Article

High-dimensional Probabilistic Fingerprinting in Wireless Sensor Networks based on a Multivariate Gaussian Mixture Model

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Abstract: The extensive deployment of wireless infrastructure provides a low-cost way to track mobile users in indoor environment. This paper demonstrates a prototype model of an accurate and reliable room location awareness system in a real public environment, where three typical problems arise. First, a massive number of access points (APs) can be sensed leading to a high-dimensional classification problem. Second, heterogeneous devices record different received signal strength (RSS) levels due to the variations in chip-set and antenna attenuation. Third, APs are not necessarily visible in every scanning cycle leading to missing data. This paper presents a probabilistic Wi-Fi fingerprinting method in a hidden Markov model (HMM) framework for mobile user tracking. Considering the spatial correlation of the signal strengths from multiple APs, a Multivariate Gaussian Mixture Model (MVGMM) is fitted to model the probability distribution of RSS measurements in each cell. Furthermore, the unseen property of invisible AP has been investigated in this research, and demonstrated the efficiency of differentiation between cells. The proposed system is able to achieve comparable localization performance. The filed test results present a reliable 97% localization room level accuracy of multiple mobile users in a real university campus WiFi network without any prior knowledge of the environment.

Keywords: Multivariate Gaussian Mixture Model (MVGMM); multivariate linear regression; Expectation-Maximization imputation; WiFi localization; Hidden Markov Model (HMM)

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1. Introduction

Global Positioning System (GPS) has been widely used to provide location information in outdoor environments, but it cannot give reliable positioning indoors [1]. WiFi based localization system has attracted continuous attention because of the prevalent deployment of Wireless Local Area Network (WLAN) infrastructure and the extensive availability of WiFi enabled mobile devices, which provides a potentially low-cost way to track a mobile user in a building [2]. The vast majority of current indoor localization systems are designed for sub-meter accuracy in position estimation which is unnecessary for most indoor navigation applications [3]. Room-level or region-level granularity of location is sufficient for most location aware services [4][5][6][7].

Received signal strength (RSS) based WiFi fingerprinting is a typical method used for location estimation, since it does not need any prior knowledge of access points (APs) deployment. The idea of the fingerprint technology is to use on-line RSS measurements to match the fingerprint database previously generated at every location in the off-line training phase. In the probabilistic fingerprint approach, the statistical distribution of the signal strength for each different location is built based on sample data collected during the training phase. In the on-line phase, Bayesian inference is used to calculate the probability that a user is at a certain location given a specified observation, and estimate the most likely location of the mobile device. The accuracy of the probability distribution model directly affects the final performance of the probabilistic fingerprint positioning [8].

Our previous research [9] employed a joint histogram model to generate the fingerprint probability distribution. However, in a complex and noisy open space environment, for example a university campus, an enormous number of APs can be scanned both during the survey and positioning phase. Matching a quantized histogram from 50 APs exactly almost never happens, rendering an AP selection rule is required to get reduced-dimensional quantized states for each cell [10]. The joint histogram probability method can achieve as high as 95% room level accuracy in a real university campus based on the data collected in static mode. The problem arises for the dynamical data that the number of visible APs is typically smaller in kinematic collection mode than the set of static data, which has been demonstrated in the experimental analysis in [11]. This is due to the fact that when the data is acquired in a moving motion, it is more difficult to collect extensive measurements than during training in a static collection mode. Furthermore, signals from wireless APs are variable and not all locations record signals from every AP, and the set of APs operating during the training phase is not necessarily the same as the set of APs at runtime [12]. Thus, the AP selection rule will pick the AP set which is not able to represent the characteristic of the on-line measurements and give biased estimation of the user position.

Most existed fingerprint-based algorithms ignore the APs that are no longer visible at training process and runtime. The conventional method of dealing with missing data is to set a low RSS value [7] or assign a penalty in the matching process [13][14]. While the incompleteness in the sensing data can lead to bias in the estimation of parameters, the expectation-maximization (EM) imputation strategy, derived from parametric statistics can be used to substitute for the missing data and learn the parameters from

the incomplete dataset [15]. Besides, learning from the RSS characteristics, different rooms have both different visible and invisible AP sets, which is also a signature that can be used to differentiate between 58 cells. By taking advantages of the 'unseen properties' of invisible APs, a conditional probabilistic 59 observation model is utilized to describe the likelihood of receiving a particular invisible AP set at a 60 certain location. The hypothesis is if an AP can not be scanned at all reference points (RPs) within a cell 61 during the training phase, an on-line observation contains RSS measurements from that AP occur rarely within that cell. This is expected as the APs with no RSS readings are less probable at the same location. In addition, most systems assume independence between the RSS measurements at a certain position 64 from the various APs. We have proved in this article that the correlation between the RSS measurements 65 taken within a cell is too high to be ignored. A simple explanation would be an object moves around a 66 certain position, the RSS measurements from all visible APs at that point will be affected simultaneously. 67 This paper proposes a statistical approach of localizing a mobile user with room-level accuracy 68 in an indoor wireless environment. By segmenting the indoor area into several cells, the system 69 fuses crowdsourcing RSS measurements from all visible APs collected within each cell. Different devices generally can provide different intensity readings due to many factors such as antenna gain and transmission power. The multivariate linear regression is used to address the RSS variance problem 72 in the crowdsourced training data caused by the device heterogeneity. The EM imputation strategy is 73 exploited to replace the missing RSS in the training data instead of assigning a low constant value. 74 Then a high-dimensional probabilistic fingerprint is constructed for each cell based on the multivariate 75 Gaussian mixture model (MVGMM) considering the correlations between APs. The Hidden Markov 76 Model (HMM) is applied to track the mobile user, where the hidden states comprise the possible room 77 locations and the WiFi RSS measurements are taken as observations. In the positioning phase, revealing 78 the trajectory of the user can be carried out with the Viterbi Algorithm. Besides, the information of invisibility of APs enabling the introduction of rigorously motivated trustworthiness for updating the 80 conditional likelihood observation function. 81

The remainder of the paper is organized as follows: section 2 briefly introduce the background and related work. Section 3 depicts the proposed system architecture. Section 4 presents the experimental results to verify the validity of the proposed algorithm. Section 5 draws the conclusion.

85 2. Related Work

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The vast majority of current indoor localization systems are designed for sub-meter accuracy in position estimation which is unnecessary for most indoor navigation [3]. Room-level or region-level granularity of location is sufficient for most location aware services [4][5][6][7].

Traditional WiFi fingerprinting method involves a site survey before the test, which needs to grid the area and construct a radio map associating each location. Conventional fingerprint localization algorithms normally average the WiFi RSS measurements for each AP in stored signatures. In practice, this is not consistent with RSS fluctuations, due to the multipath effects in complex indoor environments [16][17]. To get real-time correction of RSS variations and fluctuations, the Differential WiFi (DWiFi) scheme is proposed by analogy to Differential GPS (DGPS) where reference station network measurements are employed [18]. The recorded RSS measurements at user's end are corrected

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and the fingerprinting database is continuously updated to encounter for the possible changes in the dynamics of the environment.

In the probabilistic fingerprint techniques, a fingerprint is the probability distribution of the signal strength given the location instead of the mean during the training phase. Some approaches assume a Gaussian distribution of signal strength [19][20][21], which is not always true as the RSS distribution tends to be left-skewed, analysed in [22][23][24]. The Horus system infers the target location with the maximum posterior probability assuming a standard Gaussian distribution [21]. Another efficient approach to estimate the probability density distribution (PDF) is to use kernel functions [10]. Mirowski et al. extends this work by comparing the similarities between two PDFs using Kullback-Leibler divergence (KLD), and then performs localization through kernel regression [25].

Histogram-based probabilistic methods do not assume any known distribution and is closely related to discretization of continuous values to discrete ones [26][27][28]. However, histogram-based performance is primarily dependent on the choice of bin number and bin width. In addition, Zhang et al. pointed out that histogram-based approaches are only appropriate for low-dimensional datasets because the calculations in histogram-based techniques are exponential in the dimension of the dataset [29]. Therefore, this type of approach has low scalability to problems with larger numbers of data points and higher-dimensional spaces.

Using a subset of available APs enables reduces the number of variables and allows reliable low-dimensional quantized states for each room, which normally involves a sanity assessment to select a subset of APs for positioning [30]. The concept of 'important AP' is proposed to select significant APs for each location where the AP with the highest RSS is denoted as the important AP [31]. This method works properly for the static data, while the problem arises for the dynamical data. As demonstrated in [11][12], the visible AP set in kinematic collection mode is typically smaller compared with the AP set in static data. Thus, the AP selection rule will pick the AP set which is not suitable to represent the characteristic of the runtime measurements and give biased estimation of the user position. This paper builds up a high-dimensional probabilistic radio map by considering all available visible APs in the training data regardless of the signal quality. This is to ensure the inclusion of every possible AP that would be sensed in the observation data instead of assigning a constant probability.

Where there are missing RSS values from some APs at some locations, a heuristic method for handling the missing data is to set a constant minimal possible RSS value [32]. In this paper, we use the EM imputation method to replace the missing values in the incomplete data. In addition, we have observed that the missing APs also provide extra information because of their 'unseen' properties. A conditional likelihood observation function is presented by taking advantages of the invisibility of APs, referring to the likelihood of observing a particular invisible AP set. Similar work can be found in work [33], where an AP pickup probability is modelled using maximum entropy Gibbs distributions, indicating the beacon-visibility in each location. Bisman and Veloso neglected any unobserved or extra APs when apply the Gaussian kernel to compare different signal strengths [34]. Penalties are applying for the missing APs proposed in [35]. The concept of penalties is also used by the Redpin algorithm [36] where an extra bonus weight is added for common APs and an extra penalty for non-common APs.

Luo [37] suggested that the standard Gaussian distribution did not fully describe the signal strength in the indoor environment, a more suitable fit in the probability distribution model of signal strength is

based on the Gaussian Mixture Model (GMM), which infers an approximate probability distribution by a weighted mixture of Gaussian densities [38]. The WiGEM system employed the GMM to learn the signal propagation parameters for each AP [39]. The GMM is applied to model the probability distribution of the signal strength for each AP, assuming that the APs are independent at a particular position [40]. GMM is used to identify the RSS components of multipath decline separated from the line-of-sight (LOS) component in [41]. However, all the above systems mentioned above do not take into account the interdependencies among the RSS measurements from the various APs. Thus the system proposed a multivariate GMM by capturing the correlation between RSS measurements from pairs of APs.

3. System Overview

The proposed system is implemented to achieve high room level localization accuracy and reliability. To this end, six steps need to be taken to identify the trajectory of the mobile user, as presented in figure 1. The first step is to segment the indoor area into several cells and randomly assign multiple RPs within each cell. Then the training data collection is carried out by fusing RSS measurements taken at all RPs within each cell by all contributed devices. A multivariate linear regression is conducted to calibrate the RSS measurements collected from different devices. The missing RSS values are replaced by the new data estimated by the EM imputation method. The fifth step is to exploit the MVGMM to construct the probabilistic radio map for each cell based on the calibrated training data. Lastly an online matching process is performed which is to fit the runtime observation into the distribution model of each cell, and feed into the Viterbi Algorithm to backtrack the trajectory of the user.

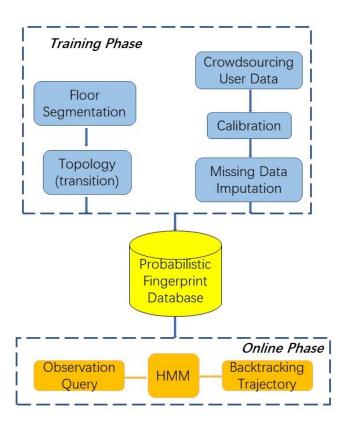


Figure 1. Framework of The System

3.1. Building Topology

Room level localization is defined in terms of cell-based localization, i.e. locations are represented as cells. A cell may correspond to a room, or a section of a hallway. In the test area depicted in figure 2, for instance, the main corridor is divided into four cells. In addition, the segmentation rule classifies the floor area into 3 categories: rooms, corridors and entrance/ exits. It constructs logical links between rooms and corridors and models the constraints to movement imposed by the building's layout. Note that surveying the scales or true dimensions of the floor is not needed for space segmentation.

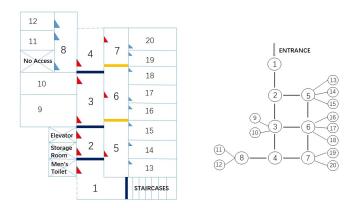


Figure 2. A schematic of the Bolz Hall, Ohio State University (not to scale) and Topology

The segmentation will define the transition matrix in the HMM such that only adjacent cells have non-zero transition probability while the transition probability between isolated cells is zero. The system does not attempt to determine the exact grid position of the mobile user but the cell that the user is in.

3.2. Cell Training Data

After cell segmentation, a WiFi database is created for each cell using the signal strength measurements collected during the training phase. Cell training data, which involves the RSS from all visible APs intensively sampled at multiple RPs within each cell and each RP is associated with manually labelled cell ID. The RPs are randomly selected within each cell and their locations need not to be known.

Given a building with a set of cells R, and the total number of visible WiFi APs is N. For a given cell $r \in R$, a WiFi measurement is a vector containing signal strength from N APs, denoted as:

$$S_{(r,j)} = \{AP_1 : Rss_{1,j}, AP_2 : Rss_{2,j}, \dots, AP_N : Rss_{N,j}\}, j = 1, \dots, M$$
(1)

M is the total number of measurements at cell r and could vary by rooms. Each AP is identified by its unique MAC address and $Rss_{i,j}$ is the signal strength value from AP_i in the j_{th} measurement. Note the RSS value is replaced with NaN for the AP unobserved in one measurement.

During the offline phase, the signal strength from all visible APs are intensively sampled at multiple RPs within each cell. The training data for cell r fused from all RPs will be stored in a $M \times N$ matrix denoted by $S_r = \{S_{(r,j)} | j = 1, \dots, M\}$.

3.2.1. Calibration

In this paper, we fused crowdsourcing training data collected by multiple devices, which is the most promising solution for reducing the site survey labour consumption [42][43]. Most existed localization systems assume that the device contributed for the training data collection is the same in positioning phase. While the fact is that every mobile user may become a potential contributor for the fingerprint database construction and the participated devices are usually different, which causes new challenges pertaining to cross-device fingerprint database construction. In addition, different devices have different sensor specifications and varying readings even at the same locations [44], a calibration process is essential in the crowdsourced radio map construction by fusing the RSS radio maps from different devices.

In order to support different devices and make the fingerprints of diverse devices compatible with each other, the calibration step is performed prior to the positioning phase. The relation in RSS values between two different devices at the same location appears to be linear, as discussed in [45]. In this case, device calibration is conducted by means of data fitting methods that create a linear transformation from the new device to the reference device. The adjusted RSS are then fused as crowdsourced training data.

In this paper, we implement the multivariate linear regression model [46] to match the signal strengths measured by the new device with the radio map constructed by the reference device. The calibration data collection is simple, the user carries the reference device and the devices that need to be calibrated, walks freely inside the area of interest, collects data from all visible APs in the environment at the same time. Given a linear mapping with parameters $a_{mvregress}$ and $b_{mvregress}$, the signal strength values reported by client X are mapped to the RSS values reported by client Y. The linear regression model is expressed as:

$$DeviceY_{RSSintensity} = a_{mvregress} \times DeviceX_{RSSintensity} + b_{mvregress}$$
 (2)

Where $DeviceX_{RSSintensity}$ is the RSS readings from device X that needs to be calibrated, $DeviceY_{RSSintensity}$ is the RSS readings from the reference device Y. $a_{mvregress}$, $b_{mvregress}$ are the calibration coefficients calculated by the linear regression model.

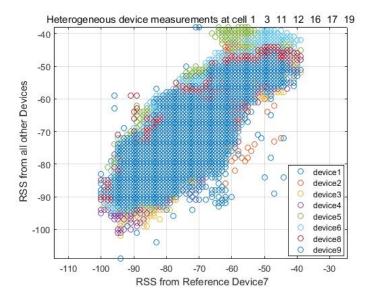


Figure 3. Heterogeneous device measurements

In order to do the linear fit between the two signal strength intensity, the user carries all devices freely walking around the test area and collects the RSS at the same time. From figure 3 we can see the signal strength from all other devices almost follows a linear match with the reference signal strength intensity at every location. Figure 4 shows a calibration example between device 1 and device 7, where device 7 is used as the reference device. The device specification is described in table 1 in the following experimental section.

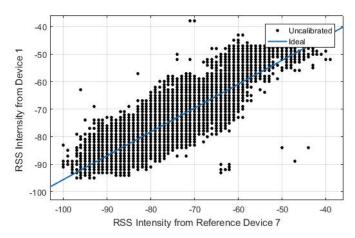


Figure 4. Multivariate Linear Regression between Device 1 and Device 7

3.2.2. Missing data Imputation

In a real open space environment, many APs can be scanned, which leads to a high-dimensional fingerprint database. Due to the RSS variability, APs may not be visible in every scan, leading to missing data. Missing data can reduce the statistical power of an investigation and produce biased estimates, leading to invalid conclusions [47]. Before we apply the MVGMM to estimate probabilistic fingerprint distributions for each cell, we need to handle the missing data problem.

In case of the missing RSS values from some APs in RSS measurement vectors, the simplest way for imputing missing values is to set a constant, the lowest possible reading or the mean RSS value of each AP [48], however this will alter the shape of the distribution and bias the covariance. A scenario may arise when RSS values from certain APs are available in the survey phase but are not observed in the on-line stage. A common approach is to find the effective APs which are visible in both the training and positioning phase [49]. The system in [50] created RSS reference surfaces for each AP using the Support Vector Regression (SVR) Machine to infer the missing data. Then, during the localization stage, the measured RSS from each AP will be searched in the corresponding surface. In [51], a multilayer perceptron (MLP) artificial neural network (ANN) with fingerprinting approach has been investigated to handle the problem of missing APs in online matching stage. All the aforementioned approaches neglect the spatial correlation to simplify generation of theoretical RSS datasets for each missing APs in the offline phase, which will result in poorer localization performance.

In this paper, we implement the expectation maximization (EM) algorithm for incomplete data parameter estimation, assuming the missing data mechanism under the missing at random (MAR) assumption. Detailed description of the algorithm can be found in [52][53][54].

The EM algorithm is an iterative process that finds the maximum likelihood estimation (MLE) of the parameters until they converge in the presence of missing data. In general, the E (expectation) step calculates the expectation of the log-likelihood function given the observed data. The M (maximization) step is to update the new parameters that maximize the expected log-likelihood from the E step. Suppose the complete data set Y is partitioned into $Y = (Y_{obs}, Y_{miss})$, where Y_{obs} represents the observed part of Y, Y_{miss} is the missing part. The unknown parameter model θ of Y can be written as:

$$P(Y|\theta) = P(Y_{obs}, Y_{miss}|\theta) = P(Y_{obs}|\theta)P(Y_{miss}|Y_{obs}, \theta)$$
(3)

Given an initial guess of $\theta^{(t)}$, it is possible to calculate the distribution of the missing data $P(Y_{miss}|Y_{obs},|\theta^{(t)})$. The E step is to calculate the expected complete data log-likelihood ratio $Q(\theta|\theta^{(t)})$ with respect to the imputation model of missing data.

$$Q(\theta|\theta^{(t)}) = \int \log[p(Y_{obs}, Y_{miss}|\theta)]P(Y_{miss}|Y_{obs}, \theta^{(t)})dY_{miss}$$
(4)

The M step maximize $Q(\theta|\theta^t)$ from the previous E step:

$$Q(\theta^{(t+1)}) = \arg\max Q(\theta|\theta^{(t)}) \tag{5}$$

3.3. Probabilistic Fingerprint

During the offline phase, a probability density function for each cell is estimated based on the MVGMM. Most current work that exploited GMM to estimate the probability distribution tends to specify a fixed mixture component, while it is important to note that the mixture component is a variable that is acting together to determine the overall estimation. The Akaike's information criterion (AIC) measures the goodness of fit of statistical models [55] and it is applied in this paper to find the optimal number of components *K* of MVGMM. The authors decided that 7 mixture components should be used in terms of optimum classification results and computation burden, as presented in figure 8 in the experimental section. Then, an online matching process fits the online observation with the optimal parameters calculated during the offline process, to identify the probability the observation belongs to each cell.

3.3.1. Multivariate Gaussian Mixture Model

The probabilistic fingerprint is the conditional probability distribution of signal strengths given the cell position $P(S_r|r)$, $r \in R$. The assumption of Gaussian distribution of the RSS is not accurate enough as proved by Kaemarungsi and Krishnamurthy [56]. A Gaussian Mixture Model allows approximation of a probability density function by a weighted sum of Gaussian densities each with different parameters. Most research work uses the GMM to approximate the RSS distribution for a single AP and ignores the interference between signals from different APs [57],[58][59]. The MVGMM is implemented to approximate the probability density distribution of the training data for each cell, which takes advantages

of correlation between the RSS from various APs within a certain area.

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Given the training data S at cell r contains M RSS measurements from N APs, (for convenience, we remove the notation r in S_r in the subsequent sections), and consider one measurement contains signals coming from N APs, the training data S is a matrix consists of multivariate random variables. The density function modelled by MVGMM can be mathematically defined as:

$$P(S|\mu, \Sigma, \pi, r) = \sum_{k=1}^{K} \pi_k \mathcal{N}(S|\mu_k, \Sigma_k)$$
(6)

Where K is the number of component of the model, $\sum_{k=1}^{K} \pi_k = 1$. π_k, μ_k, Σ_k are the mixture weight, mean and covariance matrices for the k_{th} mixture component. $\mathcal{N}(S|\mu_k, \Sigma_k)$ is the k_{th} mixture component from N-dimensional multivariate Gaussian distribution:

$$\mathcal{N}(S|\mu_k, \Sigma_k) = \frac{1}{2\pi^{N/2} |\Sigma_k|^{1/2}} \exp^{-\frac{1}{2}(S-\mu_k)^T \Sigma_k^{-1}(S-\mu_k)}$$
(7)

During fingerprinting, the signature of cell r is generated by the MVGMM parameterized by $\Phi = \{\mu_k, \Sigma_k, \pi_k\}, k = 1, \dots, K$. The EM algorithm is applied to estimate the parameters of the model.

1) E step. Calculate the responsibilities using the current parameters, which can be viewed as the posterior probability that the m_{th} measurement S_m is from the k_{th} component.

$$\gamma_k(S_m) = \frac{\pi_k \mathcal{N}(S_m | \mu_k, \Sigma_k)}{\sum\limits_{j=1}^K \pi_j \mathcal{N}(S_m | \mu_j, \Sigma_j)}$$
(8)

2) M step. Re-estimate the parameters using the responsibilities from the E step.

$$\mu_k^* = \frac{1}{M_k} \sum_{m=1}^{M} \gamma_k(S_m) S_m \tag{9}$$

$$\Sigma_k^* = \frac{1}{M_k} \sum_{m=1}^M \gamma_k(S_m) (S_m - \mu_k^*) (S_m - \mu_k^*)^T$$
(10)

$$\pi_k^* = \frac{M_k}{M} \tag{11}$$

where

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$$M_k = \sum_{m=1}^{M} \gamma_k(S_m) \tag{12}$$

3) Evaluation. Evaluate the log-likelihood

$$lnP(S|\mu, \Sigma, \pi) = \sum_{m=1}^{M} ln\{\sum_{k=1}^{K} \pi_k \mathcal{N}(S_m|\mu_k, \Sigma_k)\}$$
(13)

These three steps are iteratively repeated until the log likelihood convergences.

3.3.2. Conditional Probabilistic Observation Model

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Most work ignores the unobserved RSS value in the runtime observation, while we investigated that the missing APs also provide extra beneficial information in deciding the user position because of their 'unseen' properties. This paper presents a conditional probabilistic likelihood observation function, by taking advantages of the invisibility of APs, referring to the likelihood of observing a particular invisible AP set.

The hypothesis is that if an AP can not be scanned for the whole training data collection within cell r, then an online observation contains RSS value from that AP would have low probability belongs to that cell. In other words, if an observation contains RSS values from the APs that should not to be seen in cell r, then the probability of being located in cell r given the observation would be lower. This is expected as the APs with no RSS readings are less probable to be heard within the same area.

By splitting the observation O into O_{RSS} , the RSS measurements for the visible APs and $O_Z(I)$, a binary indicator variable for APs where $O_Z(i) = 1$ if AP i is invisible and 0 otherwise. In this case, we define the *particular invisible AP set I* as:

$$I = \cap (Invisible \ APs \ r, Visible \ APs \ O) \tag{14}$$

The observation probability would be:

$$P(O|r) = P(O_{RSS}, O_Z(I)|r)$$
(15)

$$= P(O_{RSS}|O_Z(I),r)P(O_Z(I)|r)$$
(16)

 $P(O_{RSS}|O_Z(I),r)$ is matching the online observation with the probabilistic fingerprint that discussed 274 in the MVGMM section 3.3.1. $P(O_Z(I)|r)$ is a likelihood of observing RSS from the invisible AP sets 275 of cell r. 276

$$P(O_Z(I)|r) = \prod_{i=1}^{P} P(O_Z(i)|r)$$

$$P(O_Z(i)|r) = \frac{O_Z^r(i)}{\sum_{r \in R} O_Z^r(i)}$$
(18)

$$P(O_Z(i)|r) = \frac{O_Z^r(i)}{\sum_{x \in P} O_Z^r(i)}$$
 (18)

Where P = |I|. $O_Z^r(i)$ is the invisibility of AP i at cell r, and $\sum_{x \in P} O_Z^r(i)$ is the invisibility of AP i over 277 all cells R. 278

3.4. Hidden Markov Model

The motion of the user can be modelled as a Markov process [60] and a HMM is applied to track the mobile user, where the hidden states comprise the possible cell locations and the RSS measurements are taken as observations.

The formal definition of a HMM is as follows. The set of states are identical to the set of cells. Let 283 S_1, S_2, \ldots, S_T be the sequence of hidden states in the state set R during a time sequence $t = 1, \ldots, T$, 284 which constitutes the user moving trajectory. The observed WiFi RSS sequence $O = O_1, O_2, \dots, O_T$ up 285 to time T in correspond. The model is characterized by parameters $\lambda = \{A, B, \alpha\}$. 286

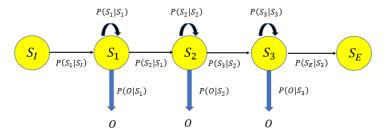


Figure 5. HMM

Given an observation sequence, the Viterbi algorithm determines the most probable hidden state sequence.

$$P(O|\lambda) \simeq \max_{S_1, S_2, \dots, S_T} P(O_1, O_2, \dots, O_T, S_1, S_2, \dots, S_T | \lambda)$$
(19)

A is the transition probability matrix. The segmentation rule based on the building topology is encoded in the state transition probability, which is the probability of the user moving from *cell i* to *cell j*, denoted as $p_{i,j} = P(S_{t+1} = S_j | S_t = S_i)$. If a given cell is linked to *n* other cells (including itself), then the probability of moving to one of these cells is defined to be 1/n, the probability of moving to other isolated cells is 0. Here we use equal probability for simplicity.

B is the emission probability, i.e. the likelihood of producing observation O_t from cell S_j , which is to fit the observation to the signature of each cell calculated by the MVGMM and the conditional observation likelihood, referring to equation (6), (15) and (17):

$$b_j(O_t) = P(O_t|S_t = S_j) = \sum_{k=1}^K \pi_{k,j} \mathcal{N}(O_t|\mu_{k,j}, \Sigma_{k,j}) \times \prod_{i=1}^P P(O_Z(i)|S_j)$$
 (20)

Where $\Phi_j = \{\mu_{k,j}, \Sigma_{k,j}, \pi_{k,j}\}, k = 1, \dots, K$ is the mixture parameters associated with *cell* S_j . α is the prior state probability, here we assign equal prior probability to each state.

296 4. Experimental Results

To verify the proposed approach, a field test was carried out on level 2, the Bolz Hall, Civil & Environmental Engineering building, Ohio State University, United States. The geometry of the building consists of labs, offices and classrooms, shown in figure 2. We divided the floor plan into 20 cells on topology. Typically there is one cell per room. We also segmented the two long hallways into cells, which are cell 1-4 and cell 5-7 denoted in figure 2. The first hallway connects the entrance of the building to the test area and the second corridor connects the right hand side 8 administration offices (cell 13-20) to the main hallway. The whole training data collection took place for 5 days covering different times of the day. During the collection, students and staffs walked around normally as usual.

In this section, we have analysed the correlation between the RSS measurements in the training data for each cell, and presented the efficiency of the proposed localization system for both stop & go movement and dynamic walking data. The minimum training size and the affect of different *K* mixture component have been investigated in order to attain certain room level localization accuracy. A

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comparison is carried out between the mean RSS and the EM imputation method in terms of replacing the missing values in the training data.

Device ID Brand Average time for one scan (seconds) 1 2.87 Samsung S8 2 Samsung Galaxy A6 3.565 3 Samsung S3 3.665 4 Samsung S3 3.46 5 Google Pixel 3.50 6 Moto G3 0.575 7 Huawei Mate 7 2.55 8 Oneplus 1 3.06 9 LG G4 3.355

Table 1. Device Specification

In the filed test, 9 android devices were used for the crowdsourcing data collection, see table 1. For WiFi data collection, the CPS App developed by Mr Hofer was used [61], each WiFi scan records the timestamp, location ID, the MAC address, the network name and the RSS values for all the visible APs. The devices collected signals from the university public base stations about which we had no prior information. The data collection consisted of 3 stages: calibration, static training and real kinematic walking data collection.

Table 2. Device Calibration Coefficients

Device ID	M	C
1	0.9336	-6.7754
2	0.8825	-11.4863
3	0.8951	-8.7945
4	0.8600	-11.9044
5	0.7802	-21.3151
6	0.8709	-13.8607
8	0.9706	-5.3788
9	0.8701	-10.6396

The calibration data collection was conducted to get the coefficients for each device with respect to the reference device, see section 3.2.1. The 9 devices were put on a trolley and one user pushed the trolley around the test area and stopped at random in various cells. Each device collected 200 scans at each location. In this paper we use device 7 as the reference device, and the calibration coefficients

for the other 7 devices are displayed in table 2. The calibrated RSS measurements for each device is calculated as:

$$\widetilde{RSS_d} = M_d \times RSS_d + C_d \tag{21}$$

Where RSS_d is the raw RSS measurements taken by device d, M_d , C_d are the calibration coefficients and $\widetilde{RSS_d}$ is the calibrated RSS measurements for device d.

After segmentation, each device was randomly assigned multiple RPs (normally 5-10) within each cell to collect training data in static mode. Note that the locations of the RPs were physically different against each device. Then the training data for each cell is obtained by fusing the data collected at all the RPs from every available device. The locations of the RPs need not to be known, and every device generally chose different times to enter into the cell to assure that the training samples are covering the whole space and time variant features. At each RP, each device was designed to collect 200-400 scans in static mode, each scan records the timestamps, point ID, the MAC address, the network name and the RSS values for all the visible APs in the environment.

4.1. AP Interdependency Analysis

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To verify the interdependence of RSS from various APs, we compare the static data collected at a single point by one device with the training data of the corresponding cell collected by the same device at multiple RPs. Intuitively, we chose 3 strongest APs (AP 19,57,74) for the two data sets to do the analysis.

RSS Properties RP1 (device 6) Cell 11 (device 6) Number of scan 400 4000 Number of visible APs 21 23 Mean RSS of AP19 (dBm) -66.90 -65.42Mean RSS of AP57 (dBm) -78.94 -78.95 Mean RSS of AP74 (dBm) -69.33-65.08 standard deviation of AP19 (dBm) 1.82 3.29 standard deviation of AP57 (dBm) 0.84 3.05 standard deviation of AP74 (dBm) 1.87 3.27 Correlation (AP19, AP57) 0.0860.21Correlation (AP19, AP74) -0.270.50 Correlation (AP57, AP74) -0.180.37

Table 3. Correlation between APs (Single Point VS Cell)

From the results shown in table 3, the RSS properties of the three APs are quite distinct at a single point from within a room. Regarding the point data, the RSS values are more stable and have smaller variations compared with the cell data. The correlations between pairs of APs at RP1 are as small as 0.086 which is similar to the results given in [56], while the correlation can also be as large as -0.27 in complex, noisy and non-line-of-sight signals. In the cell data, the correlations between pairs of APs become so large that we can no longer assume the RSS samples from the visible APs are independent,

which also explains why the proposed algorithm consider the correlations from RSS measurements between pairs of APs.

340 *4.2. AP Density*

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The exact number of training samples for each cell is displayed in table 4. Cell 1 had the most visible APs with 120, while cell 11 had the least visible APs with 34. The visibility of each AP means the number of observation from the AP respect to the total number of measurements. These 120 visible APs will be *registered* in the training data and used to extract the invisible AP sets for each cell. Figure 6 gives the example of the AP intensity and the missing data percentage for cell 1 and 11. We can clearly see that different cells have both distinct visible and invisible AP sets. The missing data percentage can be as high as 98.87% for cell 1 and 99.92% for cell 11 from AP 109 and AP 78 respectively. We manually removed APs with less than 1% visibility before applying the EM imputation to avoid singular covariance.

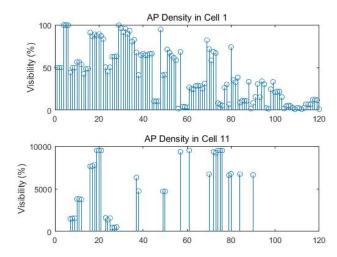


Figure 6. AP Density and Missing Data Percentage at Cell 1 and Cell 11

Table 4. Training Sample Size and Visible APs for Each Cell

Cell IDs	Training Data		Cell IDs	Training Data	
Cell IDS	Sample Size	Visible AP Number	Cell IDs	Sample Size	Visible AP Number
1	12000	120	11	10000	34
2	12000	94	12	12000	38
3	11001	85	13	10100	62
4	12004	67	14	4000	58
5	8000	82	15	5473	42
6	11800	61	16	5500	42
7	12000	61	17	7368	45
8	12103	54	18	4500	53
9	11000	55	19	4499	49
10	12000	56	20	4800	47

4.3. Stop & Go Localization Accuracy

We chose 500 scans randomly out of the crowdscourced training samples for each cell and exclude them from the training data. The remaining set was used to train the MVGMM model and get the probabilistic fingerprint for each cell. The 500 scans we removed from the training data were formed as the test set for each cell. Note the 500 test scans can be from any test device.

We constructed the stop & go movements by including transition between cells, the observed RSS sequences were simulated by randomly choosing 50 scans from the 500 test samples of each cell.

In the stop & go tests, 9 different trajectories were designed to verify the proposed algorithm. The first 6 trajectories were designed to move only between adjacent cells, covering different parts of the test area. While trajectory 7 was designed to repeat trajectory 4 but miss transition data at 3 cells. Trajectory 8 was designed to repeat trajectory 5 but miss more transition data at 6 cells. Trajectory 9 repeated trajectory 6 with 8 cells' transition data missing. These latter 3 trajectories were selected to simulate the scenarios that continuous RSS measurements cannot be obtained for a period of time during the transition between cells. This is reasonable since one WiFi scan can take around 3 seconds for some devices, while the user has already passed the transition cell.

Trajectory	Number of Covered Cells	Acc with Conditional Observation Likelihood
1	16	96.53%
2	17	98.16%
3	21	98.04%
4	23	97.02%
5	25	97.05%
6	35	97.50%
7	20	97.20%
8	19	96.99%
9	27	97.16%
Average Accuracy	23	97.29%

Table 5. Matching Accuracy for Stop & Go Trajectories

In the following analysis, the number of component K was set to K=7. Table 5 gives the average matching accuracy for the designed 9 trajectories. We performed the experiments 50 times for each trajectory, randomly choosing different test samples each time. The proposed system can still work properly when the system failed to get updated observation data for a certain time, referred to the tracking results of trajectory 7-9. The accuracy decreased if the observation data in the transition cell is missing, however, the HMM based algorithm can still recover from the losing track of position with an average matching accuracy of 97.11%. Matching accuracy is defined as the percentage of the cells correctly determined:

$$Accuracy = \frac{\sum_{t=1}^{T} Equal(s_t^{HMM}, s_t^{True})}{|T|}$$
(22)

Where

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$$Equal(a,b) = \begin{cases} 1 & a = b \\ 0 & otherwise \end{cases}$$

With the aid of the proposed conditional likelihood observation function that utilizes the information of the invisible APs, the system can achieve an average of 97.29% matching accuracy even the observed data is not continuous. Figure 7 demonstrated efficiency of the distinct invisible AP set of different cells being a significant signature which helps to improves the localization performance from an average 92.98% matching accuracy to an overall 97.29% matching accuracy.

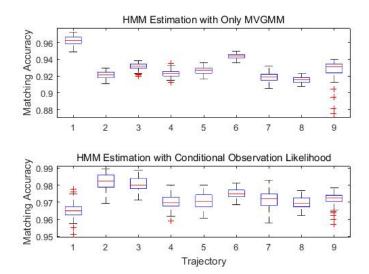


Figure 7. Conditional Likelihood Function Contributes to the Improvement of Localization Performance

4.4. Kinematic Tracking Accuracy

In addition to the stop & go simulated movement, we also conducted dynamic experiments on some devices to track a moving agent that freely moves around with normal walking speed. The user was asked to press the 'checkpoint' button in the CPS App to record the timestamps every time he entered into a new cell. Both the checkpoint time and WiFi scan time use the same nano time of the android system.

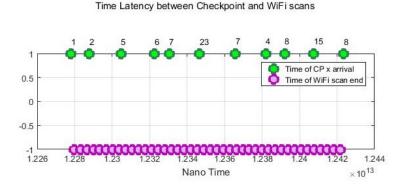


Figure 8. Time Latency between Checkpoints and WiFi Scans

Figure 8 gives an example of the recorded WiFi scan time sequences of device 4. The average time difference between the checkpoint time and the WiFi scans is 1.22 second. From figure 8 we can see that there are some latencies between the checkpoint timestamps and the WiFi scans since one WiFi scan can take between 0.6-3.7 seconds depending on devices. Considering that a WiFi scan can take few seconds and a user can change the position while scanning is done. So while a user is moving across cells, there will be always a blur in the WiFi scan and the exact cell ID.

Two kinematic trajectories were designed and repeated by different users. Each trajectory was defined as the sequence of the cell IDs along with the movement. Device 1, 2, 3, 6 and 7 repeated the trajectory 1, while device 4 and 9 repeated the trajectory 2. 7 users carried the devices starting from the same cell (normally started from cell 1), and repeated each trajectory for several times. Note that the real walking trajectory can be different as the user can walk into different locations within each cell. Here we only show part of the results due to the limited space.

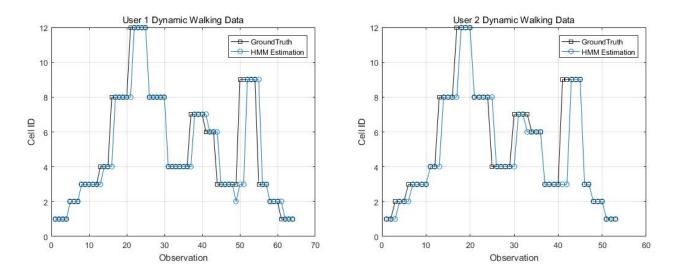


Figure 9. Kinematic Trajectory 1 Repeated by Device 1 and 2

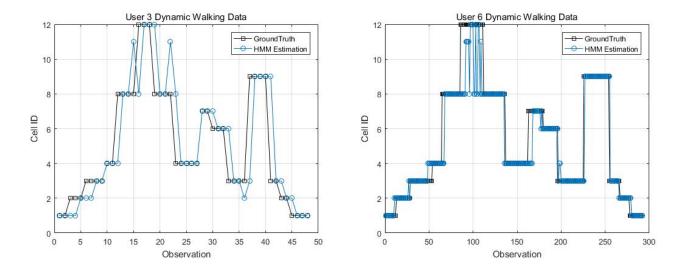


Figure 10. Kinematic Trajectory 1 Repeated by Device 3 and 6

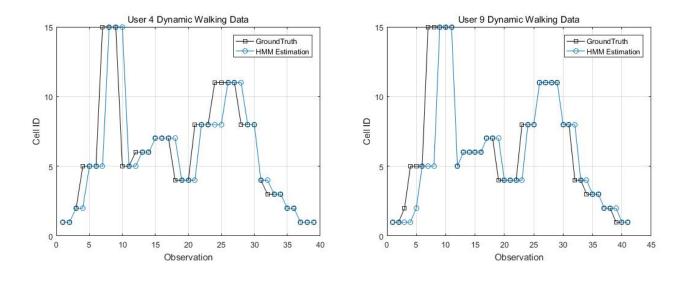


Figure 11. Kinematic Trajectory 2 Repeated by Device 4 and 9

4.5. Training Size

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Collecting enough data for creating location statistical fingerprints is the key to achieving good performance. As pointed out by Zhou [62], for a grid localization system, 5 to 6 APs deployed strategically within the test area would be ideal and each location should have enough calibration samples (e.g. 200 to 300 samples). To evaluate the performance of the proposed system with smaller training samples, we chose different training sizes ranging from 10% to 100% of the collected training data.

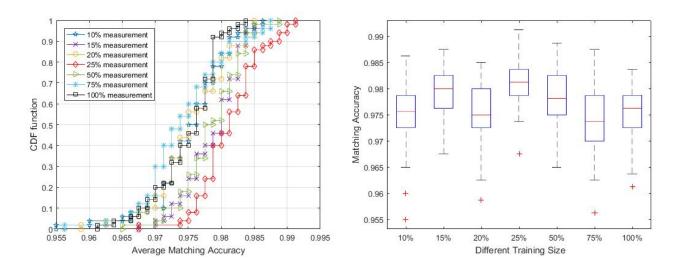


Figure 12. Matching Accuracy in Dependence of Training Size

The plot in figure 12 shows that with 25% measurements, the method can achieve the best performance with 98% accuracy in over 50% of the trials. Generally speaking, all the sample sizes are enough to train the MVGMM and can get over 97% matching accuracy for half of the trials. However, we also notice from the figure 12 that the proposed algorithm is insensitive to the size of the training samples, even presenting more robust localization accuracy to lower sample sizes. This result is similar to

the analysis in the work of Zhou [62] that denser calibration locations and more samples to the area may be confused with other areas. Elnahrawy et al. [63] also pointed out that given larger training samples, it is unlikely that additional sampling will increase accuracy. The possible reason behind might be that larger training data contains more time-varying features and signal interference from the environment. We observe that approximately 15%-25% of training data per cell is sufficient to attain comparable level of accuracy.

4.6. Optimal K

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In order to understand the optimal K, we have pre-defined different thresholds when applying the AIC rule. Figure 13 below presents a plot of the average matching accuracy of trajectory 6 when changing different K values between 6 to 30 based on 25% training data, each K runs for 50 times. Choosing a larger K will increase the accuracy at some extent with the cost of adding computation complexity. Thus, we set K = 7 for computation simplicity purposes while maintaining reasonable localization accuracy.

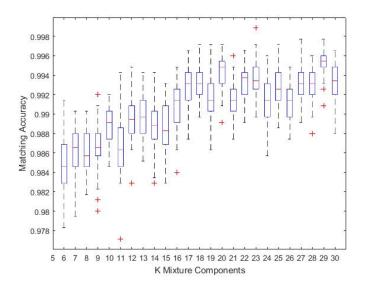


Figure 13. Optimal K in Dependence of Matching Accuracy for Trajectory 6

4.7. Comparison with Mean RSS Imputation

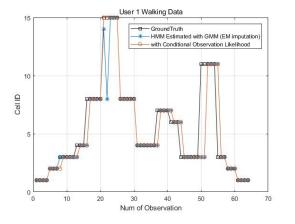
The proposed system applies the EM imputation method to deal with the missing data in the training set. This section explores the performances of the two missing data imputation methods. The only difference is the missing data in the training samples are replaced with the average RSS value for the corresponding AP.

Table 6 gives good averaged results even using the mean RSS imputation for the stop & go movements, although the accuracy is always worse than the one with the EM imputation.

Trajectory	Mismatch with Mean RSS Imputation	Mismatch with EM Imputation
1	5.10%	2.51%
2	3.53%	1.84%
3	3.68%	1.96%
4	4.05%	2.98%
5	3.85%	2.95%
6	3.29%	2.50%
7	3.96%	2.80%
8	4.32%	3.01%
9	3.94%	2.84%

Table 6. Mean RSS Imputation VS. EM Imputation

Figure 14 to 17 displayed the comparison results of the kinematic walking data. The mean imputation can still maintain good accuracy as the estimated trajectory almost matches with the ground truth with some latencies, though they normally have larger bias estimation than the EM imputation ones. In addition, figures below clearly demonstrated the efficiency of the proposed conditional likelihood function which can help to distinguish the adjacent cells and correct the position.



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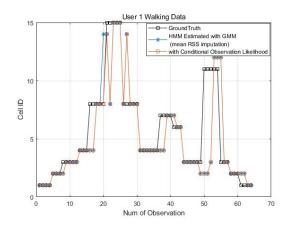
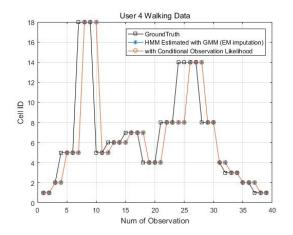


Figure 14. EM Imputation VS. Mean RSS Imputation: User 1 Dynamic Walking Data



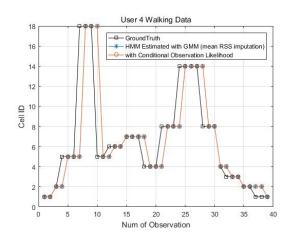
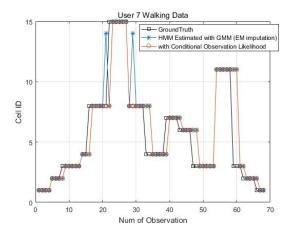


Figure 15. EM Imputation VS. Mean RSS Imputation: User 4 Dynamic Walking Data



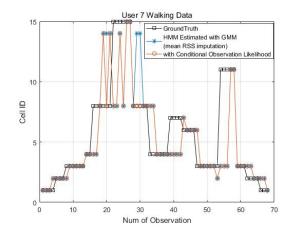
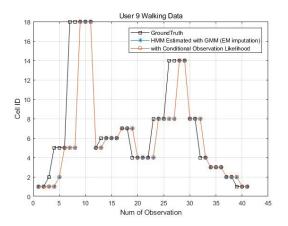


Figure 16. EM Imputation VS. Mean RSS Imputation: User 7 Dynamic Walking Data



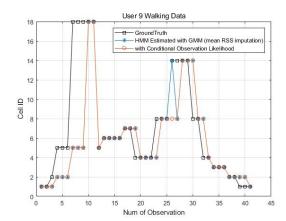


Figure 17. EM Imputation VS. Mean RSS Imputation: User 9 Dynamic Walking Data

5. Discussion

In this paper, we have validated the efficiency of the proposed conditional likelihood observation function. It correctly identify the user's position in most cases. However in some cases when the set of invisible APs for one cell is a subset of the invisible APs for another cell, no performance increase is observed.

There are some approaches dealing with the problem of the GMM parameter estimation based on incomplete data directly instead of replacing the missing ones before training [64]. The EM imputation is based on the assumption that the distribution is multivariate Gaussian, which still gives reasonable results as presented in the paper. The implementation of such algorithm would be one of the future interest to quantify the improvement in the context of the current work.

Localization accuracy in dependence of increasing training sample size is commonly discussed in literatures. While we have a different observation based on the campus wireless data verification results. The possible reason may be the crowdsourcing training data contains large variations and interference from other signal channels in the campus wireless network. In order to avoid over-training, the Gamma Test [65] will be applied to identify the optimal training data size preventing the performance from degenerating.

Cell 1 is a transition cell between the indoor and outdoor environment, which shows special characteristics in correspondence. At cell 1, the system can see 120 maximum visible APs. The property that many APs are only visible at cell 1 but invisible at all the other cells, can be used to do analysis for the transition data between indoor and outdoor.

443 **6. Conclusions**

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In this paper, we propose a statistical approach to localize the mobile user to room level accuracy based on university wireless network. The users have no basic knowledge about the base stations deployed within the environment in advance. The MVGMM is efficient to approximate the RSS distribution for each cell that takes the signals correlations into computation. The system can get an reliable 92.98% matching accuracy for half of the trials based on the crowdsourcing data.

The performance can be improved to 97.29% by introducing the conditional likelihood observation function, which takes advantages of the *unseen* signatures of APs. Instead of ignoring the invisible APs which are unobserved in the training data or the new observation, this paper investigated a conditional likelihood observation model calculated at each cell for all APs inclusive of the invisible ones, referring to a likelihood of observing an AP that is not supposed to be visible.

The proposed system demonstrates a practical prototype model of a reliable room location awareness system in a real public environment. It can handle the data uploaded by diverse devices and the noisy environment which can be widely applied in any potential public spots like guiding customers in a shopping mall or monitoring patients in a hospital. The system can be applied to a wide range of localization applications in a practical indoor environment regardless of the quality of the signals, the number of the APs, the heterogeneous devices, the interference from other channels, the time-varying phenomena or the complexity of the environment.

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464 Author Contributions

Yan Li designed and performed the experiments and wrote the paper. Bill Moran, Simon Williams, Allison Kealy and Guenther Retscher assisted with conceiving of the idea and proofreading of the paper.

467 Conflicts of Interest

The authors declare no conflict of interest.

469 References

1. Hightower, J.; Borriello, G. Location systems for ubiquitous computing. *Computer* **2001**, *34*, 57-66.

- 2. Mao, G. Localization Algorithms and Strategies for Wireless Sensor Networks: Monitoring and Surveillance Techniques for Target Tracking: Monitoring and Surveillance Techniques for Target Tracking. In *IGI Global* **2009**.
- 3. Pritt N. Indoor location with WiFi fingerprinting. *Applied Imagery Pattern Recognition Workshop* (AIPR): Sensing for Control and Augmentation **2013**, 1-8.
- 4. Chen, Y., Lymberopoulos, D., Liu, J.; Priyantha, B. FM-based indoor localization. *Proceedings of the 10th international conference on Mobile systems, applications, and services* **2012**, 169-182.
- 5. Jiang, Y.; Xiang, Y.; Pan, X.; Li, K.; Lv, Q.; Dick, R.P.; Shang, L.; Hannigan, M. Hallway based automatic indoor floorplan construction using room fingerprints. *Proceedings of the 2013 ACM international joint conference on Pervasive and ubiquitous computing* **2013**, 315-324.
- 6. Castro, P.; Chiu, P.; Kremenek, T.; Muntz, R. A probabilistic room location service for wireless networked environments. *International conference on ubiquitous computing* **2001**, 18-34.
- 7. Jiang, Y.; Pan, X.; Li, K.; Lv, Q.; Dick, R. P.; Hannigan, M.; Shang, L. Ariel: Automatic WiFi based room fingerprinting for indoor localization. *Proceedings of the 2012 ACM Conference on Ubiquitous Computing* **2012**, 441-450.
- 8. Xia, S.; Liu, Y.; Yuan, G.; Zhu, M.; Wang, Z. Indoor fingerprint positioning based on WiFi: an overview. *ISPRS International Journal of Geo-information* **2017**, *6*, 135.
- 9. Li, Y.; Williams, S.; Moran, B.; Kealy, A. Quantized RSS Based Wi-Fi Indoor Localization with Room Level Accuracy. *Proceedings of the IGNSS Conference* **2018**, 7-9.
- 10. Kushki, A.; Plataniotis, K. N.; Venetsanopoulos, A. N. Kernel-based positioning in wireless local area networks. *IEEE transactions on mobile computing* **2007**, *6*, 689-705.
- ⁴⁹³ 11. Mirowski, P.; Milioris, D.; Whiting, P.; Kam Ho, T. Probabilistic radio-frequency fingerprinting and localization on the run. *Bell Labs Technical Journal* **2014**, *18*, 111-133.
- Milioris, D.; Kriara, L.; Papakonstantinou, A.; Tzagkarakis, G.; Tsakalides, P.; Papadopouli, M. Empirical evaluation of signal-strength fingerprint positioning in wireless LANs. *Proceedings of the 13th ACM international conference on Modeling, analysis, and simulation of wireless and mobile systems* **2010**, 5-13.
- Yang, C.; Shao, H.R. WiFi-based indoor positioning. *IEEE Communications Magazine* 2015, 53,
 150-157.
- 14. Castro, P.; Chiu, P.; Kremenek, T.; Muntz, R. A probabilistic room location service for wireless networked environments. *International conference on ubiquitous computing* **2001**, 18-34.
- 503 15. Ghahramani, Z.; Jordan, M.I. Supervised learning from incomplete data via an EM approach.

 Advances in neural information processing systems **1994**, 120-127.
- Liu, H.; Gan, Y.; Yang, J.; Sidhom, S.; Wang, Y.; Chen, Y.; Ye, F. Push the limit of WiFi based localization for smartphones. *Proceedings of the 18th annual international conference on Mobile computing and networking* 2012, 305-316.
- Chintalapudi, K.; Padmanabha Iyer, A.; Padmanabhan, V. N. Indoor localization without the pain.
 Proceedings of the sixteenth annual international conference on Mobile computing and networking 2010, 173-184.
- 18. Retscher, G.; Tatschl, T. Indoor positioning using differential Wi-Fi lateration. *Journal of Applied Geodesy* **2017**, *11*, 249-269.

- 19. Yim, J.; Park, C.; Joo, J.; Jeong, S. Extended Kalman Filter for wireless LAN based indoor positioning. *Decision support systems* **2008**, *45*, 960-971.
- 515 20. Frank, K.; Krach, B.; Catterall, N.; Robertson, P. Development and evaluation of a combined WLAN and inertial indoor pedestrian positioning system. *4th International Symposium on Location and Context Awareness* **2009**.
- Youssef, M.; Agrawala, A. The Horus WLAN location determination system. *Proceedings of the*3rd international conference on Mobile systems, applications, and services **2005**, 205-218.
- 520 22. Mirowski, P.; Steck, H.; Whiting, P.; Palaniappan, R.; MacDonald, M.; Ho, T. K. KL-divergence 521 kernel regression for non-Gaussian fingerprint based localization. *Indoor Positioning and Indoor* 522 *Navigation (IPIN)* **2011**, 1-10.
- ⁵²³ 23. Chen, L.; Li, B.; Zhao, K.; Rizos, C.; Zheng, Z. An improved algorithm to generate a Wi-Fi fingerprint database for indoor positioning. *Sensors* **2013**, *13*, 11085-11096.
- Vaupel, T.; Seitz, J.; Kiefer, F.; Haimerl, S.; Thielecke, J. Wi-Fi positioning: System considerations and device calibration. *Indoor Positioning and Indoor Navigation (IPIN)* **2010**, 1-7.
- Mirowski, P.; Whiting, P.; Steck, H.; Palaniappan, R.; MacDonald, M.; Hartmann, D.; Ho, T. K.
 Probability kernel regression for WiFi localisation. *Journal of Location Based Services* 2012, 6, 81-100.
- 26. Roos, T.; Myllymäki, P.; Tirri, H.; Misikangas, P.; Sievänen, J. A probabilistic approach to WLAN user location estimation. *International Journal of Wireless Information Networks* **2002**, *9*, 155-164.
- Meng, W.; Xiao, W.; Ni, W.; Xie, L. Secure and Robust Wi-Fi Fingerprinting Indoor Localization.
 Indoor Positioning and Indoor Navigation (IPIN) 2011, 1-7.
- ⁵³⁴ 28. Youssef, M.A.; Agrawala, A.; Shankar, A.U. WLAN location determination via clustering and probability distributions. *Pervasive Computing and Communications (Percom)* **2003**, 143-150.
- Zhang, T.; Zhao, Q.; Shin, K.; Nakamoto, Y. Bayesian Optimization Based Peak Searching
 Algorithm for Clustering in Wireless Sensor Networks. *Journal of Sensor and Actuator Networks* 2018, 7, 2.
- 539 30. Khalajmehrabadi, A.; Gatsis, N.; Akopian, D. Modern WLAN fingerprinting indoor positioning 540 methods and deployment challenges. *IEEE Communications Surveys & Tutorials* **2017**, *19*, 541 1974-2002.
- Jiang, P.; Zhang, Y.; Fu, W.; Liu, H.; Su, X. Indoor mobile localization based on Wi-Fi fingerprint's important access point. *Journal of Distributed Sensor Networks* **2015**, *11*, 429104.
- Stella, M., Russo, M. and Begušić, D. RF Localization in Indoor Environment. *Radioengineering* 2012, 21.
- 33. Beder, C.; Klepal, M. Fingerprinting based Localisation Revisited: A Rigorous Approach for Comparing RSSI Measurements Coping with Missed Access Points and Differing Antenna Attenuations. *Indoor Positioning and Indoor Navigation (IPIN)* **2012**, 1-7.
- J., Biswas; M., Veloso. WiFi localization and navigation for autonomous indoor mobile robots.
 Robotics and Automation (ICRA) 2010, 4379-4384.
- 35. Weyn, M. Opportunistic Seamless Localization. In *Lulu. com* **2011**.

- 552 36. P., Bolliger. Redpin-adaptive, zero-configuration indoor localization through user collaboration.

 553 Proceedings of the first ACM international workshop on Mobile entity localization and tracking in

 554 GPS-less environments 2008, 55-60.
- Luo, J.; Zhan, X. Characterization of Smart Phone Received Signal Strength Indication for WLAN Indoor Positioning Accuracy Improvement. *JNW* **2014**, *9*, 739-746.
- Bilmes, J.A. A gentle tutorial of the EM algorithm and its application to parameter estimation for Gaussian mixture and hidden Markov models. *International Computer Science Institute* **1998**, *4*, 126.
- Goswami, A.; Ortiz, L.E.; Das, S.R. WiGEM: A learning-based approach for indoor localization.
 Proceedings of the Seventh Conference on emerging Networking Experiments and Technologies
 2011, 3.
- 40. Alfakih, M.; Keche, M.; Benoudnine, H. Gaussian mixture modeling for indoor positioning WIFI systems. *Control, Engineering & Information Technology (CEIT)* **2015**, 1-5.
- Tseng, C.H.; Yen, J.S. Enhanced Gaussian mixture model of RSSI purification for indoor positioning. *Journal of Systems Architecture* **2017**, 1-6.
- 42. Wu, C.; Yang, Z.; Liu, Y. Smartphones based crowdsourcing for indoor localization. *IEEE Transactions on Mobile Computing* 2015, *14*, 444-457.
- ⁵⁶⁹ 43. Zhou, B.; Li, Q.; Mao, Q.; Tu, W.; Zhang, X.; Chen, L. ALIMC: Activity landmark-based indoor mapping via crowdsourcing. *IEEE Transactions on Intelligent Transportation Systems* **2015**, *16*, 2774-2785.
- 572 44. Käjergaard, M.B. and Munk, C.V. Hyperbolic location fingerprinting: A calibration-free solution 573 for handling differences in signal strength (concise contribution). *Pervasive Computing and* 574 *Communications* **2008**,110-116.
- Haeberlen A.; Flannery E.; Ladd A.M.; Rudys A.; Wallach D.S.; Kavraki L.E. Practical robust localization over large-scale 802.11 wireless networks. *Proceedings of the 10th annual international conference on Mobile computing and networking* **2004**, 70-84.
- 46. Alexopoulos, E.C. Introduction to Multivariate Regression Analysis. *Hippokratia* **2010**, *14*, 23.
- 579 47. Kang, H. The prevention and handling of the missing data. *Korean Journal of Anesthesiology* **2013**, 64, 402-406.
- 48. Li, C.; Xu, Q.; Gong, Z.; Zheng, R. TuRF: Fast data collection for fingerprint-based indoor localization. *Indoor Positioning and Indoor Navigation (IPIN)* **2017**, 1-8.
- 49. Milioris, D.; Tzagkarakis, G.; Papakonstantinou, A.; Papadopouli, M.; Tsakalides, P.
 Low-dimensional signal-strength fingerprint-based positioning in wireless LANs. Ad hoc networks
 2014, 12, 100-114.
- 50. Hernández, N.; Ocaña, M.; Alonso, J.M.; Kim, E. Continuous space estimation: Increasing WiFi-based indoor localization resolution without increasing the site-survey effort. *Sensors* **2017**, 17, 147.
- 51. Ahmad, U.; Gavrilov, A.; Nasir, U.; Iqbal, M.; Cho, S.J.; Lee, S. In-building localization using neural networks. *Engineering of Intelligent Systems* **2006**, 1-6.
- 591 52. Meng, X.L.; Rubin, D.B. Maximum likelihood estimation via the ECM algorithm: A general framework. *Biometrika* **1993**, *80*, 267-278.

- 53. Dempster, A.P.; Laird, N.M.; Rubin, D.B. Maximum likelihood from incomplete data via the EM algorithm. *Journal of The Royal Statistical Society. Series B (methodological)* **1977**, 1-38.
- 595 54. Dong, Y.; Peng, C.Y.J. Principled missing data methods for researchers. *SpringerPlus* **2013**, 2, 222.
- 55. H. Akaike. Akaike's information criterion. *International Encyclopedia of Statistical Science* **2011**, 25.
- 599 56. Kaemarungsi, K.; Krishnamurthy, P. Properties of indoor received signal strength for WLAN location fingerprinting. *Mobile and Ubiquitous Systems: Networking and Services* **2004**, 14-23.
- 57. Correa, A.; Munoz Diaz, E.; Bousdar Ahmed, D.; Morell, A.; Lopez Vicario, J. Advanced Pedestrian Positioning System to Smartphones and Smartwatches. *Sensors* **2016**, *16*, 1903.
- 58. Pfaff, P.; Plagemann, C.; Burgard, W. Gaussian mixture models for probabilistic localization. *Robotics and Automation* **2008**, 467-472.
- Kaji, K.; Kawaguchi, N. Design and implementation of WiFi indoor localization based on Gaussian
 mixture model and particle filter. *Indoor Positioning and Indoor Navigation (IPIN)* 2012, 1-9.
- 60. Rabiner, L. R. A tutorial on hidden Markov models and selected applications in speech recognition.

 Proceedings of the IEEE 1989, 77, 257-286.
- 609 61. Retscher, G.; Hofer, H. Wi-Fi Location Fingerprinting Using an Intelligent Checkpoint Sequence.

 Journal of Applied Geodesy **2017**, 11, 197-205.
- 611 62. Zhou, R. Wireless indoor tracking system (WITS). *Aktuelle Trends in der Softwareforschung,*612 Tagungsband zum doIT Software-Forschungtag **2006**, 163-177.
- 613 63. Elnahrawy, E.; Li, X.; Martin, R.P. The limits of localization using signal strength: A comparative study. *Sensor and Ad Hoc Communications and Networks* **2004**, 406-414.
- 64. Eirola, E.; Lendasse, A.; Vandewalle, V; Biernacki, C. Mixture of gaussians for distance estimation with missing data. *Neurocomputing* **2014**, *131*, 32-42.
- 65. Stefánsson, A.; Končar, N.; Jones, A.J. A note on the gamma test. *Neural Computing & Applications* **1997**, *5*(*3*), 131-133.