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# A Software Reliability Model with a Weibull Fault Detection Rate Function Subject to Operating Environments

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**ABSTRACT:** The main focus when developing software is to improve the reliability and stability of a software system. When software systems are introduced, these systems are used in field environments that are the same as or close to those used in the development-testing environment; however, they may also be used in many different locations that may differ from the environment in which they were developed and tested. In this paper, we propose a new software reliability model that takes into account the uncertainty of operating environments. The explicit mean value function solution for the proposed model is presented. Examples are presented to illustrate the goodness-of-fit of the proposed model and several existing non-homogeneous Poisson process (NHPP) models and confidence intervals of all models based on two sets of failure data collected from software applications. The results show that the proposed model fits the data more closely than other existing NHPP models to a significant extent.

**Keywords:** non-homogeneous poisson process, software reliability; weibull function; mean squared error

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## 1. Introduction

Software systems have become an essential part of our lives. These systems are very important because they are able to ensure the provision of high-quality services to customers because of their reliability and stability. However, software development is a difficult and complex process. Therefore, the main focus of software companies is on improving the reliability and stability of a software system. This has prompted research in software reliability engineering and many software reliability growth models (SRGM) have been proposed over the past decades. Many existing non-homogeneous Poisson process (NHPP) software reliability models have been developed through the fault intensity rate function and the mean value functions  $m(t)$  within a controlled testing environment to estimate reliability metrics such as the number of residual faults, failure rate, and reliability of the software. Generally, the reliability increases more quickly and later the improvement slows down. Software reliability models are used to estimate and predict the reliability, number of remaining faults, failure intensity, total software development cost, etc., of software. Various software reliability models and application studies have been developed to date. Discovering the confidence intervals of software reliability is done in the field of software reliability because it can enhance the decision of software releases and control the related expenditures for software testing [1]. First, Yamada and Osaki [2] considered that the maximum likelihood estimates concerning the confidence interval of the mean value function can be estimated. Yin and Trivedi [3] present the confidence bounds for the model parameters via the Bayesian approach. Huang [4] also present a graph to illustrate the confidence interval of the mean value function.

Chatterjee and Singh [5] proposed a software reliability model based on NHPP that incorporates a logistic-exponential testing coverage function with imperfect debugging. In addition, Chatterjee and Shukla [6] developed a software reliability model that considers different types of faults incorporating both imperfect debugging and a change point. Generally, existing models are applied to software testing data and then used to make predictions on the software failures and reliability in the field. Here, the important point is that the test environment and operational environment are different from each other. Once software systems are introduced, the software systems used in the field environments are the same as or close to those used in the development-testing environment; however, the systems may be used in many different locations. Several researchers started applying the factor of operational environments. A few researchers, Yang and Xie, Huang et al., and Zhang et al. [7-9], proposed a method of predicting the fault detection rate to reflect changes in operating environments, and used methodology that modifies the software reliability model in the operating environments by introducing a calibration factor. Teng and Pham [10] discussed a generalized model that captures the uncertainty of the environment and its effects upon the software failure rate. Pham [11-12] and Chang et al. [13] developed a software reliability model incorporating the uncertainty of the system fault-detection rate per unit of time subject to the operating environment. Honda et al. [14] proposed a generalized software reliability model (GSRM) based on a stochastic process and simulated developments that include uncertainties and dynamics. Pham [15] recently presented a new generalized software reliability model subject to the uncertainty of operating environments. And also, Song et al. [16] presented a new model with consideration of a three-parameter fault-detection rate in the software development process, and relate it to the error detection rate function with consideration of the uncertainty of operating environments.

In this paper, we discuss a new model with consideration for the Weibull function in the software development process and relate it to the error detection rate function with consideration of the uncertainty of operating environments. We examine the goodness-of-fit of the fault-detection rate software reliability model and other existing NHPP models based on several sets of software testing data. The explicit solution of the mean value function for the new model is derived in Section 2. Criteria for model comparisons and confidence interval for selection of the best model are discussed in Section 3. Model analysis and results are discussed in Section 4. Section 5 presents the conclusions and remarks.

## 2. A New software reliability model

### 2.1 Non-homogeneous Poisson process model

The software fault detection process has been widely formulated by using a counting process. A counting process  $\{N(t), t \geq 0\}$ , is said to be a non-homogeneous Poisson process with intensity function  $\lambda(t)$ , if  $N(t)$  follows a Poisson distribution with the mean value function  $m(t)$ , i.e.,

$$\Pr\{N(t) = n\} = \frac{\{m(t)\}^n}{n!} \exp\{-m(t)\}, n = 0, 1, 2, 3 \dots$$

The mean value function  $m(t)$ , which is the expected number of faults detected at time  $t$  can be expressed as

$$m(t) = \int_0^t \lambda(s) ds$$

where  $\lambda(t)$  represents the failure intensity.

A general framework for NHPP-based SRGM has been proposed by Pham et al. [17]. They have modeled  $m(t)$  using the differential equation

$$\frac{dm(t)}{dt} = b(t)[a(t) - m(t)]. \quad (1)$$

Solving Equation (1) makes it possible to obtain different values of  $m(t)$  using different values for  $a(t)$  and  $b(t)$ , which reflects various assumptions of the software testing process.

## 2.2. Weibull fault detection rate function model

A generalized NHPP model incorporating the uncertainty of operating environments can be formulated as follows [12]:

$$\frac{dm(t)}{dt} = \eta[b(t)][N - m(t)], \quad (2)$$

where  $\eta$  is a random variable that represents the uncertainty of the system fault detection rate in the operating environments with a probability density function  $g$ ,  $N$  is the expected number of faults that exists in the software before testing,  $b(t)$  is the fault detection rate function, which also represents the average failure rate of a fault, and  $m(t)$  is the expected number of errors detected by time  $t$  or the mean value function. We propose an NHPP software reliability model including the uncertainty of the operating environment using Equation (2) and the following assumptions:

- (a) The occurrence of software failures follows an NHPP.
- (b) Software can fail during execution, caused by faults in the software.
- (c) The software-failure detection rate at any time is proportional to the number of remaining faults in the software at that time.
- (d) When a software failure occurs, a debugging effort removes the faults immediately.
- (e) For each debugging effort, regardless of whether the faults are successfully removed, some new faults may be introduced into the software system.
- (f) The environment affects the unit failure detection rate,  $b(t)$ , by multiplying by a factor  $\eta$ .

The solution for the mean value function  $m(t)$ , where the initial condition  $m(0) = 0$ , is given by [12]:

$$m(t) = \int_{\eta} N \left( 1 - e^{-\eta \int_0^t b(x) dx} \right) dg(\eta). \quad (3)$$

Pham [15] recently developed a generalized software reliability model incorporating the uncertainty of fault-detection rate per unit of time in the operating environments where the random variable  $\eta$  has a generalized probability density function  $g$  with two parameters  $\alpha \geq 0$  and  $\beta \geq 0$  and the mean value function from Equation (3) is given by:

$$m(t) = N \left( 1 - \frac{\beta}{\beta + \int_0^t b(s) ds} \right)^{\alpha}, \quad (4)$$

where  $b(t)$  is the fault detection rate per fault per unit of time.

In this paper, we consider a Weibull fault detection rate function  $b(t)$  to be as follows:

$$b(t) = a^b t^{b-1}, \quad a, b > 0, \quad (5)$$

where,  $a$  and  $b$  are known as the scale and shape parameters, respectively. A Weibull fault detection rate function  $b(t)$  is decreasing for  $b < 1$ , increasing for  $b > 1$ , and constant when  $b = 1$ . We obtain a new NHPP software reliability model subject to the uncertainty of the environments,  $m(t)$ , that can be used to determine the expected number of software failures detected by time  $t$  by substituting the function  $b(t)$  above into Equation (4):

$$m(t) = N \left( 1 - \frac{\beta}{\beta + (at)^b} \right)^{\alpha}, \quad (6)$$

## 3. Model comparisons

### 3.1 Criteria for model comparisons

Once the analytical expression for the mean value function  $m(t)$  is derived, the model parameters to be estimated in the mean value function can then be obtained with the help of a developed Matlab program based on the least-squares estimate (LSE) method. Five common criteria [18-19], namely the mean squared error (MSE), the sum absolute error (SAE), the predictive ratio risk (PRR), the predictive power (PP), and Akaike's information criterion (AIC) will be used as criteria for the model estimation of the goodness-of-fit and to compare the proposed model and other existing models as listed in Table 1.

The mean squared error is given by

$$MSE = \frac{\sum_{i=0}^n (m(t_i) - y_i)^2}{n - m}.$$

The sum absolute error is given by

$$SAE = \sum_{i=0}^n |m(t_i) - y_i|.$$

The predictive ratio risk and the predictive power are given as follows:

$$PRR = \sum_{i=0}^n \left( \frac{\hat{m}(t_i) - y_i}{\hat{m}(t_i)} \right)^2, \quad PP = \sum_{i=0}^n \left( \frac{\hat{m}(t_i) - y_i}{y_i} \right)^2.$$

To compare the all model's ability in terms of maximizing the likelihood function (MLF) while considering the degrees of freedom, Akaike's information criterion (AIC) is applied.

$$AIC = -2 \log|MLF| + 2m.$$

where  $y_i$  is the total number of failures observed at time  $t_i$ ;  $m$  is the number of unknown parameters in the model; and  $m(t_i)$  is the estimated cumulative number of failures at  $t_i$  for  $i = 1, 2, \dots, n$ .

The mean squared error measures the distance of a model estimate from the actual data with the consideration of the number of observations,  $n$ , and the number of unknown parameters in the model,  $m$ . The sum absolute error is similar to the sum squared error, but the way of measuring the deviation is by the use of absolute values, and sums the absolute value of the deviation between the actual data and the estimated curve. The predictive ratio risk measures the distance of model estimates from the actual data against the model estimate. The predictive power measures the distance of model estimates from the actual data against the actual data. For all five these criteria - MSE, SAE, PRR, PP and AIC - the smaller the value, the closer the model fits relative to other models run on the same data set.

### 3.2 Estimation of the confidence intervals

In this section, we use Equation (7) to obtain the confidence intervals [18] of the software reliability models in Table 1. The confidence interval is given by

$$\hat{m}(t) \pm Z_{\alpha/2} \sqrt{\hat{m}(t)}, \quad (7)$$

where,  $Z_{\alpha/2}$  is  $100(1 - \alpha)$  percentile of the standard normal distribution. Table 1 summarizes the proposed model and several existing well-known NHPP models with different mean value functions. Note that models 9 and 10 in Table 1 did consider environmental uncertainty.

## 4. Numerical examples

The software in this numerical example runs on an element within a wireless network switching center. Its main function includes routing voice channels and signaling messages to relevant radio resources and processing entities. Dataset #1, field failure data for Release 1 listed in Table 2, was reported by Jeske and Zhang [26]. Release 1 were Year 2000 compatibility modifications, operating system upgrade, and some new features pertaining to the signaling message processing. Release 1 had a life cycle of 13 months in the field. The cumulative field exposure time of the software was 167,900 system-days and a total of 115 failures were observed in the field. Table 4 shows the field failure data for Release 1 for each of the 13 months. Dataset #2, test data for Release 2 listed in Table 3, was also reported by Jeske and Zhang [26]. The test data is the

set of failures that were observed during a combination of feature testing and load testing. The test interval that was used in this analysis was a 38 week period between. At times, as many as 11 different BCF frames were being used in parallel to test the software. Thus, to obtain an overall number of days spent testing the software we aggregated the number of days spent testing the software on each frame. The table 3 shows how many days in each week of the test interval that each frame was used for Release 2 testing. The 38 weeks of Release 2 testing accumulates to 1001 days of exposure time.

**Table 1.** Software reliability models

No.	Model	$m(t)$
1	G-O model [20]	$m(t) = a(1 - e^{-bt})$
2	Delayed S-shaped SRGM [21]	$m(t) = a(1 - (1 + bt)e^{-bt})$
3	Inflection S-shaped SRGM [22]	$m(t) = \frac{a(1 - e^{-bt})}{1 + \beta e^{-bt}}$
4	Yamada Imperfect Debugging Model [23]	$m(t) = a[1 - e^{-bt}] \left[ 1 - \frac{\alpha}{b} \right] + \alpha t$
5	PNZ Model [17]	$m(t) = \frac{a[1 - e^{-bt}] \left[ 1 - \frac{\alpha}{b} \right] + \alpha t}{1 + \beta e^{-bt}}$
6	Pham-Zhang Model [24]	$m(t) = \frac{\left( (c + a)[1 - e^{-bt}] - \left[ \frac{ab}{b - \alpha} (e^{-\alpha t} - e^{-bt}) \right] \right)}{1 + \beta e^{-bt}}$
7	Dependent-Parameter Model1 [25]	$m(t) = \alpha(1 + \gamma t)(\gamma t + e^{-\gamma t} - 1)$
8	Dependent-Parameter Model2 [25]	$m(t) = m_0 \left( \frac{\gamma t + 1}{\gamma t_0 + 1} \right) e^{-\gamma(t-t_0)} + \alpha(\gamma t + 1)(\gamma t - 1 + (1 - \gamma t_0)e^{-\gamma(t-t_0)})$
9	Testing Coverage Model [13]	$m(t) = N \left[ 1 - \left( \frac{\beta}{\beta + (at)^b} \right)^\alpha \right]$
10	Three-parameter Model [16]	$m(t) = N \left[ 1 - \left( \frac{\beta}{\beta - \frac{a}{b} \ln \left( \frac{(1+c)e^{-bt}}{1+ce^{-bt}} \right)} \right) \right]$
11	Proposed New Model	$m(t) = N \left( 1 - \frac{\beta}{\beta + (at)^b} \right)^\alpha$

**Table 2.** Field failure data for Release 1 – Dataset #1

Month Index	System days (Days)	System days (Cumulative)	Failures	Cumulative Failures
1	961	961	7	7
2	4170	5131	3	10
3	8789	13920	14	24

Month Index	System days (Days)	System days (Cumulative)	Failures	Cumulative Failures
4	11858	25778	8	32
5	13110	38888	11	43
6	14198	53086	8	51
7	14265	67351	7	58
8	15175	82526	19	77
9	15376	97902	17	94
10	15704	113606	6	100
11	18182	131788	11	111
12	17760	149548	4	115
13	18352	167900	0	115

Table 3. Test data for Release 2 – Dataset #2

Week Index	System days (Cumulative)	Cumulative Failures	Week Index	System days (Cumulative)	Cumulative Failures
1	5	5	19	383	105
2	9	6	20	418	110
3	13	13	21	467	117
4	18	13	22	519	123
5	28	22	23	570	128
6	33	24	24	619	130
7	43	29	25	657	136
8	63	34	26	699	141
9	88	40	27	733	148
10	123	46	28	775	156
11	153	53	29	798	156
12	178	63	30	845	164
13	203	70	31	892	166
14	238	71	32	934	169
15	263	74	33	955	170
16	288	78	34	977	176
17	318	90	35	999	180
18	348	98	36	1001	181

Table 4 and 5 summarize the results of the estimated parameters of all 11 models in Table 1 using the least-squares estimation (LSE) technique and the values of the five common criteria (MSE, SAE, PRR, PP and AIC). We obtained the five common criteria when  $t = 1, 2, \dots, 13$  from Dataset #1 (Table 2), with exposure time (Cum. System days) from Dataset #2 (Table 3), As can be seen from Table 4, the MSE, SAE, PRR, PP and AIC value for the proposed new model are the lowest values compared to all models. As can be seen from Table 5, the MSE, SAE and PRR value for the proposed new model are the lowest values, and the PP and AIC value for the proposed new model are the second lowest values compared to all models. Table 7 and 8 show confidence intervals of all 11 models from Dataset #1 and #2, respectively ( $\alpha = 0.05$ ).

Figure 1 show the graph of the mean value functions for all 11 models for Datasets #1 and #2, respectively. Figure 2 and 3 show that the relative error value of the software reliability model can quickly approach zero in comparison with the other models confirming its ability to provide more accurate prediction. Figure 4 and 5 show the graph of the mean value function and confidence interval each of all 11 models for Datasets #1 and #2, respectively.

Table 4. Model parameter estimation and comparison criteria from Dataset #1

Model	LSE's	MSE	SAE	PRR	PP	AIC
G-O Model	$\hat{a}=2354138, \hat{b}=0.000004$	43.6400	72.2548	0.3879	1.0239	98.7606

Delayed S-shaped SRGM	$\hat{a}=168.009, \hat{b}=0.195$	20.7414	43.2510	2.3107	0.4295	92.2587
Inflection S-shaped SRGM	$\hat{a}=134.540, \hat{b}=0.336, \hat{\beta}=8.939$	15.3196	37.2090	0.2120	0.1587	85.3000
Yamada Imperfect Debugging Model	$\hat{a}=1.130, \hat{b}=1.110, \hat{\alpha}=9.129$	33.3890	51.0913	0.3027	0.2495	100.7378
PNZ Model	$\hat{a}=134.549, \hat{b}=0.3359, \hat{\alpha}=0.0, \hat{\beta}=8.940$	17.0223	37.2442	0.2124	0.1588	87.3098
Pham-Zhang Model	$\hat{a}=51.455, \hat{b}=0.336, \hat{\alpha}=289998.1, \hat{\beta}=8.939, \hat{c}=83.085$	19.1495	37.2091	0.2120	0.1587	89.3019
Dependent-Parameter Model1	$\hat{\alpha}=0.0088, \hat{\gamma}=9.996$	370.8651	207.3750	60.5062	2.6446	164.5728
Dependent-Parameter Model2	$\hat{\alpha}=672.637, \hat{\gamma}=0.04, t_0=0.027, m_0=23.541$	215.7784	133.2294	1.1037	8.6260	168.846
Testing Coverage Model	$\hat{a}=0.242, \hat{b}=1.701, \hat{\alpha}=17.967, \hat{\beta}=73.604, \hat{N}=149.410$	25.9244	41.8087	1.4473	0.3601	95.5655
Three-parameter Model	$\hat{a}=2.980, \hat{b}=0.336, \hat{\beta}=0.080, \hat{c}=1105.772, \hat{N}=135.142$	19.1517	37.2107	0.2119	0.1588	89.3053
Proposed New Model	$\hat{a}=0.095, \hat{b}=15.606, \hat{\alpha}=0.085, \hat{\beta}=1.855, \hat{N}=116.551$	11.2281	26.5568	0.2042	0.1558	79.3459

Table 5. Model parameter estimation and comparison criteria from Dataset #2

Model	LSE's	MSE	SAE	PRR	PP	AIC
G-O Model	$\hat{a}=291.768, \hat{b}=0.001$	95.3796	299.7160	24.7924	3.4879	198.5419
Delayed S-shaped SRGM	$\hat{a}=168.568, \hat{b}=0.0057$	178.4899	387.7724	7368.5885	7.4923	317.8791
Inflection S-shaped SRGM	$\hat{a}=200.110, \hat{b}=0.002, \hat{\beta}=0.059$	43.2888	182.9709	10.4336	2.1725	202.0752
Yamada Imperfect Debugging Model	$\hat{a}=81.999, \hat{b}=0.0063, \hat{\alpha}=0.0014$	18.9651	119.1208	3.1804	1.0871	187.7564
PNZ Model	$\hat{a}=67.132, \hat{b}=0.009, \hat{\alpha}=0.0019, \hat{\beta}=0.0001$	18.2406	119.7722	1.5566	0.6869	188.9438
Pham-Zhang Model	$\hat{a}=200.057, \hat{b}=0.002, \hat{\alpha}=9999.433, \hat{\beta}=0.058, \hat{c}=0.001$	46.0819	183.0449	10.4090	2.1698	206.0887
Dependent-Parameter Model1	$\hat{\alpha}=0.0003, \hat{\gamma}=0.866$	2075.6677	1411.8412	1165906.4021	17.1338	554.6335
Dependent-Parameter Model2	$\hat{\alpha}=9.035, \hat{\gamma}=0.005, t_0=48.975, m_0=49.004$	1379.2331	1134.6843	13.0318	156.8519	572.8343
Testing Coverage Model	$\hat{a}=0.002, \hat{b}=0.646, \hat{\alpha}=0.137, \hat{\beta}=8.920, \hat{N}=7973.501$	16.5529	116.0937	0.3033	0.4499	187.4100
Three-parameter Model	$\hat{a}=0.011, \hat{b}=0.707, \hat{\beta}=8.029, \hat{c}=0.000001, \hat{N}=300.684$	34.5762	154.1593	7.7768	1.8500	199.3282
Proposed New Model	$\hat{a}=0.004, \hat{b}=1.471, \hat{\alpha}=0.430, \hat{\beta}=78.738, \hat{N}=504.403$	9.8789	90.3633	0.2944	0.5159	187.4204



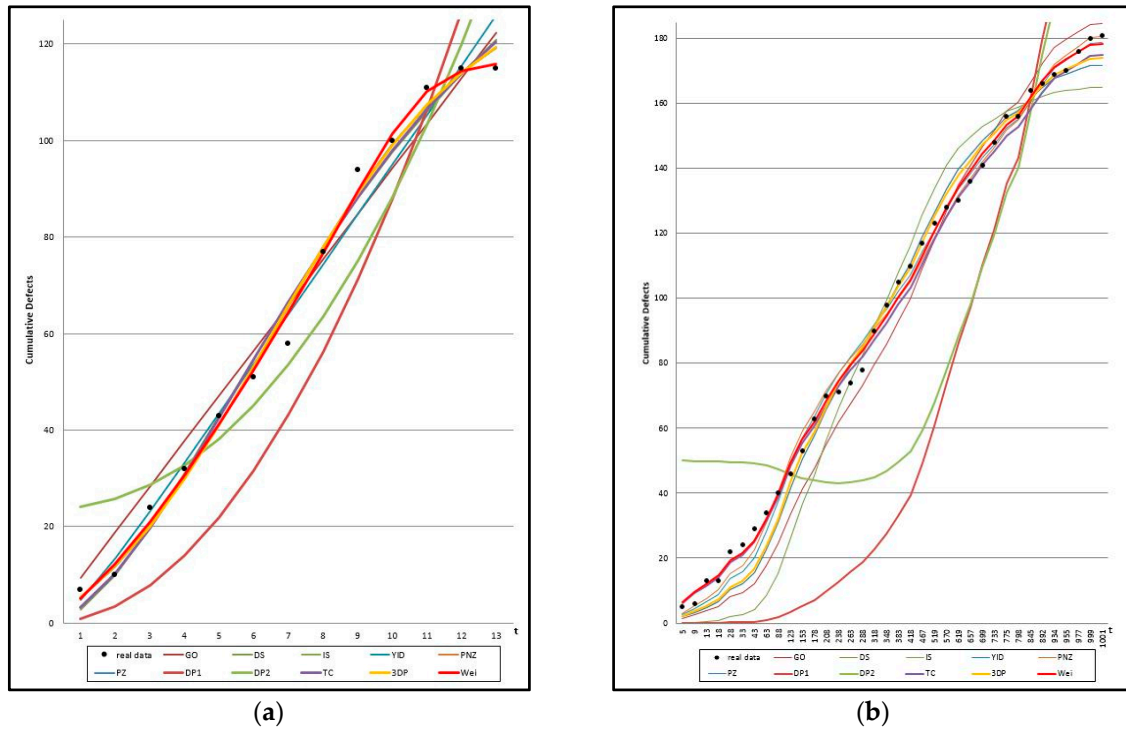


Figure 1. Mean value function of all 11 models; (a) Dataset #1; (b) Dataset #2.

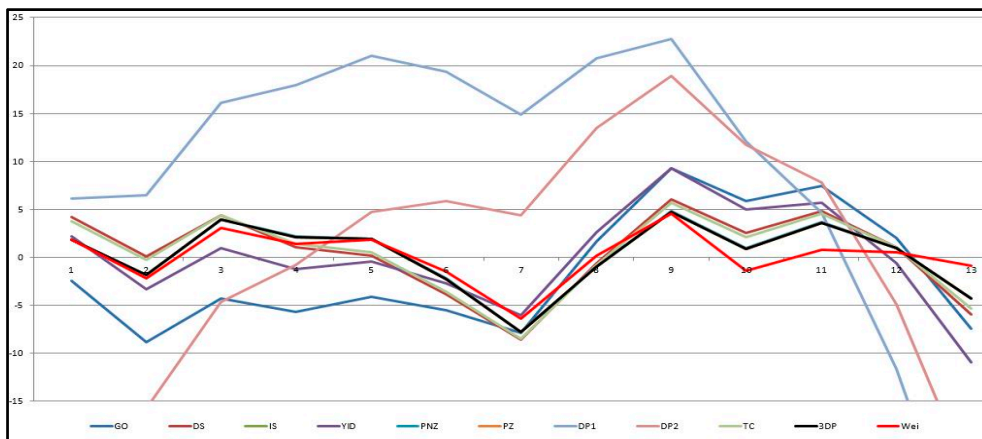


Figure 2. Relative error value of 11 models in Table 1 for Dataset #1.

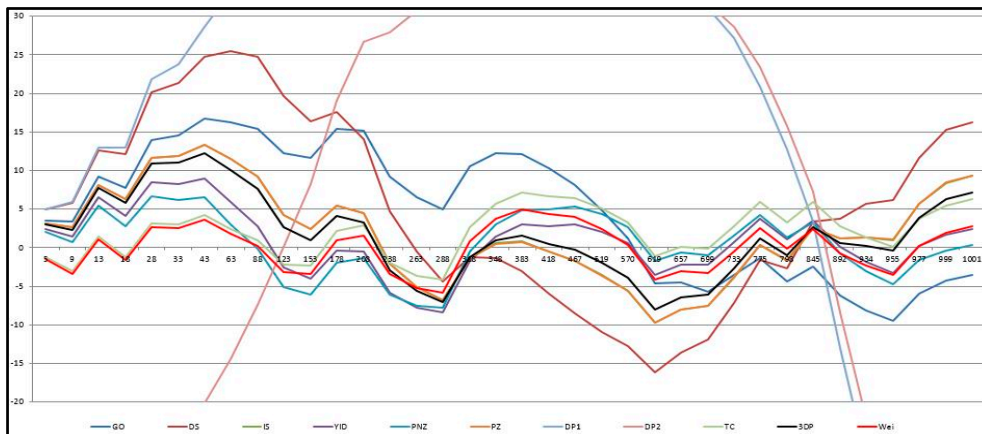
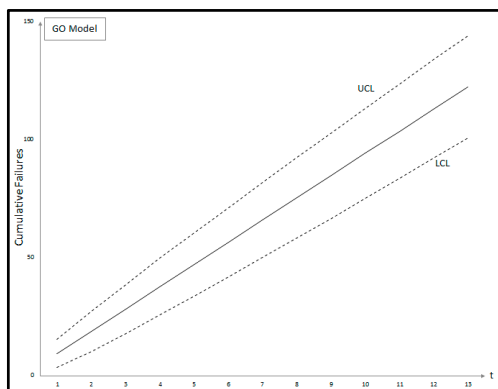


Figure 3. Relative error value of 11 models in Table 1 for Dataset #2.

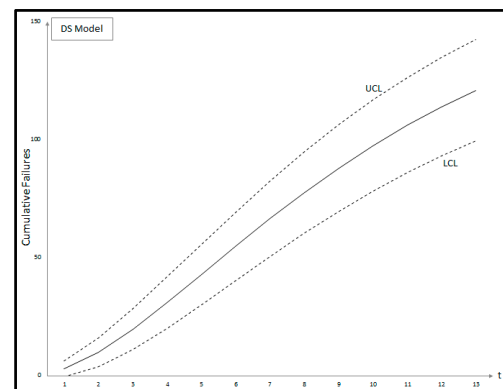
Table 6. Confidence interval of all 11 models from Dataset #1 ( $\alpha = 0.05$ )



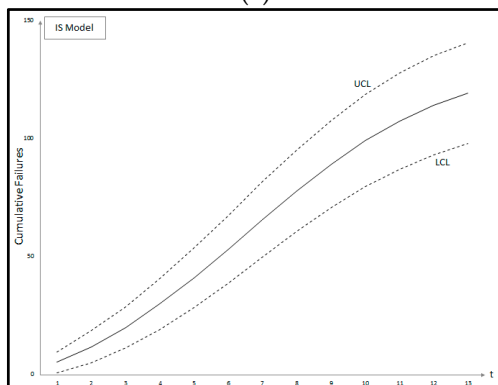
Time index Model		1	2	3	4	5	6	7	8	9	10	11	12	13
	GO	LCL	3.40	10.33	17.83	25.64	33.63	41.77	50.00	58.32	66.70	75.14	83.63	92.16
	UCL	15.43	27.34	38.67	49.69	60.53	71.23	81.83	92.34	102.79	113.18	123.53	133.83	144.10
DS	LCL	-0.48	3.73	10.97	20.02	30.02	40.34	50.55	60.36	69.59	78.14	85.96	93.03	99.38
	UCL	6.09	16.06	28.34	41.82	55.68	69.37	82.52	94.90	106.36	116.85	126.35	134.88	142.49
IS	LCL	0.73	5.09	11.27	19.14	28.50	38.94	49.88	60.67	70.72	79.59	87.07	93.13	97.91
	UCL	9.66	18.57	28.82	40.56	53.62	67.54	81.67	95.29	107.74	118.61	127.68	135.00	140.72
YID	LCL	0.53	6.19	13.66	21.90	30.53	39.37	48.36	57.46	66.64	75.89	85.20	94.55	103.95
	UCL	9.16	20.52	32.49	44.49	56.37	68.11	79.74	91.27	102.71	114.10	125.42	136.70	147.94
PNZ	LCL	0.73	5.09	11.26	19.13	28.49	38.92	49.86	60.65	70.69	79.57	87.05	93.12	97.90
	UCL	9.66	18.56	28.81	40.55	53.60	67.51	81.64	95.26	107.72	118.58	127.66	134.98	140.71
PZ	LCL	0.73	5.09	11.27	19.14	28.50	38.94	49.88	60.67	70.72	79.59	87.07	93.13	97.91
	UCL	9.66	18.57	28.82	40.56	53.62	67.54	81.67	95.29	107.74	118.61	127.68	135.00	140.72
DP1	LCL	-0.96	-0.16	2.39	6.71	12.79	20.62	30.21	41.56	54.67	69.54	86.17	104.56	124.70
	UCL	2.70	7.18	13.42	21.41	31.16	42.67	55.94	70.97	87.75	106.30	126.60	148.66	172.48
DP2	LCL	14.46	15.79	18.12	21.53	26.13	32.00	39.24	47.92	58.11	69.89	83.32	98.44	115.31
	UCL	33.69	35.68	39.08	43.96	50.37	58.34	67.93	79.16	92.08	106.73	123.14	141.36	161.41
TC	LCL	-0.29	3.96	10.96	19.79	29.70	40.09	50.49	60.53	69.93	78.53	86.23	93.01	98.87
	UCL	6.75	16.49	28.34	41.49	55.25	69.05	82.45	95.10	106.77	117.32	126.68	134.85	141.87
3PDF	LCL	0.73	5.09	11.27	19.15	28.50	38.94	49.88	60.67	70.72	79.59	87.07	93.13	97.91
	UCL	9.67	18.57	28.82	40.56	53.62	67.54	81.67	95.29	107.74	118.61	127.68	135.00	140.73
NEW	LCL	0.69	5.37	11.96	19.79	28.61	38.27	48.64	59.62	70.95	81.69	89.62	93.47	94.79
	UCL	9.57	19.07	29.88	41.49	53.77	66.66	80.09	93.97	108.03	121.17	130.77	135.41	136.99



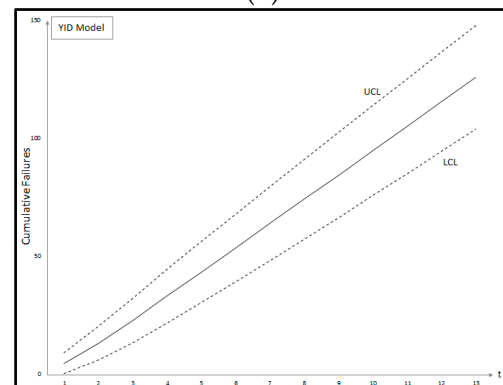
(a)



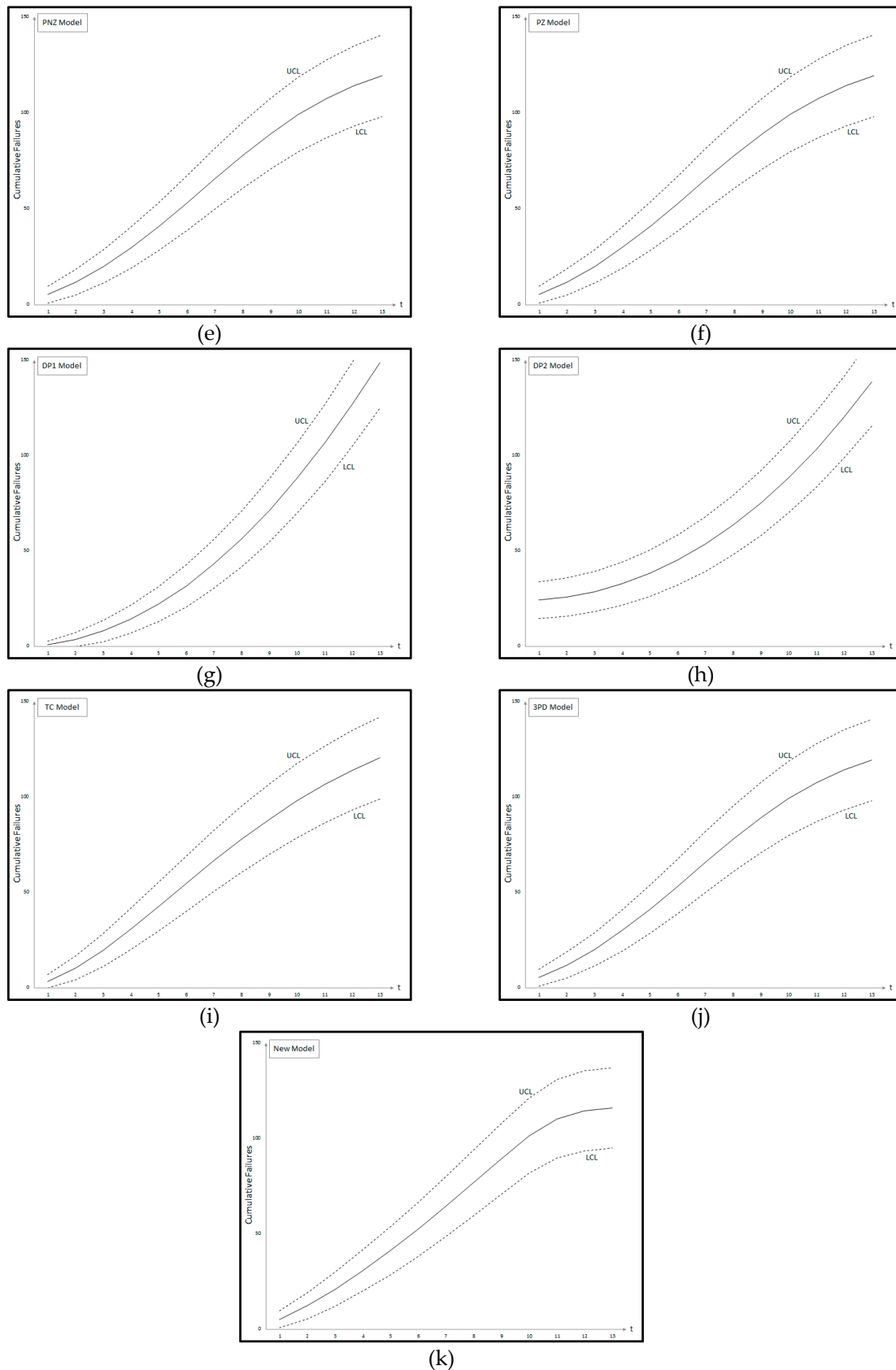
(b)



(c)



(d)

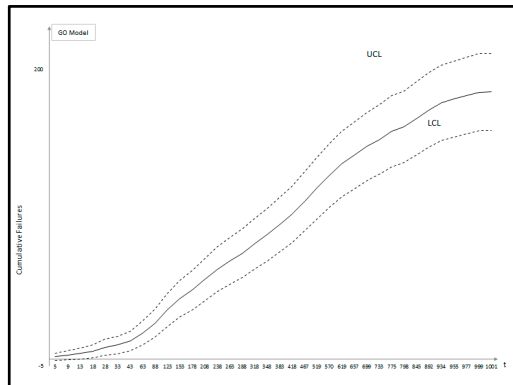


**Figure 4.** Confidence intervals of all 11 models Dataset #1: (a) GO model; (b) Delayed S-shaped SRGM; (c) Inflection S-shaped SRGM; (d) Yamada Imperfect Debugging Model; (e) PNZ Model; (f) Pham-Zhang Model; (g) Dependent-Parameter Model1; (h) Dependent-Parameter Model2; (i) Testing Coverage Model; (j) Three-parameter Model (k) Proposed New Model.

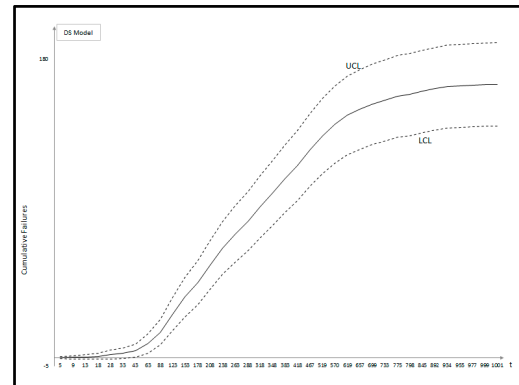
**Table 7.** Confidence interval of all 11 models from Dataset #2 ( $\alpha = 0.05$ )



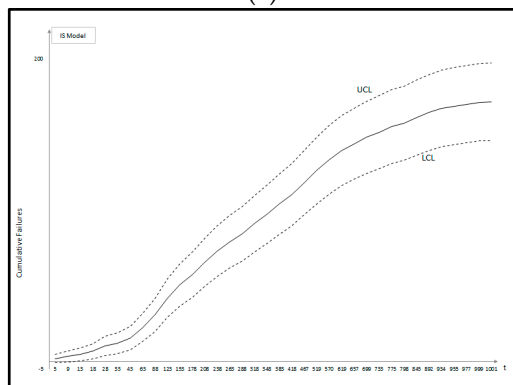
GO	LCL	117.28	122.99	127.45	132.76	135.58	141.15	146.47	151.03	153.23	155.50	157.72	157.91
	UCL	163.75	170.48	175.71	181.93	185.23	191.72	197.91	203.19	205.75	208.37	210.93	211.17
DS	LCL	125.68	128.70	130.77	132.95	133.98	135.79	137.25	138.32	138.78	139.22	139.61	139.65
	UCL	173.64	177.17	179.61	182.15	183.36	185.47	187.18	188.42	188.96	189.47	189.93	189.97
IS	LCL	120.52	124.62	127.70	131.23	133.05	136.53	139.70	142.30	143.52	144.75	145.93	146.03
	UCL	167.57	172.38	176.00	180.14	182.27	186.33	190.03	193.06	194.48	195.91	197.28	197.40
YID	LCL	115.14	119.78	123.51	128.08	130.57	135.64	140.70	145.20	147.46	149.81	152.17	152.38
	UCL	161.22	166.70	171.08	176.45	179.37	185.30	191.19	196.44	199.05	201.79	204.52	204.77
PNZ	LCL	113.71	118.66	122.67	127.62	130.33	135.87	141.41	146.37	148.85	151.45	154.05	154.29
	UCL	159.53	165.38	170.10	175.91	179.08	185.56	192.02	197.79	200.67	203.69	206.70	206.97
PZ	LCL	120.52	124.61	127.69	131.22	133.04	136.51	139.68	142.28	143.50	144.73	145.91	146.01
	UCL	167.57	172.38	175.99	180.13	182.26	186.32	190.01	193.04	194.46	195.89	197.25	197.37
DP1	LCL	77.80	89.38	99.33	112.35	119.81	135.80	152.79	168.81	177.12	186.03	195.17	196.01
	UCL	116.43	130.48	142.43	157.92	166.73	185.49	205.24	223.73	233.27	243.48	253.91	254.86
DP2	LCL	78.53	88.88	97.96	110.03	117.04	132.22	148.53	164.05	172.14	180.85	189.79	190.62
	UCL	117.33	129.87	140.79	155.17	163.46	181.29	200.29	218.24	227.55	237.54	247.77	248.72
TC	LCL	113.09	117.82	121.56	126.08	128.52	133.40	138.17	142.33	144.39	146.52	148.63	148.82
	UCL	158.79	164.38	168.79	174.11	176.96	182.68	188.24	193.10	195.49	197.96	200.41	200.63
3PDF	LCL	119.05	123.32	126.60	130.46	132.48	136.43	140.16	143.32	144.84	146.39	147.91	148.04
	UCL	165.83	170.86	174.72	179.24	181.60	186.22	190.57	194.25	196.01	197.82	199.58	199.73
NEW	LCL	115.97	120.80	124.61	129.21	131.68	136.62	141.43	145.62	147.68	149.81	151.92	152.11
	UCL	162.20	167.89	172.38	177.78	180.67	186.44	192.05	196.92	199.31	201.79	204.23	204.45



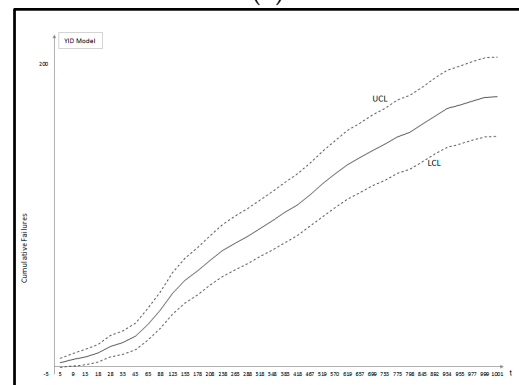
(a)



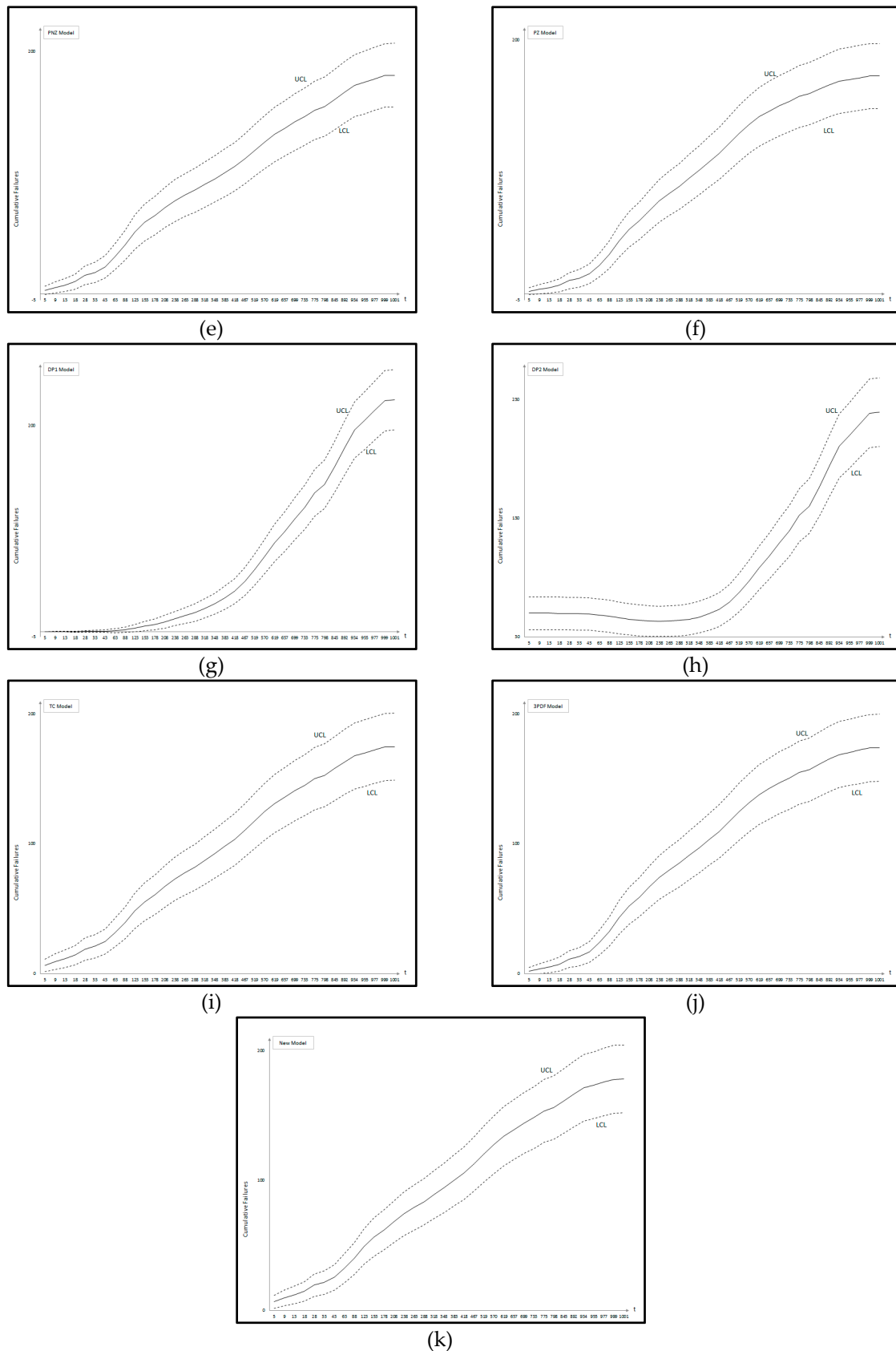
(b)



(c)



(d)



**Figure 5.** Confidence intervals of all 11 models Dataset #2: (a) GO model; (b) Delayed S-shaped SRGM; (c) Inflection S-shaped SRGM; (d) Yamada Imperfect Debugging Model; (e) PNZ Model; (f) Pham-Zhang Model; (g) Dependent-Parameter Model1; (h) Dependent-Parameter Model2; (i) Testing Coverage Model; (j) Three-parameter Model (k) Proposed New Model.

## 5. Conclusions

Generally, existing models are applied to software testing data and then used to make predictions on the software failures and reliability in the field. Here, the important point is that the test environment and operational environment are different from each other. In this paper, we discussed a new software reliability model based on a Weibull fault detection rate function subject to the uncertainty of operating environments. Table 4 and 5 summarized the results of the estimated parameters of all 11 models in Table 1 using the LSE technique and the five common criteria (MSE, SAE, PRR, PP and AIC) value for two data sets. As can be seen from Table 4, the MSE, SAE, PRR, PP and AIC value for the proposed new model are the lowest values compared to all models. As can be seen from Table 5, the MSE, SAE and PRR value for the proposed new model are the lowest values compared to all models. Finally, we show confidence intervals of all 11 models from Dataset #1 and #2, respectively, in Table 7 and 8. By estimating the confidence interval, we will help to find the optimal software reliability model at different confidence levels.

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