L-G.; data curation, A.E. and T.L-G; writing—original draft preparation, A.E.; writing—review and editing, A.E., C.C. and T.L-G.; visualization, C.C. and T.L-G.; project administration, A.E.; funding acquisition, A.E., C.C. and T.L-G. All authors have read and agreed to the published version of the manuscript.

Funding: GLOBFAM project, Ministerio de Ciencia e Innovación de España, grant number RTI2018-096730-B-I00; CERCA Programme, Generalitat de Catalunya; Talent Programme, Universitat Autònoma de Barcelona; eS-MiData Project, Comunidad de Madrid, grant number 447379001.

Data Availability Statement: The data and scripts are available in the Spatial and Regional Economics research group (ECONRES) of the Universidad Autónoma de Madrid at https://econ-resuam.wordpress.com/opendata.

Conflicts of Interest: The authors declare no conflicts of interest.

References

1. Harris, D.R.; Ono, H. How Many Interracial Marriages Would There Be If All Groups Were of Equal Size in All Places? A New Look at National Estimates of Interracial Marriage. *Soc. Sci. Res.* **2005**, *34*, 236–251, doi:10.1016/j.ssresearch.2004.01.002.

2. Rontos, K.; Salvati, L. Space Matters? Exploring Gender Differentials in the Age at Marriage, Greece (1980–2017). *Soc. Sci.* **2020**, *9*, 59, doi:10.3390/socsci9040059.

3. Boschman, S. Residential Segregation and Interethnic Contact in the Netherlands. *Urban Stud.* **2012**, *49*, 353–367, doi:10.1177/0042098011402236.

4. Asibey, M.O.; Poku-Boansi, M.; Adutwum, I.O. Residential Segregation of Ethnic Minorities and Sustainable City Development. Case of Kumasi, Ghana. *Cities* **2021**, *116*, 103297, doi:10.1016/j.cities.2021.103297.

5. Kalmijn, M. Intermarriage and Homogamy: Causes, Patterns, Trends. *Annu. Rev. Sociol.* **1998**, *24*, 395–421, doi:10.1146/annurev.soc.24.1.395.

6. Lichter, D.T.; Qian, Z. The Study of Assortative Mating: Theory, Data, and Analysis. In *Analytical Family Demography*; Schoen, R., Ed.; The Springer Series on Demographic Methods and Population Analysis; Springer International Publishing: Cham, 2019; pp. 303–337 ISBN 978-3-319-93227-9.

7. Rosenfeld, M.J. *The Age of Independence: Interracial Unions, Same-Sex Unions, and the Changing American Family*; Harvard University Press: Cambridge, Mass, 2007; ISBN 978-0-674-02497-7.

8. Zhang, X.; Sassler, S. The Age of Independence, Revisited: Parents and Interracial Union Formation Across the Life Course. *Sociol. Forum* **2019**, *34*, 361–385, doi:10.1111/socf.12501.

9. Blau, P.M.; Blum, T.C.; Schwartz, J.E. Heterogeneity and Intermarriage. *Am. Sociol. Rev.* **1982**, *47*, 45–62, doi:10.2307/2095041.

10. Hwang, S.-S.; Saenz, R.; Aguirre, B.E. Structural and Assimilationist Explanations of Asian American Intermarriage. *J. Marriage Fam.* **1997**, *59*, 758–772, doi:10.2307/353959.

11. Kalmijn, M.; Van Tubergen, F. A Comparative Perspective on Intermarriage: Explaining Differences Among National-Origin Groups in the United States. *Demography* **2010**, *47*, 459–479.

12. Bohm-Jordan, M.; Yang, P.Q. Determinants of Intermarriages among Foreign-Born Asians in the United States. *Marriage Fam. Rev.* **2021**, *57*, 59–91, doi:10.1080/01494929.2020.1753137.

13. Corti, G.; Scherer, S. Mating Market and Dynamics of Union Formation. *Eur. J. Popul.* **2021**, *37*, 851–876, doi:10.1007/s10680-021-09592-2.

14. Massey, D.S.; Denton, N.A. The Dimensions of Residential Segregation. *Soc. Forces* **1988**, *67*, 281–315, doi:10.2307/2579183.

15. Telles, E.; Esteve, A. Racial Intermarriage in the Americas. *Sociol. Sci.* **2019**, *6*, 293–320, doi:10.15195/v6.a12.

16. Bossard, J.H.S. Residential Propinquity as a Factor in Marriage Selection. *Am. J. Sociol.* **1932**, *38*, 219–224.

17. Haandrikman, K.; Harmsen, C.; van Wissen, L.J.G.; Hutter, I. Geography Matters: Patterns of Spatial Homogamy in the Netherlands. *Popul. Space Place* **2008**, *14*, 387–405, doi:10.1002/psp.487.

18. Hannemann, T.; Kulu, H.; Rahnu, L.; Puur, A.; Hărăguş, M.; Obućina, O.; González-Ferrer, A.; Neels, K.; Van den Berg, L.; Pailhé, A.; et al. Co-Ethnic Marriage versus Intermarriage among Immigrants and Their Descendants: A Comparison across Seven European Countries Using Event-History Analysis. *Demogr. Res.* **2018**, *39*, 487–524, doi:10.4054/DemRes.2018.39.17.

19. Qian, Z.; Lichter, D.T.; Tumin, D. Divergent Pathways to Assimilation? Local Marriage Markets and Intermarriage among U.S. Hispanics. *J. Marriage Fam.* **2018**, *80*, 271–288, doi:10.1111/jomf.12423.

20. Mare, R.D. Five Decades of Educational Assortative Mating. *Am. Sociol. Rev.* **1991**, *56*, 15–32, doi:10.2307/2095670.

21. Qian, Z. Changes in Assortative Mating: The Impact of Age and Education, 1970-1990. *Demogr. Pre-2011* **1998**, *35*, 279–292.

22. Chiappori, P.-A. The Theory and Empirics of the Marriage Market. *Annu. Rev. Econ.* **2020**, *12*, 547–578, doi:10.1146/annurev-economics-012320-121610.

23. Akers, D.S. On Measuring the Marriage Squeeze. *Demography* **1967**, *4*, 907–924, doi:10.2307/2060328.

24. Schoen, R. Measuring the Tightness of a Marriage Squeeze. *Demography* **1983**, *20*, 61–78, doi:10.2307/2060901.

25. McCaa, R. Ethnic Intermarriage and Gender in New York City. *J. Interdiscip. Hist.* **1993**, *24*, 207–231, doi:10.2307/205357.

26. Choi, K.H.; Tienda, M. Marriage-Market Constraints and Mate-Selection Behavior: Racial, Ethnic, and Gender Differences in Intermarriage. *J. Marriage Fam.* **2017**, *79*, 301–317, doi:10.1111/jomf.12346.

27. Weiss, I.; Stecklov, G. Assimilation and Ethnic Marriage Squeeze in Early 20th Century America: A Gender Perspective. *Demogr. Res.* **2020**, *42*, 99–132, doi:10.4054/demres.2020.42.4.

28. Camarero, L.; Sampedro, R. ¿Por qué se van las mujeres? El continuum de movilidad como hipótesis explicativa de la masculinización rural (Why Are Women Leaving? The Mobility Continuum as an Explanation of Rural Masculinization Process). *Reis* **2008**, 73, doi:10.2307/40184907.

29. Eckhard, J.; Stauder, J. Migration and the Partner Market: How Gender-Selective Relocations Affect Regional Mating Chances in Germany. *Eur. J. Popul.* **2018**, *34*, 59–86, doi:10.1007/s10680-017-9422-8.

30. Arango, J. Becoming a Country of Immigration at the End of the Twentieth Century: The Case of Spain. In *Eldorado or Fortress? Migration in Southern Europe*; King, R., Lazaridis, G., Tsardanidis, C., Eds.; Palgrave Macmillan UK: London, 2000; pp. 253–276 ISBN 978-0-333-98252-5.

31. Recaño Valverde, J.; Domingo i Valls, A. Evolución de la distribución territorial y la movilidad geográfica de la población extranjera en España. *Pap. Demogr.* **2006**, 1.

32. González-Ferrer, A.; Obućina, O.; Cortina, C.; Castro-Martín, T. Mixed Marriages between Immigrants and Natives in Spain: The Gendered Effect of Marriage Market Constraints. *Demogr. Res.* **2018**, *39*, 1–32, doi:10.4054/DemRes.2018.39.1.

33. Vila, M.R.; Castro-Martín, T. Childbearing Patterns of Foreign Women in a New Immigration Country. *Population* **2007**, *62*, 351–379.

34. Desiderio, R. *The Impact of International Migration on Fertility: An Empirical Study*; Knomad Paper 36; World Bank Group: Washington, D.C., 2020; p. 32;.

35. Esteve, A.; López-Gay, A.; Serret, J. Revelaciones de una década de matrimonios entre españoles y extranjeros, 1998-08. *Estad. Esp.* **2013**, *55*, 213–230.

36. Cortina, A.E., Clara Demographic Research - Changes in Educational Assortative Mating in Contemporary Spain Available online: https://www.demographic-research.org/volumes/vol14/17/default.htm (accessed on 10 February 2022).

37. Esteve, A.; Bueno, X. Tras El Rastro Estadístico de Las Parejas de Inmigrantes En España. *Estad. Esp.* **2010**, *52*, 91–125.

38. Cabré, A.; Cortina Trilla, C.; Esteve, A. ¿Con quién se unen los latinoamericanos en España? Respuestas a partir de tres fuentes estadísticas. *Pap. Demogr.* **2006**, *300*, 1–21.

39. Sánchez-Domínguez, M.; Valk, H. de; Reher, D. Marriage Strategies among Immigrants in Spain. *Rev. Int. Sociol.* **2011**, *69*, 139–166, doi:10.3989/ris.2011.iM1.389.

40. Medrano, J.D.; Cortina, C.; Safranoff, A.; Castro-Martín, T. *Euromarriages* in Spain: Recent Trends and Patterns in the Context of European Integration: Euromarriages in Spain. *Popul. Space Place* **2014**, *20*, 157–176, doi:10.1002/psp.1774.

41. Rodríguez-García, D.; Luken, V. de M.; Solana-Solana, M. Las uniones mixtas y sus descendientes en España: evolución y consideraciones sobre la mixticidad. *Anu. CIDOB Inmigr.* **2020**, 167–196, doi:10.24241/AnuarioCIDOBInmi.2020.168.

42. Cortina, C.; Esteve, A.; Domingo, A. Crecimiento y singularidades demográficas de los matrimonios de extranjeros en España. *Migr. Publ. Inst. Univ. Estud. Sobre Migr.* **2006**, 75–105.

43. LeSage, J.; Pace, R.K. *Introduction to Spatial Econometrics*; Chapman and Hall/CRC: New York, 2009; ISBN 978-0-429-13808-9.

44. Anselin, L.; Rey, S.J. *Modern Spatial Econometrics in Practice: A Guide to GeoDa, GeoDaSpace and PySAL*; Chicago, IL, 2014; ISBN 978-0-9863421-0-3.

45. Lambert, D. Zero-Inflated Poisson Regression, with an Application to Defects in Manufacturing. *Technometrics* **1992**, *34*, 1–14, doi:10.2307/1269547.

46. Loeys, T.; Moerkerke, B.; De Smet, O.; Buysse, A. The Analysis of Zero-Inflated Count Data: Beyond Zero-Inflated Poisson Regression. *Br. J. Math. Stat. Psychol.* **2012**, *65*, 163–180, doi:10.1111/j.2044-8317.2011.02031.x.

47. Zero-Inflated Poisson Regression | R Data Analysis Examples Available online: https://stats.oarc.ucla.edu/r/dae/zip/ (accessed on 13 February 2022).

48. Quiroz, J.L.; Peeters, L.; Chasco, C.; Aroca, P. Equal Access to University Education in Chile? An Application Using Spatial Heckman Probit Models. *Mathematics* **2022**, *10*, 280, doi:10.3390/math10020280.

49. Anselin, L. Spatial Externalities, Spatial Multipliers, And Spatial Econometrics. *Int. Reg. Sci. Rev.* **2003**, *26*, 153–166, doi:10.1177/0160017602250972.

50. Epifani, I.; Ghiringhelli, C.; Nicolini, R. Population Distribution over Time: Modelling Local Spatial Dependence with a CAR Process. *Spat. Econ. Anal.* **2020**, *15*, 120–144, doi:10.1080/17421772.2020.1708442.

51. Halleck Vega, S.; Elhorst, J.P. The SLX Model. *J. Reg. Sci.* **2015**, *55*, 339–363, doi:10.1111/jors.12188.

52. Mella-Márquez, J.M.; Chasco, C. Urban Growth and Territorial Dynamics: A Spatial-Econometric Analysis of Spain. In *Spatial dynamics, networks and modelling*; Edward Elgar Publishing, 2006; p. 42 ISBN 978-1-78100-747-1.

53. Gourieroux, C. *Econometrics of Qualitative Dependent Variables*; Cambridge University Press: Cambridge, MA, USA, 2000; ISBN 978-0-511-80560-8.

54. Tobin, J. Estimation of Relationships for Limited Dependent Variables. *Econometrica* **1958**, *26*, 24–36, doi:10.2307/1907382.

55. Ferrari, S.; Cribari-Neto, F. Beta Regression for Modelling Rates and Proportions. *J. Appl. Stat.* **2004**, *31*, 799–815, doi:10.1080/0266476042000214501.

56. Greene, W. *Econometric Analysis*; New York, NY, 2017; ISBN 978-0-13-446136-6.

57.   Berkson, J. Tables for Use in Estimating the Normal Distribution Function by Normit Analysis: Part I. Description and Use of Tables Part II. Comparison Between Minimum Normit X2 Estimate and the Maximum Likelihood Estimate. *Biometrika* **1957**, *44*, 411–435, doi:10.2307/2332873.

58.   Chasco, C.; Aroca, P.; Anselin, L. Probit Models for Grouped-Data Migration Flows: A Theoretical Note. *Economia* **2019**, *42*, 1–8, doi:10.18800/economia.201902.001.

59.   Esteve, A.; Chasco, C.; López-Gay, A. Modeling Spatial Local Variations in Intermarriage Available online: https://b2share.eudat.eu/records/e0006de39f5b4d138fe41e6acde57cb6 (accessed on 28 February 2022).