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Article

Clinical Implications of Inappropriate Therapy in Community-Onset Urinary Tract Infections and Development of a Bayesian Hierarchical Weighted Incidence Syndromic Combination Antibigram

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Abstract: Background/Objectives: The rise of multidrug-resistant pathogens complicates the management of urinary tract infections (UTIs), particularly in empirical antibiotic therapy. This study aimed to develop Weighted Incidence Syndromic Combination Antibigram (WISCA) models to optimize antibiotic therapy for adult patients with community-onset UTIs presenting to emergency care settings at our institution. **Methods:** A retrospective study was conducted using a hierarchical Bayesian model to develop WISCA models for UTI management. Data were extracted from microbiology laboratory records and medical databases, focusing on adult patients stratified by age, prior antibiotic exposure, and clinical characteristics. Clinical outcomes, such as extended hospital stays and in-hospital mortality, were evaluated before WISCA model development. Eleven monotherapies and 18 combination therapies were tested against 15 pathogens, with posterior coverage probabilities and 95% highest density intervals (HDIs) used to assess antibiotic coverage. **Results:** Inappropriate antibiotic treatment (both empirical and directed) was initially associated with worse clinical outcomes. After adjusting for covariates, inappropriate directed treatment during the UTI episode was significantly associated with in-hospital mortality, highlighting the need for tailored therapeutic strategies. WISCA models provided robust estimates of antibiotic coverage probabilities, revealing subgroup differences. Combination regimens demonstrated superior coverage, particularly in patients with multidrug-resistant pathogens, older adults, and those with conditions linked to MDR organisms. **Conclusions:** WISCA models show promise in optimizing antibiotic selection for UTIs requiring hospitalization by integrating local resistance patterns and patient-specific factors. These findings support WISCA-informed strategies to potentially improve clinical outcomes and mitigate antibiotic resistance in UTIs.

Keywords: antibiotic resistance; urinary tract infections; antimicrobial coverage; WISCA; weighted incidence syndromic combination antibiograms; antibiotic stewardship; Bayesian analysis

1. Introduction

Urinary tract infections (UTIs) are common bacterial infections that can lead to significant morbidity if not treated properly [1]. In outpatient settings, the increasing resistance to commonly used oral antibiotics poses a challenge [2]. Patients who receive antibiotics to which the pathogen is

resistant are more likely to require additional antibiotic courses or hospitalization [1,2]. Community-onset urinary tract infections (CoUTIs) that require hospitalization often present with a challenging resistance profile and significant clinical consequences [3]. *Escherichia coli* is the most common uropathogen in CoUTIs, followed by *Klebsiella pneumoniae*, *Enterococcus spp.*, and *Pseudomonas aeruginosa* [4]. Resistance to commonly used antibiotics such as ampicillin, cefazolin, cefuroxime, and co-trimoxazole is frequently observed, complicating the choice of empiric therapy [5]. In particular, extended-spectrum cephalosporin-resistant Enterobacterales (ESC-R EB) are associated with poor clinical outcomes, including increased rates of clinical failure and inappropriate initial antibiotic therapy (IIAT) [3].

The global burden of antimicrobial resistance (AMR) in UTIs is exacerbated by the widespread use of antibiotics, which accelerates the development of resistance mechanisms, such as β -lactamase production [6–9]. This issue is particularly pronounced in community-acquired UTIs, where resistance to fluoroquinolones has been documented to increase significantly over time in various regions, including Europe, Asia, and North America [10]. De forma similar Patients with CoUTIs due to multidrug-resistant (MDR) bacteria often experience IIAT, leading to longer hospital stays and increased healthcare costs [11]. A significant proportion of uropathogens are MDR, with rates exceeding 40% in some studies [12]. ESBL-producing strains are increasingly common, particularly among *E. coli* and *K. pneumoniae* [13]. The presence of carbapenem-resistant (CR) organisms further exacerbates these issues, as these infections are associated with longer hospital stays, higher readmission rates, and increased mortality, especially in cases with concurrent bacteremia [14].

The SENTRY Antimicrobial Surveillance Program and other studies have documented these resistance patterns, emphasizing the need for continuous surveillance and tailored antimicrobial stewardship programs to manage and mitigate the impact of AMR in UTIs across Latin America [15]. Additionally, the presence of quinolone-resistant and ESC-R EB in both community and hospital settings, with resistance rates reaching up to 64% and 49%, respectively, further emphasizes the significant public health challenge posed by AMR in our country. [16]. This heterogeneity in local epidemiology and bacterial resistance patterns complicates the selection of appropriate empirical antibiotics, reinforcing the need for customized tools based on local data [9,12,13,17].

The development of weighted-incidence syndromic combination (WISCAs) has led to significant advancements in empirical antibiotic treatment, especially in the context of multidrug-resistant organisms (MDROs) [17,18]. Traditional antibiograms, while useful, often lack syndrome specificity and fail to consider the variability in bacterial prevalence across different populations and age groups [17]. Unlike traditional antibiograms, WISCAs provide tailored recommendations by weighting the coverage probability according to pathogen prevalence [17]. This methodology has shown potential in various clinical scenarios, including pediatric urinary tract infections [19], health care associated urinary tract infections [20,21], neonatal sepsis [22] neonatal meningitis [23], febrile neutropenia in oncologic pediatric patients [24], ventilator associated pneumonia [18], and pediatric bloodstream infections [25].

WISCAs incorporating Bayesian hierarchical models, which adjust for patient-specific factors such as age, sex, or prior antibiotic exposure, provide further refinement of these coverage estimates [21,24,26]. This methodology, adaptable to various clinical scenarios involving infectious syndromes, offers probability distributions for coverage adjusted to local resistance patterns and accounts for patient-specific factors, resulting in more accurate and tailored antibiotic recommendations [19,24,26]. Integrating WISCAs into clinical practice supports antimicrobial stewardship by providing timely, evidence-based guidance for antibiotic regimen selection, potentially reducing resistance, and preserving antibiotic efficacy [19,22,24].

Most of the previous studies on WISCA design have been conducted in high-income countries with varying levels of AMR [17–27]. This study aims to address the knowledge gap regarding WISCA development for UTIs in middle-income Latin American countries, where high rates of AMR among Enterobacterales are prevalent [15,16]. Specifically, our objective was to evaluate the impact of inappropriate antimicrobial therapy in patients with community-onset UTIs who required

hospitalization at our institution and to develop a WISCA designed to enhance appropriate use of antibiotics, thereby supporting antimicrobial stewardship initiatives.

2. Materials and Methods

2.1. Population and Eligibility Criteria

We conducted a retrospective review of medical records for patients admitted to the Emergency Department and medical wards of our university hospital with International Classification of Diseases 10th Edition (ICD-10) codes N39.0 (Urinary Tract Infection), N30.0 (Acute Cystitis), and N10 (Acute Pyelonephritis). Urinary culture results from the hospital's microbiology laboratory were analyzed, including samples obtained via clean-catch midstream or catheterization from patients presenting to the Emergency Department or hospitalized in medical wards. Only initial cultures reported as positive were included. Records from patients admitted between June 2021 and December 2024 were reviewed. Patients included in the analysis met the following criteria: they presented to the Emergency Department with community-onset symptoms of a UTI, their hospitalization was due to a UTI diagnosis, and they had a positive urinary culture at admission, which was defined as a UTI event. Patients younger than 18 years of age, those with more than 10% missing data from clinical records, and those with episodes of asymptomatic bacteriuria were excluded.

Demographic data, comorbidities, the Charlson Comorbidity Index (CCI), the Sequential Organ Failure Assessment (SOFA) score, and other clinical and microbiological parameters were systematically collected from patient medical records and microbiology registries. The collected data included variables such as the type of urinary infection (complicated versus uncomplicated), site of UTI (pyelonephritis or cystitis), use of a permanent urinary catheter, and presence of recurrent UTI (defined as at least three episodes within 12 months or at least two episodes within six months). UTI-related local complications (e.g., renal or perinephric abscess, emphysematous cystitis or pyelonephritis, xanthogranulomatous pyelonephritis, malakoplakia, and renal papillary necrosis) and systemic complications (e.g., bacteremia, UTI-related hypotension, and UTI-related need for vasopressors) were also recorded. Additional variables included the initial empiric treatment used, time to correction of inappropriate initial antibiotic therapy, prior antibiotic use within 90 days, prior hospitalization within 90 days, presence of hypotension, need for vasopressors, in-hospital mortality, length of hospital stay, and susceptibilities of the isolated microorganisms. If the isolated microorganism was not susceptible to the initial empiric treatment, the event was classified as inappropriate initial antibiotic treatment. Furthermore, if inappropriate initial antibiotic treatment was not corrected based on urinary culture results during the UTI episode, it was classified as inappropriate final antibiotic treatment (IFAT).

For the type of UTI, we considered complicated UTIs in patients with functional, metabolic, or anatomical factors such as diabetes mellitus; immunocompromised status; nephrolithiasis; urinary obstruction; neurogenic bladder; indwelling urinary catheter or other device; and male sex [28–30].

The bacterial isolates were identified using standardized techniques and methods, and their antibiotic susceptibility was tested with the VITEK 2 system by *Biomerieux (Marcy l'Etoile, France)*. Specific panels were used in accordance with the Clinical and Laboratory Standards Institute (CLSI) guidelines [31]. Results were interpreted based on CLSI breakpoints, with extrapolations made for agents lacking established breakpoints or minimum inhibitory concentration (MIC) values, as detailed in **Table S1** of the supplementary materials [31].

2.2. WISCA Development

The WISCA model was designed on the basis of pathogens isolated from patients with UTIs. As previously described, a UTI event was considered if the patient had symptoms corresponding to a clinical syndrome of UTI and had a positive urine culture on admission. If the isolated organism was susceptible to an empiric antibiotic, it was reported as susceptible to the empirical antimicrobial

regimen employed. Intermediate susceptibility to a particular antimicrobial was reported as resistant for the model design. If there was more than one microorganism in the same urine culture, it was considered contamination and was discarded for the model design [31]. For patients with multiple subsequent urine cultures, it was verified that these patients experienced a new UTI event. Subsequent cultures taken during the initial UTI episode were excluded to avoid results influenced by pharmacological pressure. If they corresponded to a new UTI event, they were included in the analysis.

We studied the coverage of regimens according to local and international guidelines and for which automated sensitivity testing is routinely performed [31–33]. These regimens included fosfomycin (FOS), nitrofurantoin (NIT), trimethoprim/sulfamethoxazole (SXT), amikacin (AN), ciprofloxacin (CIP), piperacillin/tazobactam (TZP), ceftazidime (CAZ), ceftriaxone (CRO), cefepime (FEP), ertapenem (ETP), and meropenem (MEM). Additionally, combination therapies, including ceftazidime + amikacin (CAZ + AN), ceftriaxone + amikacin (CRO + AN), piperacillin/tazobactam + amikacin (TZP + AN), cefepime + amikacin (FEP + AN), meropenem + amikacin (MEM + AN), ertapenem + amikacin (ETP + AN), ceftazidime + linezolid (CAZ + LNZ) or vancomycin (CAZ + VA), ceftriaxone + linezolid (CRO + LNZ) or vancomycin (CRO + VA), piperacillin/tazobactam + linezolid (TZP + LNZ) or vancomycin (TZP + VA), cefepime + linezolid (FEP + LNZ) or vancomycin (FEP + VA), ertapenem + linezolid (ETP + LNZ) or vancomycin (ETP + VA), and meropenem + linezolid (MEM + LNZ) or vancomycin (MEM + VA), were studied.

2.3. Statistical Analysis

Demographic data are reported as simple relative frequencies. The normality of the data distribution was assessed via the Shapiro–Wilk test. Pearson’s chi-square test and Fisher’s exact test were used to compare proportions, as appropriate. For comparisons of quantitative variables, Student’s *t* tests and Wilcoxon–Mann–Whitney tests were used for normally and nonnormally distributed data, respectively.

To identify the impact of an inadequate antibiotic regimen on clinical outcomes, a multivariate logistic regression model was used. The clinical outcomes were in-hospital mortality and extended hospital stays, defined as a stay exceeding the 90th percentile of our patient population. The variables included in the model had a *p* value of <0.1 and were considered to be biologically plausible. A stepwise method for the selection of variables in the model was used. The Hosmer–Lemeshow test was used to determine the goodness of fit, with values greater than 0.1 considered appropriate.

To identify factors associated with antimicrobial resistance and determine subgroups where the WISCA tool could be applied to promote the rational use of antibiotics, we performed univariate analyses of demographic, clinical, and comorbidity factors with biological plausibility for the presence of resistance. Based on the methodologies described, in the development of a WISCA (which considers all organisms causing a particular infectious syndrome), we used fluoroquinolone non-susceptibility and third-generation cephalosporin non-susceptibility in the entire set of patients (and consequently all microorganisms) as dependent variables, given their established role as markers of resistance in microorganisms associated with UTIs. [16,17,26,34,35]. Variables identified as associated with resistance, along with those with biological plausibility, were subsequently used in the exploration of subgroups within the WISCA framework.

Finally, in accordance with previous studies [17,18,20,24,26] we employed a Bayesian hierarchical logistic regression model to estimate the probability that a given antibiotic regimen covers a specific pathogen causing a urinary tract infection (UTI), covariates for age categories (<30 years, 30–65 years and >65 years) and sex were included to account for potential variability. In particular, we introduced:

- Fixed effects for age-stratum (<30 years, 30–65 years and >65 years) and sex, capturing systematic coverage differences across these patient subgroups.

- Random effects for both regimen and pathogen, allowing each regimen and pathogen to share information with one another, thereby reducing the uncertainty in regimens or pathogens with few observations.

Such an approach extends the Weighted Incidence Syndromic Combination Antibigram (WISCA) framework by combining the partial pooling advantage with a straightforward means to incorporate patient covariates.

Model Specification

Let $i=1,\dots,N$ index the observations, each corresponding to a patient with covariates, a chosen antibiotic regimen $j[i]$, and an identified pathogen $k[i]$. Define:

- $\text{coverage}_i \in \{0,1\}$: an indicator that regimen j successfully covers pathogen k in the i -th observation.
- Age and sex: patient-level covariates (age group, sex).
- α_0 : overall intercept.
- $\alpha_{\text{pathogen}[k]}$: random intercept for pathogen k .
- $\alpha_{\text{regimen}[j]}$: random intercept for regimen j .
- β_1, β_2 : fixed-effect coefficients for age group and sex, respectively.

The probability of coverage was modeled via a Bernoulli likelihood with a logit link:

$$\text{logit}(P(\text{coverage}_i=1)) = \alpha_0 + \alpha_{\text{pathogen}[k[i]]} + \alpha_{\text{regimen}[j[i]]} + \beta_1 \text{age group}_i + \beta_2 \text{sex}_i.$$

Hence,

$$P(\text{coverage}_i=1) = \text{logit}^{-1}(\alpha_0 + \alpha_{\text{pathogen}[k[i]]} + \alpha_{\text{regimen}[j[i]]} + \beta_1 \text{age group}_i + \beta_2 \text{sex}_i).$$

Random effects, we assume:

$$\alpha_{\text{pathogen}[k]} \sim N(0, \sigma_{\text{pathogen}}), \alpha_{\text{regimen}[j]} \sim N(0, \sigma_{\text{regimen}}).$$

We placed moderately informative priors on model parameters, ensuring stability yet allowing the data to drive estimates:

- Intercept (α_0): e.g., $\alpha_0 \sim \text{Student-}t(\nu=3, \mu=0, \sigma=10)$.
- Fixed effects (β_1, β_2): e.g., $\beta \sim \text{Student-}t(3,0,1)$ or $\text{normal}(0, 1)$.
- Random effects' standard deviations ($\sigma_{\text{pathogen}}, \sigma_{\text{regimen}}$): e.g., half-Student- t or half-normal, such as:

$$\sigma_{\text{pathogen}} \sim \text{normal}^+(0,0.3), \sigma_{\text{regimen}} \sim \text{normal}^+(0,0.3).$$

Prior Distributions

We placed moderately informative priors on model parameters, ensuring stability and allowing the data to drive estimates:

- **Intercept** (α_0): e.g., $\alpha_0 \sim \text{Student-}t(\nu=3, \mu=0, \sigma=10)$.
- **Fixed effects** (β_1, β_2): e.g., $\beta \sim \text{Student-}t(3,0,1)$ or $\text{normal}(0, 1)$.
- **Random effects' standard deviations** ($\sigma_{\text{pathogen}}, \sigma_{\text{regimen}}$): e.g., half-Student- t or half-normal, such as:

$$\sigma_{\text{pathogen}} \sim \text{normal}^+(0,0.3), \sigma_{\text{regimen}} \sim \text{normal}^+(0,0.3).$$

These priors regularized extreme intercepts and keep the model stable when data were sparse in certain pathogen-regimen combinations.

The Hamiltonian Monte Carlo (HMC) algorithm, implemented via Stan software, was used for Bayesian inference. The algorithm executed four chains, each with 10,000 iterations, discarding the first 3,000 iterations as a warm-up period. Convergence was assessed using trace plots, R-hat diagnostic indices, and autocorrelation evaluations.

The parameters and coverage estimates were presented as medians accompanied by 95% highest density intervals (HDIs), where broader HDIs reflect higher levels of uncertainty. These summaries are robust to potential skewness in the posterior distributions, providing reliable and informative estimates of antibiotic regimen coverage. Autocorrelation for the parameters was evaluated at distal lags (e.g., between X_t and X_{t+h} , where $h \geq 2$) to ensure the independence of samples across iterations.

The statistical analysis was conducted using R software (version 4.3.0). Data management and manipulation were primarily performed with the *dplyr* and *tidyr* packages. The Bayesian hierarchical models were fitted via *brms* (version 2.19.0), which interfaces with the Stan framework for model specification and Hamiltonian Monte Carlo sampling. Model diagnostics and posterior summaries were completed using the *posterior*, *bayesplot*, and *tidybayes* libraries (versions 1.3.0, 1.10.0, and 3.0.5, respectively), ensuring rigorous evaluation of convergence and robust inference.

3. Results

According to the specified criteria, we documented 242 episodes of patients with CoUTIs who presented to the emergency department, as depicted in **Figure 1**.

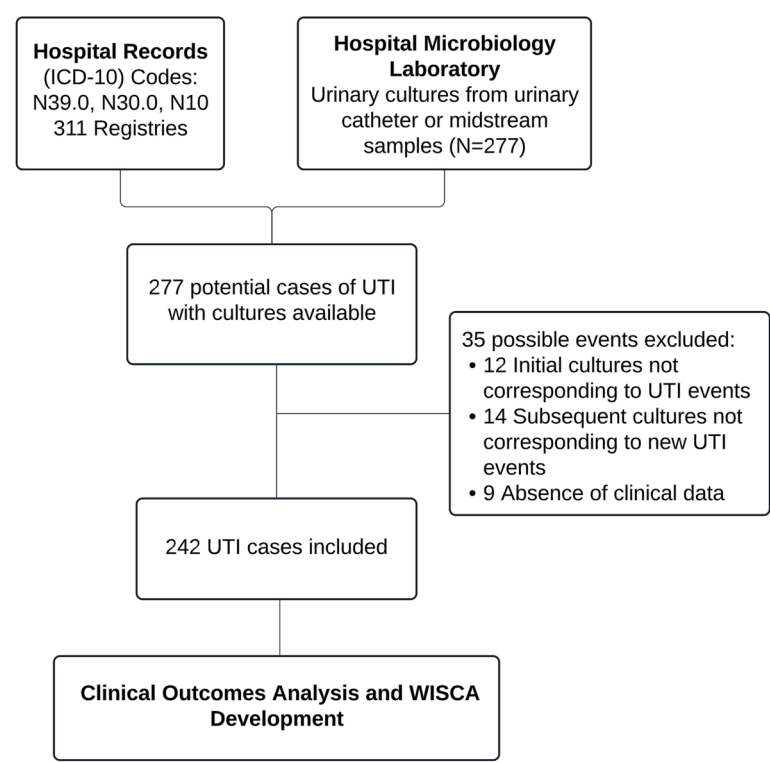


Figure 1. Flowchart of case selection.

The median age of the patients was 53.0 years (IQR 40.0-63.0), and 68.18% were female (n = 165). Among the documented CoUTI events, 99.17% were complicated urinary tract infections (n = 240), and 97.11% were reported as pyelonephritis (n = 235). Permanent urinary catheters were used in 11.57% of patients due to underlying anatomical or functional conditions (n = 28). The median Charlson Comorbidity Index was 3 (IQR 1–5), and the median SOFA score was 2 (IQR 0–6). Hypotension associated with infection was observed in 23.14% of patients (n = 56), and 18.59% required vasopressors (n = 45). The median length of hospital stay was 6.5 days (IQR 4–10), with the 90th percentile of hospital stays at 15 days. The in-hospital mortality rate was 14.05% (n = 34). The remaining sociodemographic data, comorbidities, and clinical characteristics are presented in **Table 1**.

Table 1. Sociodemographic, comorbidities and clinical characteristics of patients who presented to the emergency department with community-onset UTIs.

Variable	Obs* (n =242)	Total (n =242)	Women (n = 165)	Men (n = 77)	p value
Sociodemographic and clinical characteristics					

Age - mean (IQR)	165 / 77	53 (40.0-63.0)	51 (38.0-63.0)	55.0 (47.0-67.0)	0.036
Complicated UTI - n (%)	165 / 77	240 (99.17)	163 (98.78)	77 (100.00)	0.101
Permanent urinary catheter - n (%)	165 / 77	28 (11.57)	11 (6.66)	17 (22.08)	0.001
Recurrent UTI - n (%)	165 / 77	93 (38.43)	58 (35.2)	35 (45.45)	0.164
Pyelonephritis - n (%)	165 / 77	235 (97.11)	158 (95.8)	77 (100.00)	0.101
Initial inappropriate antibiotic treatment - n (%)	164 / 75	72 (29.75)	44 (26.67)	28 (36.36)	0.136
Days to correction of empirical treatment - median (IQR)	44 / 28	3.0 (2.0-4.0)	3.0 (2.0-4.0)	3.00 (2.75-4.0)	0.965
Inappropriate final antibiotic treatment in UTI event - n (%)	164 / 77	20 (8.26)	11 (6.67)	9 (11.69)	0.263
Previous antibiotic treatment within 90 days - n (%)	165 / 77	88 (36.36)	50 (30.30)	38 (49.35)	0.033
Previous hospitalization within 90 days - n (%)	165 / 77	108 (44.63)	69 (41.82)	39 (50.65)	0.251
Severity of Illness					
Hypotension - n (%)	165 / 77	56 (23.14)	37 (22.42)	19 (24.68)	0.823
Need for vasopressors - n (%)	165 / 77	45 (18.59)	31 (18.79)	14 (18.18)	1.000
SOFA Score - median (IQR)	165 / 77	2 (0.0-6.0)	2 (0.0-6.0)	3.0 (0.0-7.0)	0.758
Comorbidities					
Charlson comorbidity index - median (IQR)	165 / 77	3 (1.0-5.0)	3 (1.0-5.0)	3 (2.0-5.0)	0.858
Diabetes mellitus - n (%)	165 / 77	150 (61.98)	105 (63.64)	45 (58.44)	0.527
Hypertension - n (%)	165 / 77	82 (33.88)	62 (37.58)	20 (25.97)	0.103
Cardiovascular disease - n (%)	165 / 77	48 (19.83)	35 (21.21)	13 (16.88)	0.540
Acute kidney injury - n (%)	165 / 77	105 (43.39)	65 (39.39)	40 (51.95)	0.090
Chronic kidney disease - n (%)	165 / 77	68 (28.10)	49 (29.70)	19 (24.68)	0.512
Chronic liver disease - n (%)	165 / 77	4 (1.65)	3 (1.82)	1 (1.30)	1.000
Pregnancy - n (%)	165 / 77	9 (3.72)	9 (5.45)	0.0 (0.0)	0.061
Immunosuppression - n (%)	165 / 77	14 (5.79)	10 (6.06)	4 (5.95)	1.000
Cancer - n (%)	165 / 77	15 (6.20)	11 (6.67)	4 (5.95)	0.780
Central nervous system neurological disease- n (%)	165 / 77	28 (11.57)	11 (6.67)	17 (22.08)	0.001
Previous diagnosis of peripheral neuropathy - n (%)	165 / 77	26 (10.74)	23 (13.94)	3 (3.90)	0.033
Outcome					
Urinary tract infection related complications	165 / 77	68 (28.10)	47 (28.48)	21 (27.27)	0.967
Local complication of UTI - n (%)	165 / 77	29 (11.98)	20 (12.12)	9 (11.69)	1.000
System complication of UTI n (%)	165 / 77	48 (19.83)	34 (20.60)	14 (18.18)	0.789
Hospital stay - median (IQR)	165 / 77	6.5 (4.0-10.0)	7.0 (4.0-10.0)	6.0 (3.0-12.0)	0.899
Hospital stay > 15 days - n (%)	165 / 77	29 (11.98)	17 (10.30)	12 (15.56)	0.334
In-hospital mortality - n (%)	165 / 77	34 (14.05)	24 (14.55)	10 (12.99)	0.899

*Observations frequency.

In terms of local epidemiology and antibiotic resistance, the most frequently isolated organisms were *Escherichia coli* (71.07%, n = 172), *Klebsiella pneumoniae* (9.91%, n=24), *Pseudomonas aeruginosa* (4.54%, n = 11), and *Proteus mirabilis* (2.89%, n=7). According to standard definitions for acquired resistance, 46.44% of Enterobacterales isolates were classified as MDR (n=98), and 61.61% of the isolates were resistant to third-generation cephalosporins (n=130). No carbapenem-resistant Enterobacterales were identified. The rest of the isolated organisms and their resistance profiles are detailed in **Table 2**.

Table 2. Organisms isolated in patients with community-onset UTI during the study period.

Variable	Obs* (n =242)	Total (n = 242)	Women (n = 165)	Men (n = 77)	p value
Microorganism - n (%)					
<i>Escherichia coli</i>	165 / 77	172 (71.07)	129 (78.18)	43 (55.84)	<0.001
<i>Klebsiella pneumoniae</i>	165 / 77	24 (9.91)	14 (8.48)	10 (12.98)	0.389
<i>Pseudomonas aeruginosa</i>	165 / 77	11 (4.54)	3 (1.81)	8 (10.38)	0.006
<i>Proteus mirabilis</i>	165 / 77	7 (2.89)	6 (3.63)	1 (1.29)	0.436
<i>Candida glabrata</i>	165 / 77	6 (2.47)	5 (3.03)	1 (1.29)	0.667
<i>Enterococcus faecalis</i>	165 / 77	5 (2.06)	1 (0.60)	4 (5.19)	0.037
<i>Citrobacter freundii</i>	165 / 77	4 (1.65)	0 (0.00)	4 (5.19)	0.010
<i>Staphylococcus aureus</i>	165 / 77	3 (1.23)	2 (1.21)	1 (1.29)	1.000
<i>Candida tropicalis</i>	165 / 77	3 (1.23)	1 (0.60)	2 (2.59)	0.238
<i>Morganella morganii</i>	165 / 77	2 (0.82)	0 (0.00)	2 (2.59)	0.100
<i>Enterobacter cloacae</i>	165 / 77	1 (0.41)	1 (0.60)	0 (0.00)	1.000
<i>Staphylococcus saprophyticus</i>	165 / 77	1 (0.41)	1 (0.60)	0 (0.00)	1.000
<i>Acinetobacter baumannii</i> complex	165 / 77	1 (0.41)	1 (0.60)	0 (0.00)	1.000
<i>Candida parapsilosis</i>	165 / 77	1 (0.41)	0 (0.00)	1 (1.29)	0.318
<i>Enterococcus faecium</i>	165 / 77	1 (0.41)	1 (0.60)	0 (0.00)	1.000
Resistance profile - n (%) **					
MDR Enterobacterales	151 / 60	98 (46.44%)	70 (46.35)	28 (46.66)	1.000
3rd Gen Cephalosporin-resistant Enterobacterales	151 / 60	130 (61.61)	91 (60.26)	39 (65.0)	0.536
MDR Pseudomonas	8 / 3	3 (27.27)	0 (0.00)	3 (100)	0.055
MDR Enterococcus	2 / 4	1 (16.66)	1 (50.0)	0 (0.00)	0.429
MRSA	2 / 1	2 (66.66)	2 (100)	0 (0.00)	1.000

*Observations frequency. ** Proportions were calculated based on the total isolates of *Enterobacterales*, *Pseudomonas*, *Enterococcus*, and *Staphylococcus aureus*, respectively. MDR: Multidrug-Resistant; 3rd Gen Cephalosporin: Third-Generation Cephalosporins; MRSA: Methicillin-Resistant *Staphylococcus aureus*.

The most frequently used initial empiric therapy regimens were ertapenem (35.95%, n=87), ceftriaxone (24.38%, n=59), meropenem (9.92%, n=24), and levofloxacin (8.68%, n=21). Other regimens were used in 21.07% (n=51) of the patients. The remaining antimicrobial regimens are listed in **Table S2** of the supplementary material. Initial empiric therapy was inappropriate for 29.75% of the UTI episodes (n=72). However, in 91.74% of the UTI events, patients eventually received an appropriate regimen during the entire UTI episode. The median number of days to correct the initial empirical inappropriate empiric treatment was 3 days (IQR 2-4) (**Table 1**).

According to the univariate analyses of in-hospital mortality and extended2 hospital stay, older age, higher SOFA scores, the presence of hypotension, the need for vasopressors, a higher Charlson comorbidity index, the development of acute kidney injury (AKI) during the UTI event, a history of cardiovascular disease, chronic kidney disease, a previous cancer diagnosis, and complications related to urinary tract infection were positively associated with in-hospital mortality ($p < 0.05$). Additionally, the development of AKI during the UTI event and higher SOFA scores were positively associated with an extended hospital stay ($p < 0.01$). The remaining comparisons are detailed in **Table S3** of the supplementary material.

According to the multivariate logistic regression analyses regarding in-hospital mortality, we developed two models. Model A, which included inappropriate initial antibiotic treatment, identified higher SOFA scores (OR = 1.735, 95% CI: 1.348–2.396, $p < 0.001$) and the presence of complications related to urinary tract infection (OR = 7.207, 95% CI: 1.674–41.478, $p = 0.013$) as significant predictors of increased mortality, while initial inappropriate antibiotic treatment was not significantly associated (OR = 1.439, 95% CI: 0.416–5.057, $p = 0.563$). Model B, which incorporated inappropriate final antibiotic treatment, revealed that inappropriate final antibiotic treatment was significantly

associated with higher mortality (OR = 10.506, 95% CI: 1.311–135.265, $p = 0.042$), alongside elevated SOFA scores (OR = 1.719, 95% CI: 1.333–2.368, $p < 0.001$) and complications related to urinary tract infection (OR = 10.758, 95% CI: 2.218–84.270, $p = 0.008$). Both models demonstrated strong explanatory power (Pseudo $R^2 = 0.708$ for Model A and 0.728 for Model B) and exhibited good fit (Hosmer-Lemeshow tests, $p = 0.8546$ and $p = 0.9754$, respectively). Both models are illustrated in **Figure 2a**, while the complete set of variables for both models is provided in **Table S4** of the supplementary material.

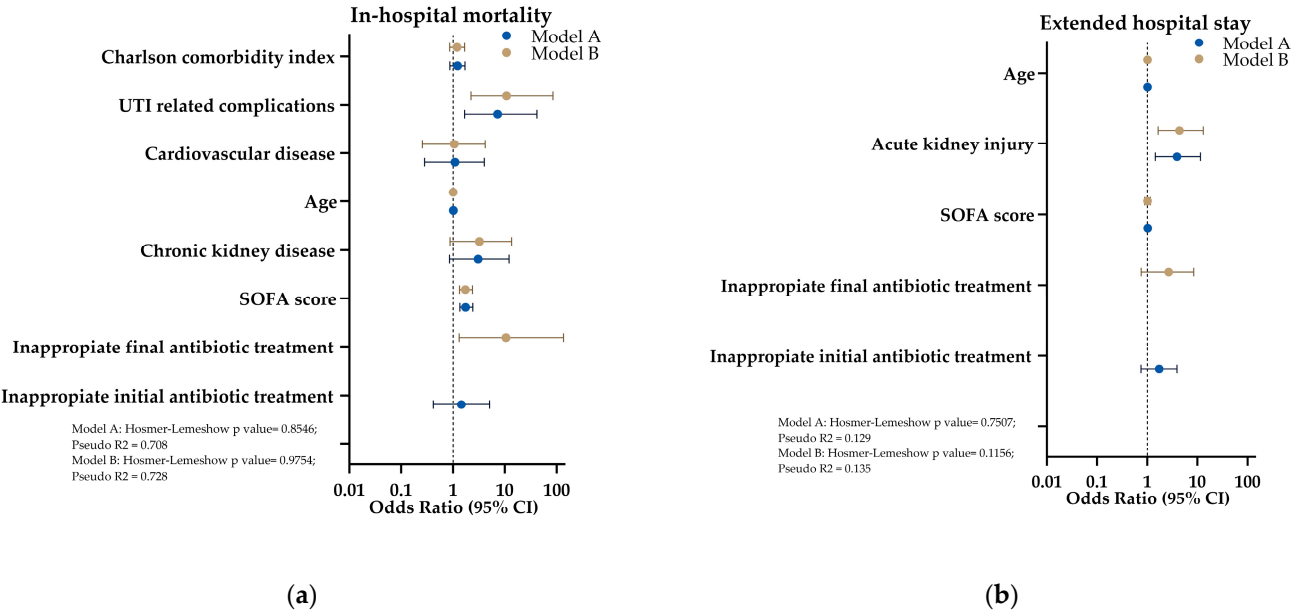


Figure 2. Logistic regression analysis for clinical outcomes. Model A includes initial empirical inappropriate therapy (blue dots and lines) as a covariate, while Model B includes inappropriate final treatment (red dots and lines) as a covariate. **(a)** Multivariate models for hospital mortality; **(b)** Multivariate models for extended hospital stay.

Regarding factors associated with extended hospital stays, we again developed models using initial empirical inappropriate treatment and final inappropriate treatment as covariates. Model A, which included inappropriate initial antibiotic treatment, revealed that only AKI was significantly associated with extended hospitalization (Odds Ratio [OR] = 3.887; 95% Confidence Interval [CI]: 1.445–11.455; $p = 0.0094$). Neither SOFA score (OR = 1.017; 95% CI: 0.891–1.155; $p = 0.7926$), increasing age (OR = 1.012; 95% CI: 0.984–1.040; $p = 0.4092$), nor initial empirical inappropriate treatment showed significant associations with extended stays. Model B, which incorporated inappropriate final antibiotic treatment, similarly identified AKI as a significant predictor of extended hospitalization (OR = 4.387; 95% CI: 1.640–13.013; $p = 0.0047$). However, SOFA score (OR = 1.013; 95% CI: 0.886–1.151; $p = 0.8482$), age (OR = 1.010; 95% CI: 0.982–1.039; $p = 0.4946$), and final inadequate treatment were not significantly associated with extended stays. Both models demonstrated modest explanatory power, with Pseudo R^2 values of 0.129 for Model A and 0.135 for Model B, and exhibited good fit (Hosmer-Lemeshow tests, $p = 0.7507$ and $p = 0.1156$, respectively). Both models are illustrated in **Figure 2b**, while the complete set of variables for both models is provided in **Table S5** of the supplementary material.

Figure 3 shows the WISCA for hospitalized urinary tract infection (UTI) cases, comparing results obtained using a hierarchical Bayesian model and a non-hierarchical approach. For low-coverage agents such as ciprofloxacin (CIP), ceftazidime (CAZ), and ceftriaxone with vancomycin (CRO+VA), the Bayesian model produced even lower median estimates. This downward adjustment likely reflects the Bayesian model’s conservative handling of sparse data, which may “shrink” estimates toward the overall mean when evidence is limited. Conversely, for regimens with high static coverage in the traditional approach (red dots in Figure 3), such as carbapenem combinations

(MEM+AN, ETP+AN), the hierarchical Bayesian estimates are slightly higher, with substantially narrower HDI ranges. For example, MEM+AN achieved a median coverage of 95.1% (HDI: 84.2–98.7%) under the Bayesian model, compared to the traditional static estimate of 93.4%. Similarly, ETP+AN’s Bayesian estimate was 95.1% (HDI: 84.2–98.6%), which aligns closely with traditional results but benefits from increased precision. Intermediate-coverage regimens such as nitrofurantoin (NIT) and aminoglycosides (AN) demonstrated a similar pattern, with Bayesian estimates refining their static counterparts by providing narrower intervals. For instance, NIT’s coverage under the Bayesian model was 52.8% (HDI: 26.9–77.3%), reflecting greater uncertainty compared to AN, which showed a higher median coverage of 64.5% (HDI: 37.7–84.3%). The Bayesian WISCA approach also highlighted meaningful differences in regimens with overlapping ranges of effectiveness. For regimens like TZP+VA and TZP+LNZ, Bayesian estimates showed narrower HDIs (e.g., TZP+LNZ: median 82.7%, HDI: 58.2–94.4%) compared to their static counterparts, emphasizing the hierarchical model’s advantage in accounting for variability across pathogen-regimen combinations. The complete set of medians of the posterior distribution and the associated 95% HDIs are shown in **Figure 3** and **Table S6** of the supplementary material.

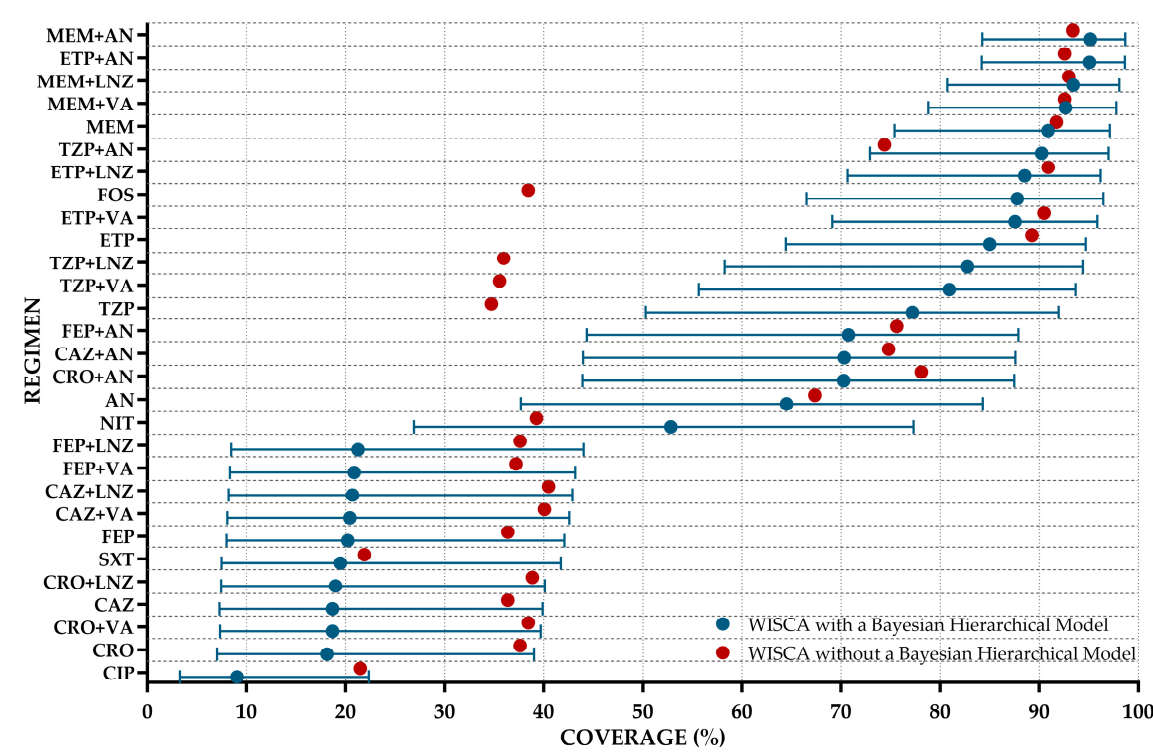


Figure 3. Weighted incidence syndromic combination antibiogram for community-onset urinary tract infections. Blue dots represent the median of the posterior distribution of antibiotic regimens (coverage), and the blue lines indicate the 95% Highest Density Interval (HDI). Red dots represent the WISCA without a Bayesian hierarchical model. Regimens are displayed in descending order of coverage.

The algorithm used to sample from the posterior distributions of the parameters achieved optimal convergence, with $R^{\wedge}$ index values near 1 and good mixing for all chains, as shown in **Table S7** of the supplementary material. The Monte Carlo Markov chain (MCMC) trace plots of the WISCA model parameters and the density plots of the posterior distributions for the general UTI population are depicted in **Figures S1** and **S2** of the supplementary materials. Autocorrelation plots demonstrated minimal autocorrelation at distal lags ($h \geq 2$), as shown in **Figure S3** of the supplementary material.

The age-stratified WISCA is presented in **Figure 4**. Some clear differences were evident among the three age groups regarding the estimated coverage rates of various antibiotic regimens. Younger

patients (<30 years) consistently demonstrated the highest coverage rates across most regimens, with carbapenem-based schemes (MEM, ETP) and their combinations (e.g., MEM+AN, ETP+AN) achieving coverage rates exceeding 90%, as reflected in the posterior median estimates. In contrast, the intermediate group (30–65 years) shows moderately lower coverage rates, with these same combinations maintaining favorable coverage (around 70–80%) but reflecting a decline in efficacy compared to younger patients. For patients over 65 years, the effectiveness of many regimens further declined, with some antibiotics—such as ceftriaxone (CRO)—showing notably low coverage levels, dropping below 40%. Moreover, specific regimens such as TZP+LNZ and TZP+VA also exhibit gradual reductions in efficacy from younger to older groups. Interestingly, fosfomycin (FOS) remained a relatively effective option even in older populations, with coverage rates nearing 70% in the >65-year group. The complete coverage data for regimens by age group are provided in **Table S10** of the supplementary materials.

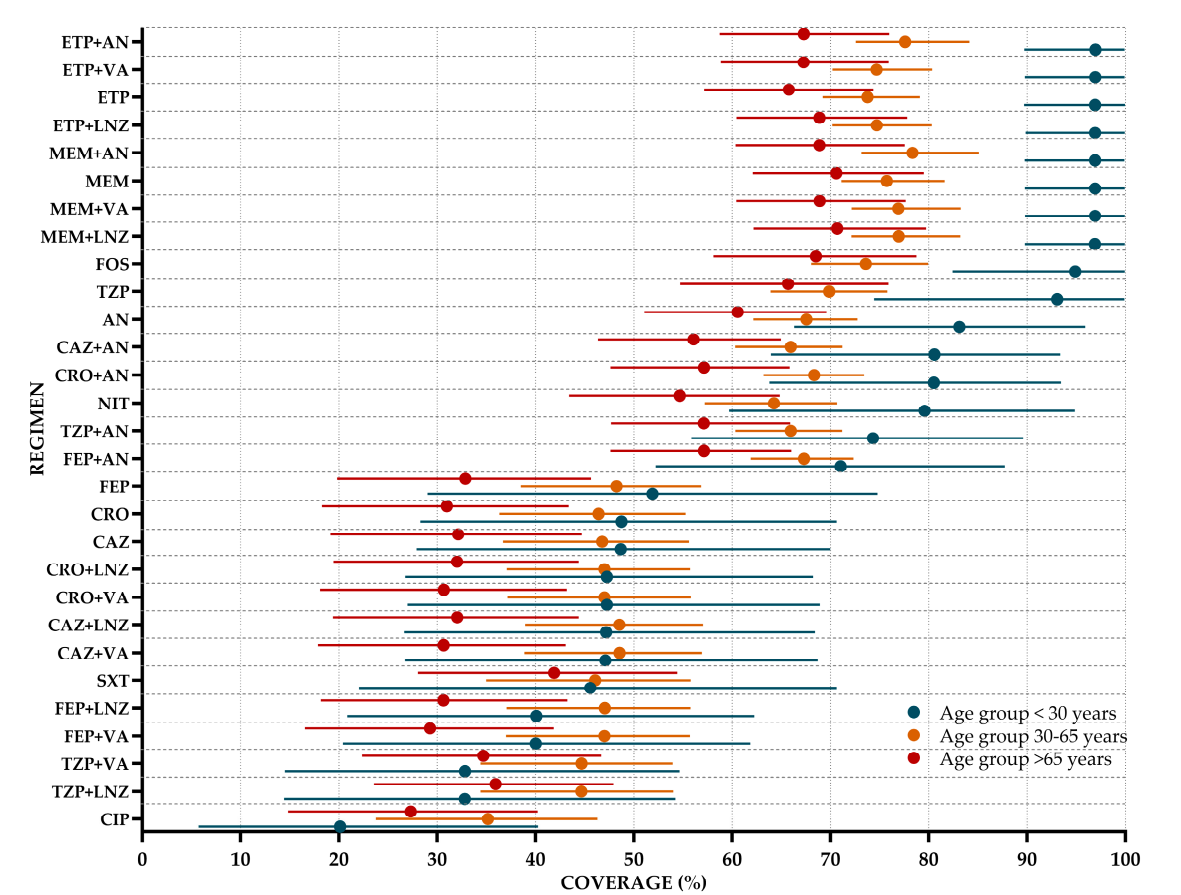


Figure 4. Weighted Incidence Syndromic Combination Antibigram for community-onset urinary tract infections across different age subgroups. Dots represent the median of the posterior distribution of antibiotic regimens (coverage), and the lines indicate the 95% Highest Density Interval (HDI). Regimens are displayed in descending order of coverage for the subgroup under 30 years of age.

According to the univariate analyses of non-susceptibility to fluoroquinolones and third-generation cephalosporins in all organisms (**Tables S8** and **S9**, respectively), factors related to comorbidities, disease severity, complications associated with urinary tract infections, and the presence of recurrent UTIs were significantly associated with the development of non-susceptibility in the general UTI population ($p < 0.05$). Building on these findings, and considering the biological plausibility of factors associated with resistance, we developed a WISCA stratified by subgroups to identify populations that could benefit from narrower-spectrum antimicrobials, while accounting for the resistance patterns observed in our population. The defined subgroups were based on the

presence (and absence) of recurrent UTIs, antibiotic use within the previous 90 days, prior hospitalization (within 90 days), and UTI-related hypotension.

The analysis of WISCA across subgroups revealed important trends in antibiotic regimen coverage, despite the 95% HDIs overlapping in comparisons. For regimens with lower median coverage, non-recurrent urinary tract infections (UTIs) consistently showed higher median coverage compared to recurrent cases. For instance, ciprofloxacin (CIP) had a median coverage of 32.9% (HDI: 21.6–44.5%) in non-recurrent UTIs versus 29.4% (HDI: 18.0–41.2%) in recurrent cases. Similarly, piperacillin-tazobactam with vancomycin (TZP+VA) exhibited higher median coverage in non-recurrent cases (median 43.5%, HDI: 33.3–52.8%) compared to recurrent ones (median 35.9%, HDI: 24.5–46.9%). For high-coverage regimens, such as carbapenem-based combinations (e.g., MEM+AN), recurrent UTIs often showed slightly higher median coverage compared to non-recurrent cases, as in the case of MEM+AN (80.6% vs. 75.1%, respectively). While the overlapping HDIs indicate that these differences may not be statistically significant, the trends observed highlight potential variability influenced by regimen-pathogen interactions.

Subgroup analyses also highlighted the influence of prior hospitalization and recent antibiotic exposure on regimen coverage. The analysis of WISCA based on prior hospitalization status revealed a trend where antibiotic regimens generally demonstrated higher median coverage in patients without a history of hospitalization compared to those with prior hospitalization. For example, fosfomycin (FOS) showed a median coverage of 72.5% (HDI: 67.7–78.9%) in patients without prior hospitalization, compared to 65.5% (HDI: 55.8–75.8%) in those previously hospitalized. Similarly, nitrofurantoin (NIT) and aminoglycosides (AN) displayed better coverage in patients without hospitalization history, with median coverage rates of 60.1% (HDI: 52.9–66.9%) and 63.6% (HDI: 57.4–68.9%), respectively, compared to 56.3% (HDI: 46.8–65.5%) and 61.8% (HDI: 53.5–70.4%) in those with prior hospitalization. However, the 95% HDIs overlapped in most cases, indicating that these differences may not be statistically significant and should be interpreted cautiously. Similarly, prior antibiotic treatment within 90 days was associated with reduced coverage for multiple regimens. For instance, the coverage of Piperacillin-Tazobactam (TZP) decreased from 69.09% (HDI: 64.21–75.34%) in patients without recent antibiotic exposure to 64.66% (HDI: 54.99–73.73%) in those recently treated.

The analysis of WISCA in patients presenting with hypotension due to urinary tract infection (UTI), with or without the need for vasopressors, revealed a consistent trend of lower median coverage across nearly all antibiotic regimens compared to those without hypotension. For example, fosfomycin (FOS) demonstrated a median coverage of 74.5% (HDI: 69.5–80.3%) in patients without hypotension, compared to 66.9% (HDI: 55.2–78.5%) in those with hypotension. Similarly, nitrofurantoin (NIT) showed a median coverage of 64.4% (HDI: 57.6–70.4%) in patients without hypotension, while it decreased to 53.3% (HDI: 42.0–64.3%) in those with this condition. Among regimens with high coverage, meropenem combined with aminoglycosides (MEM+AN) achieved a median coverage of 78.4% (HDI: 73.7–84.9%) in patients without hypotension, compared to 68.5% (HDI: 59.3–78.2%) in those with hypotension. Similarly, meropenem combined with linezolid (MEM+LNZ) demonstrated a median coverage of 78.4% (HDI: 73.7–84.9%) in patients without hypotension, dropping to 66.8% (HDI: 57.5–76.3%) in hypotensive patients. Ceftriaxone combined with aminoglycosides (CRO+AN) showed a similar pattern, with a median coverage of 67.9% (HDI: 62.8–72.7%) in patients without hypotension, compared to 55.8% (HDI: 46.3–65.3%) in those with this condition. While the 95% HDIs for most regimens overlapped between groups, the consistent trend of reduced coverage in hypotensive patients underscores the need for careful regimen selection in this clinical scenario. The WISCA across subgroups related to recurrent UTI, prior hospitalization, recent antibiotic use within 90 days, and hemodynamic stability is depicted in **Figure 5** and detailed in **Table S10** of the supplementary materials.

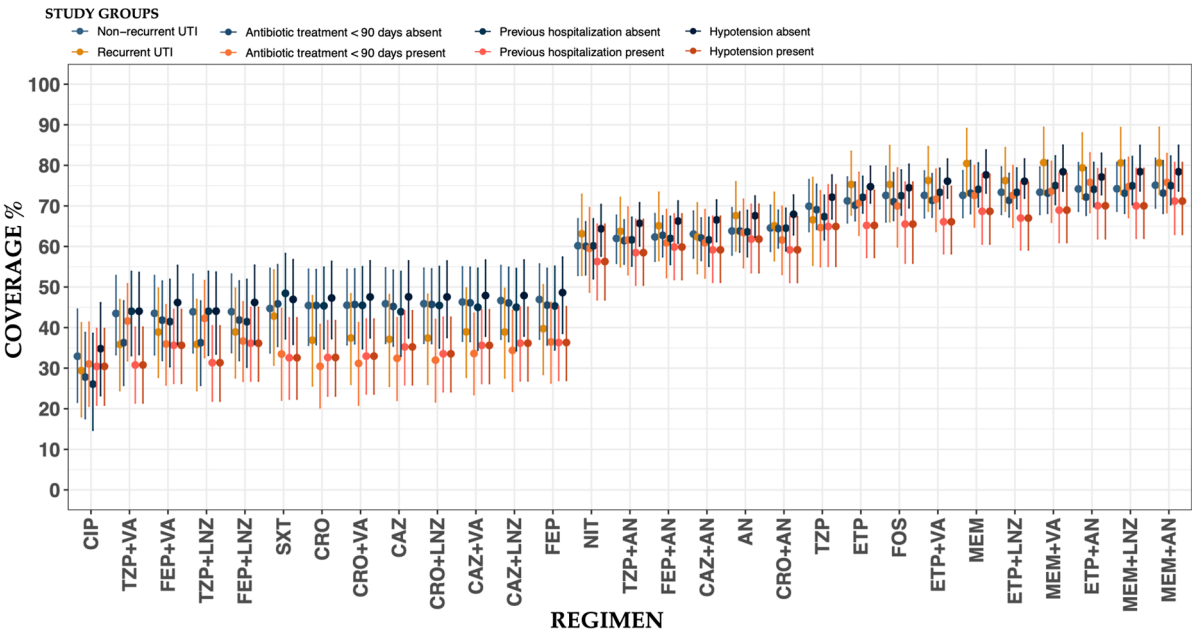


Figure 5. Weighted incidence syndromic combination antibiogram for community-onset urinary tract infections across subgroups, ordered from left to right: recurrent UTI, prior hospitalization, recent antibiotic use within 90 days, and hemodynamic stability. The X-axis represents the treatment regimens, while the Y-axis indicates the percentage of posterior coverage (dots) and the 95% highest density intervals (HDIs) (lines).

Finally, based on the previous results and to promote the rational use of antibiotics, we identified two subgroups of patients for the development of a specific WISCA. Group 1 consisted of patients without a history of recurrent UTI, prior hospitalization, recent antibiotic use within 90 days, and who were hemodynamically stable. Group 2 included patients who were hemodynamically stable and did not have comorbidities. Overall, antibiotic coverage was higher in these subgroups, particularly for regimens that previously demonstrated lower coverage, such as piperacillin-tazobactam plus aminoglycosides (TZP+AN), ceftriaxone plus aminoglycosides (CRO+AN), and piperacillin-tazobactam (TZP). In these subgroups, TZP+AN showed 73.85% coverage (HDI: 53.79–78.92%), CRO+AN reached 75.7% (HDI: 55.9–80.7%), and TZP exhibited improved coverage with a median of 79.7% (HDI: 59.3–85.7%). Despite the wide HDIs, these regimens demonstrated notable improvements in coverage compared to previous analyses. These findings suggest that these specific populations may benefit from regimens that were previously considered less effective, supporting the rational selection of antibiotics in targeted clinical scenarios. The full results are presented in **Figure 6** and **Table S10** of the supplementary materials.

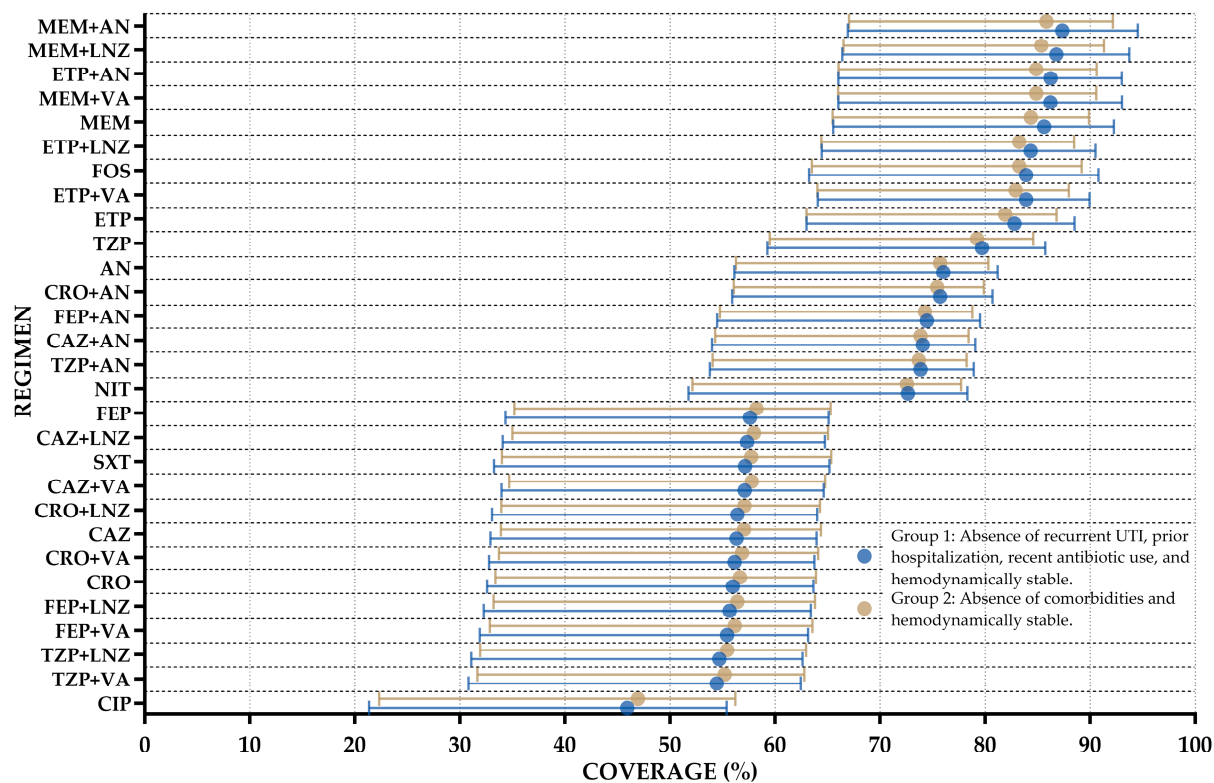


Figure 6. Weighted Incidence Syndromic Combination Antibigram (WISCA) for community-onset urinary tract infections across two subgroups with favorable resistance profiles. The X-axis represents the treatment regimens, while the Y-axis indicates the percentage of posterior coverage (dots) and the 95% highest density intervals (HDIs) (lines).

3. Discussion

In this study, we described the sociodemographic and clinical characteristics of adult patients hospitalized with urinary tract infections (UTIs), along with the resistance profiles of the causative microorganisms, and evaluated the impact of inappropriate therapy. A high frequency of antimicrobial resistance was observed, whether acquired or intrinsic to the microorganisms, particularly against fluoroquinolones (78.52%) and third-generation cephalosporins (62.40%) among the pathogens responsible for UTIs (Table S6). Furthermore, our findings revealed that inappropriate final antibiotic treatment during a UTI episode, elevated SOFA scores, and UTI-related complications—both local and systemic—significantly increased the risk of in-hospital mortality, which is consistent with reports on UTI-related mortality [36,37]. To address this, we developed a Weighted Incidence Syndromic Combination Antibigram tool, which extends the framework of traditional hospital antibiograms by providing weighted coverage estimates based on the frequency of identified pathogens. This approach was implemented using a Bayesian hierarchical model with random effects for both pathogens and treatment regimens. Our study contributes to the limited understanding of WISCA design for UTIs in middle-income countries. To the best of our knowledge, this is the first WISCA designed specifically for community onset UTIs in adult patients within a Latin American country. Previous WISCA efforts have primarily been conducted in high-income countries, particularly in Europe and the United States, or as part of international collaborations between research groups. [17,18,20,24].

The resistance profile of our institution is positioned at the higher end of the spectrum commonly reported in recent literature. [1,6,15,16,38]. Specifically, the proportion of multidrug-resistant (MDR) Enterobacterales (46.44%) and the elevated resistance to third-generation cephalosporins (61.61%) are concerning from both a clinical management and public health perspective [39,40]. These findings align with the global trend of increasing antimicrobial resistance among Enterobacterales, which has been documented for fluoroquinolones, third-generation

cephalosporins, and, more recently, carbapenems [39]. The high usage of carbapenems for initial empirical therapy (46.86%) may reflect both the perceived risk and actual likelihood of encountering resistant pathogens. However, such empiric strategies may inadvertently propagate further resistance if not closely monitored and guided by local microbiological surveillance [41]. The proportion of initial empirical regimens deemed inappropriate (29.75%) underscores a critical gap between prescribing practices and actual pathogen susceptibility patterns, reinforcing the need for timely culture and susceptibility testing to inform definitive therapy. These resistance rates emphasize the importance of antimicrobial stewardship programs, which should include active local surveillance, regular antibiogram updates, and targeted interventions (e.g., de-escalation strategies and strict infection control measures) [42].

However, in our multivariate regression analysis, initial inappropriate empirical treatment was not associated with increased mortality. Instead, other determinants, such as disease severity (SOFA score) and UTI-related complications, were associated with in-hospital mortality. This finding may be explained by the rapid correction of the antimicrobial regimen in most patients with initially inappropriate empirical treatment, with a median correction time of 3 days (IQR 2–4), a relatively young patient population (mean age 53 years, IQR 40–63), and fewer comorbidities (Charlson comorbidity index 3, IQR 1–5). In contrast, inadequate final antibiotic treatment during the UTI episode was associated with increased mortality, highlighting the importance of achieving an appropriate antimicrobial regimen over the course of the UTI episode [36]. This findings aligns with international literature regarding disease severity and inadequate treatment as key risk factors for in-hospital mortality in UTIs [36,43–46]. Additionally, acute kidney injury was positively associated with an extended hospital stay in our multivariate model, a finding consistent with international literature [47,48]. Neither initial inappropriate empirical treatment nor inadequate directed treatment during the UTI episode was associated with extended hospital stay in our models.

From the WISCA design for the general UTI population, we can infer that the most commonly used therapies in our institution, for which automated sensitivity testing is routinely performed, have coverage rates ranging from 9.04% for ciprofloxacin to 95.13% for meropenem + amikacin. Due to the high incidence of third-generation cephalosporin-resistant Enterobacterales in our institution, fosfomycin and carbapenem-based regimens demonstrated better coverage compared to non-carbapenem-based regimens.

The non-Bayesian WISCA design yielded coverage rates comparable to those of the Bayesian hierarchical WISCA model, with most coverages rates falling within the 95% HDI of the Bayesian model. However, certain discrepancies were observed in specific scenarios, where coverage rates fell outside the HDIs of their Bayesian counterparts. The Bayesian hierarchical WISCA approach generally produces different coverage estimates compared to the traditional (non-Bayesian) WISCA, as it explicitly accounts for inherent variability and correlations among samples or subpopulations through “partial pooling” of information across strata. In contrast, the conventional WISCA approach treats observations independently, which may lead to under- or overestimation of coverage for certain antibiotic regimens. For example, regimens such as piperacillin/tazobactam or fosfomycin exhibited markedly higher coverage under the Bayesian model, likely due to the hierarchical prior structure that moderates extreme estimates, shrinking them toward a more stable central tendency. Conversely, some regimens show lower coverage estimates in the Bayesian model when the conventional method overestimates their efficacy. These discrepancies highlight the Bayesian framework’s advantage of integrating prior knowledge with observed data, resulting in more robust interval estimates and better accounting for small sample sizes or heterogeneous resistance patterns across clinical isolates. This approach is particularly valuable for managing complex infections, such as UTIs requiring hospitalization, where local resistance patterns significantly influence pathogen susceptibility. As a result, Bayesian hierarchical modeling provides more adaptive and subgroup-specific coverage estimates, in contrast to the “static coverage” rates generated by traditional WISCA designs.

Regarding coverage across regimens in the general population, it is important to highlight the role of fosfomycin, which demonstrated high coverage and consistent performance, even in patients previously treated with antibiotics. Recent studies support fosfomycin as an effective option for treating urinary tract infections, including those caused by ESBL-producing bacteria, and as a step-down therapy for complicated cases [49,50]. Currently, fosfomycin is FDA-approved in the United States for the treatment of uncomplicated UTIs caused by *Escherichia coli* and *Enterococcus faecalis* in women, but it is not approved for complicated UTIs [51,52]. While fosfomycin is generally well-tolerated and effective, its use in upper UTIs should be carefully considered, particularly in complex cases such as those involving kidney transplant recipients [53,54]. The potential emergence of resistance during treatment remains a concern, emphasizing the need for further research to optimize dosing regimens and confirm its efficacy in broader clinical settings [53]. Nevertheless, these findings position fosfomycin as a viable alternative in settings with a high prevalence of multidrug-resistant organisms. [55,56].

Coverage estimates varied across age groups. Older adults (>65 years) exhibited reduced coverage for beta-lactam-based regimens, likely due to higher comorbidities and resistant organisms, while younger patients (<30 years) generally showed higher coverage, possibly linked to fewer antibiotic exposures. This is consistent regarding antimicrobial resistance is indeed a significant concern in elderly populations, particularly in the context of UTIs. Several factors contribute to this increased resistance, including age, comorbidities, and prior antimicrobial use [57]. Combination therapies with carbapenems or piperacillin/tazobactam, particularly when paired with amikacin or linezolid, demonstrated strong coverage across all age groups.

Coverage also varied by study subgroup, with lower rates generally observed in patients with recurrent UTIs, recent hospitalization, or recent antibiotic use, likely reflecting the selection of resistant strains. Notably, patients with UTI-related hypotension consistently showed lower coverage compared to those without this condition, highlighting the judicious use of broad-spectrum regimens in such cases [58–60]. Broad-spectrum agents, such as carbapenems (with or without amikacin) and fosfomycin, consistently provided robust coverage across subgroups, underscoring the importance of considering prior antibiotic history and clinical status when selecting empirical therapy [58].

Finally, based on previous analyses, we defined two groups to facilitate practical decision-making for clinicians initiating antimicrobial regimens in our institution, aiming to promote rational antibiotic use informed by microbiological evidence. Group 1 (patients without recurrent UTI, recent hospitalization, or recent antibiotic use, and who are hemodynamically stable) and Group 2 (patients without comorbidities and also hemodynamically stable) consistently achieved higher coverage across most regimens compared to previously analyzed subgroups. This pattern likely reflects a lower baseline risk of encountering resistant organisms due to fewer predisposing factors. These findings could encourage clinicians to initially use regimens such as TZP, TZP+AN, CRO+AN, and FEP+AN, thereby sparing carbapenems and later de-escalating to fosfomycin, particularly in clinically stable patients with a lower risk of MDR organisms causing UTIs.

The limitations of our study are primarily due to its retrospective nature. First, during data collection, some variables were missing and had to be excluded from the analyses. Additionally, the single-center nature of this study limits the generalizability of the results to other settings, particularly those with differing patient demographics, healthcare resources, or antimicrobial resistance patterns. However, this limitation aligns with one of the primary purposes of WISCA: to generate an antibiogram tailored to the specific conditions and local scenarios, especially considering the particular resistance profiles. To ensure accurate and reliable results regarding the impact of inappropriate treatment on clinical outcomes, further studies using prospective cohorts or clinical trials within the same population and region are recommended. This approach will enable the inclusion of more comprehensive data and improved control of variables. Moreover, while the study relied on antimicrobial susceptibility testing results, these do not always translate directly to clinical efficacy due to the complex interplay of pharmacokinetic and pharmacodynamic factors in vivo.

Finally, while the WISCA model is specifically designed to capture the dynamic interactions between antimicrobial regimens and pathogens, adjusting for patient-specific characteristics and clinical scenarios, its performance remains dependent on the quality and comprehensiveness of the input data. Although the model accounts for variables such as age, sex, and prior antibiotic exposure to tailor predictions to specific patient populations, it may not fully encompass unmeasured or unknown confounding factors that could influence outcomes, such as underlying comorbidities, immune status, or variations in healthcare practices.

Overall, these findings underscore the significant challenges in managing antibiotic resistance. As a reference center serving the uninsured population in western Mexico, our university hospital unit encounters a higher incidence of patients requiring hospitalization for urinary tract infections compared to what is typically reported. A notable strength of this study was the development of WISCA models utilizing a Bayesian hierarchical framework, which were specifically tailored to the local epidemiology characterized by a high resistance profile. The analysis of posterior coverages and 95% high-density intervals (HDIs) revealed significant variability in antibiotic coverage, predominantly influenced by resistance patterns. These results emphasize the critical importance of carefully selecting empirical therapies and continuously monitoring and adapting antibiotic policies to combat resistance effectively. Moreover, the expansion of WISCA models to other infectious syndromes will further enhance the capacity for evidence-based empiric therapy regimens, addressing the heterogeneity of local epidemiology. These efforts collectively support antimicrobial stewardship by providing actionable tools to implement rational empiric antibiotic treatments, ultimately improving patient outcomes in our population.

4. Conclusions

This study highlights the significant challenges posed by antimicrobial resistance in the treatment of urinary tract infections (UTIs) in a middle-income Latin American context. By developing a Bayesian hierarchical Weighted Incidence Syndromic Combination Antibigram model tailored to the unique resistance profiles of the study population, we demonstrated its potential to guide empirical antibiotic therapy effectively. The model's capacity to provide precise, subgroup-specific coverage estimates underscores its utility in optimizing antimicrobial stewardship strategies. Our findings reveal a high prevalence of multidrug-resistant pathogens, particularly Enterobacterales, and underscore the impact of inappropriate final antibiotic treatment on clinical outcomes, particularly in-hospital mortality. These results emphasize the necessity of incorporating local epidemiological data into antibiotic policies and the value of tools like WISCA in enhancing evidence-based decision-making. Future efforts should focus on expanding the application of WISCA models to other infectious syndromes and integrating prospective data to strengthen the evidence base for empirical treatment regimens in diverse clinical settings.

Supplementary Materials: The following supporting information can be downloaded at the website of this paper posted on Preprints.org. Table S1. Extrapolations and assumptions applied in cases of missing breakpoints or MIC values for antimicrobial agents on clinical microbiology reports. Table S2. Frequencies of initial empirical antibiotic regimens in patients with UTI requiring hospitalization. Table S3. Univariate analysis of factors associated with mortality and extended hospital stays. Table S4. Multivariate logistic regression analyses regarding in-hospital mortality. Table S5. Multivariate logistic regression analyses regarding extended hospital stay. Table S6. Weighted incidence syndromic combination antibiogram for urinary tract infection. Median estimated coverages (expressed as percentages) and 95% HDIs for monotherapies and combined regimens in the WISCA model. Table S7. R² Index values of WISCA model parameters in patients with community-onset urinary tract infections. Table S8. Univariate analysis of non-susceptibility to fluoroquinolones in patients with community-onset urinary tract infections. Table S9. Univariate analysis of non-susceptibility to third-generation cephalosporins in patients with community-onset urinary tract infections. Table S10. Weighted incidence syndromic combination antibiogram in patients with community-onset urinary tract infections across sub-groups. Table S11. R² Index Values of WISCA model parameters in patients with community-onset urinary tract infections across sub-groups. Figure S1. Monte Carlo Markov Chains (MCMC) traceplots of the CoUTI WISCA model parameters. Figure S2. Density plots of the posterior distributions of the CoUTI WISCA model

parameters. Figure S3. Autocorrelation plots for distal lags ($h \geq 2$) in posterior samples of the CoUTI WISCA model parameters.

Author Contributions: Conceptualization, AG-Q, BBA-C, and JB-R; methodology, LP-G and JB-R; software, JB-R; validation JB-R; formal analysis JB-R; investigation AG-Q, BBA-C, JCDA-J, PM-A, LP-G and JB-R; resources, JCDA-J, JB-R; data curation, AG-Q, BBA-C, JCDA-J, and JB-R; writing—original draft preparation, AG-Q, BBA-C, JCDA-J, LP-G and PM-A; writing—review and editing, JB-R; visualization, JB-R; supervision, PM-A and JB-R; project administration, PM-A and JB-R.

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Informed Consent Statement: As this study was conducted retrospectively and involved only deidentified data, informed consent was waived in accordance with the guidelines of the “Comité de ética en investigación en ciencias de la salud del Centro Universitario de Tlajomulco, Universidad de Guadalajara” (ethical approval number CUTLAJO/DS/CEICS/016/24) and relevant ethical regulations.

Data Availability Statement: The raw data supporting the conclusions of this article will be made available by the authors on request.

Conflicts of Interest: The authors declare no conflicts of interest.

Abbreviations

The following abbreviations are used in this manuscript:

AKI	Acute Kidney Injury
AMR	Antimicrobial resistance
AN	Amikacin
CAZ	Ceftazidime
CCI	Charlson Comorbidity Index
CI	Confidence Interval
CIP	Ciprofloxacin
CKD	Chronic kidney disease
CLSI	Clinical and Laboratory Standards Institute
CoUTIs	Community-onset urinary tract infections
CRO	Ceftriaxone
CRBS	Catheter-Related Bloodstream Infection
CVD	Cardiovascular Disease
ESBL	Extended-Spectrum Beta-Lactamase
ESC-R EB	Extended-Spectrum Cephalosporin-Resistant Enterobacterales
ETP	Ertapenem
FEP	Cefepime
FOS	Fosfomycin
HDI	Highest Density Interval
HMC	Hamiltonian Monte Carlo
IIAT.	Inappropriate Initial Antibiotic Treatment
IFAT.	Inappropriate Final Antibiotic Treatment
ICD-10	International Classification of Diseases 10th Edition
ICU	Intensive Care Unit
IQR	Interquartile Range
LNZ	Linezolid
MCMC	Monte Carlo Markov chain
MDR	Multi Drug Resistant
MDROs	Multidrug-resistant organisms
MEM	Meropenem
MIC	Minimum Inhibitory Concentration
MRSA	Methicillin-resistant <i>Staphylococcus aureus</i>
NIT	Nitrofurantoin
OR	Odds Ratio

SD	Standard Deviation
SOFA	Sequential Organ Failure Assessment
SXT	Trimethoprim/Sulfamethoxazole
TZP	Piperacillin/Tazobactam
UTI	Urinary Tract Infection
VA	Vancomycin
WISCA	Weighted Incidence Syndromic Combination Antibigram

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